# Analysis and Projection of The Influent Rate and Quality Parameters of Water in King Talal Dam 

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## DIDICATION

# To my famify <br> To whom I am proud to belong 

With Best Regards
Eng. A. Bani Fani

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## XVIII

## LIST OF SYMBOLS

P : The estimated probability of an X value
M : The Rank
n : The number of data
S : Standard deviation

X : The mean
$\mathrm{C}_{\mathrm{k}}$ : Coefficient of Kurtosis.

K : The fourth moment about the mean.
$X_{i}: \quad$ Observed data at time i.
Wi : Shapiro-Wilk statistics
$\mathrm{U}_{\mathrm{t}}: \quad$ Random process with mean equals to zero and variance equals to one.
$\beta_{\mathrm{r}}$ : Unknown parameter with $\beta_{0}=1$ and $\left|\beta_{\mathrm{r}}\right|<1$.
$Y_{i}: \quad$ Observation (data) at time equals to i.
$\lambda: \quad$ The lag operator
$\alpha_{\mathrm{q}}$ : Unknown parameter, $\alpha_{0}=1$ and $\left|\alpha_{\mathrm{q}}\right|<1$.
$\mathrm{v}_{\mathrm{t}}$ : Is a sequence of independent random variables with $\xi \mathrm{v}_{\mathrm{t}}=0$ and $\xi \mathrm{v}_{\mathrm{t}}^{2}=\sigma^{2}$.
$\mathrm{W}_{\mathrm{t}}=\nabla:$ Differentiation of Yi

P : Unknown parameter with $\rho_{0}=1$ and $\mid \rho_{\mathrm{r}}<1$.
a : The regression constant
$\mathrm{b}_{\mathrm{y}}$ : The regression coefficient
$\theta$ : Smoothing constant, $0<\theta<1$
$\rho_{\mathrm{xy}}$ : $\quad$ The correlation coefficient
p,d,q : ARIMA coefficients

# Analysis and Projection of The Influent Rate and Quality Parameters of Water in King Talal Dam 

By

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#### Abstract

King Talal Reservoir (KTR) is one of the most important projects in Jordan. This project is used to irrigate wide regions in the Jordan Valley. Studying the water quality and quantity in King Talal Reservoir (KTR), and the trends of influent rate and quality should be a priority in Jordan.


This study concentrates on studying and projection quantity and quality variables in Zarka River (main branch and supplier of KTR). The quantity variable used was the flow, and, the quality variables were (TSS, BOD5, COD, T-P, and T-N) in Zarka River. The data collected for each variable was containing 156 months from the year 1988 till the end of the year 2000. The procedure used in analyzing the six variables in Zarka River is through auto, cross, and vertical distance correlation between a point in Zarka River and another point, which was a point in the Samra's effluent. Deterministic and stochastic (ARIMA model) forecasting of the six variables were used in finding the best model to be used in projection.

The study results indicate that the ARIMA model is a good model in predicting most of the six variables. In forecasting the BOD5 variable non of the modeling satisfied the 10 percentage of mean error but ARIMA model gave the best model and the least percentage of mean error. However, ARIMA model did not give the best modeling in
the COD variable. The least percentage of mean error of all variables by ARIMA modeling was equal to $4.8 \%$ in the T-P variable. The cross and distance correlation gave information about the variables, such as: the relation of the variables together, the form of the variables in Zarka River, the source of the variables, and other information.

## 1. INTRODUCTION

### 1.1 Introduction

Jordan is one of the countries of the Middle East. Jordan has an area of 90 thousand $\mathrm{Km}^{2}$ and population of about 4.9 million in the year 1999 according to (General Statistical Department, 2000). The population growth rate is about $3.6 \%$ per year. One of the problems facing Jordan is water; Jordan's climate is classified as a semi-arid one. The climate is mainly characterized by low precipitation and humidity. Approximately $80 \%$ of Jordan's area receives an average precipitation of less than 100 $\mathrm{mm} /$ year. The evaporation rate in some parts of Jordan ranges from 5 to 80 times the average amount of precipitation (Salameh, 1996).

Water consumption categories in Jordan are mainly divided into three types. These include domestic, industrial, and agricultural water consumption. These categories constitute $25 \%, 5 \%$, and $70 \%$ of the total water consumption respectively. It should be noted that the annual agricultural water consumption in 1995 was about 639.7 MCM (RSS, 1998).

Many projects were done to increase the usage of water. One of the most important projects was King Talal Reservoir (KTR). (KTR) is one of the most important projects in Jordan and is used to irrigate wide regions in the Jordan Valley, which is estimated to be about $100 \mathrm{Km}^{2}$. The capacity of (KTR) is estimated to be about 86 million $\mathrm{m}^{3}$ with a depth of 108 m and a catchment area of $3175 \mathrm{Km}^{2}$. The inlet of water to King Talal Reservoir is mainly from Zarka Stream and Ramemen Wadi. The influent
water to the reservoir is a mixture of rainfall and springs water mixed with domestic and industrial waste (WAJ, 1998).

The difference in quality and quantity of water and wastewater in (KTR) mandates the implementation of a continuous monitoring program of quantity and quality. Monitoring activities should aim at investigating whether the effluent of (KTR) is within the limits of the irrigation standards. Forecasting the water quantity and quality should help in establishing some precautional measures, which should assist in resolving the anticipated problems.

The term water quality refers to the suitability of water to a particular purpose. Any physical, chemical, or biological property that influences the use of water is called water quality variable. Water quality standards have been developed to serve as guidelines for selecting water supplies for various activities (Boyd, 2000).

The water quality variables that will be forecasted in this thesis are TSS, BOD, COD, T-P, and T-N. Total suspended solids (TSS) is the most important physical characteristics of wastewater, which is composed of floating matter, settleable matter, colloidal matter, and matter in solution. Analytically, the total suspended solids content is defined as all matter that remains as residue upon filtration and then evaporation at 103 to $105{ }^{\circ} \mathrm{C}$. Biochemical oxygen demand (BOD) is a chemical characteristic. The most widely used parameter of organic pollution applied to both wastewater and surface water is the 5-days $\mathrm{BOD}\left(\mathrm{BOD}_{5}\right)$. This determination involves the measurements of the dissolved oxygen used by microorganisms in the biochemical oxidation of organic matter. Chemical oxygen demand (COD) is a chemical characteristic used to measure the organic and inorganic pollutants in wastewater. Finally the total Nitrogen and

Phosphorus are chemical characteristics and essential to the growth of the protista and plants. Inspite of the fact that they are chemical characteristics, they are necessary to evaluate the treatability of wastewater by biological processes (Metcalf \& Eddy, 1991).

These five water quality variables will be analyzed and forecasted using statistical methods. Any statistical analysis of data has to be based upon some assumed probability model for those data (Green \& Margerison, 1977). If a series has shown some trend or persistent pattern in its variations for long period of time in the past, it will be sensible to assume that such patterns or regularities will continue in the future (Chao, 1974). But one should take into consideration that in decision-making, forecasts are usually wrong. The magnitude of the forecasting errors experienced will depend upon the forecasting system used (Montogomery \& Johnson 1976).

The main objectives of forecasting, for environmental quality can be summarized as; firstly, the data is frequently expensive to accumulate so that using forecasting can minimize it. Secondly, correlation between the constituents may help the filling of missing data or the identification of outlier data. Finally, it is useful in detection of any deterioration in environmental quality, because early detection may provide the opportunity for controlling the problem at a lower cost before the problem magnifies (Bean \& Rover 1998).

The aim of this research is to study the water quality and quantity in (KTR). Trends of influent rate and quality will be established and analyzed as a function of time. A time series model will then be derived to represent the trends and help in making future projections.

### 1.2 Literature Review

Many researchers have investigated the subject of forecasting in the field of water quantity and quality. Some of the topics investigated include water quality of surface and ground water, water level, water flow, floods and rainfall, design and operation of biological wastewater treatments, wastewater treatment plant performance, long term effluent quality, and other papers in forecasting field.

Miao-Hsiang PENG and Jin-King LIU (2000) have studied the groundwater level forecasting with time series analysis. This study investigates the application of time series analysis methods for forecasting groundwater levels. The study site is located in western Taiwan where serious land subsidence has occurred. A series of monthly groundwater level observations made during the period 1993 and 1999 is used for the experiments. Univariate time series models including ARIMA models and the time series decomposition method are applied and the resulting accuracy is compared. Empirical results indicate that groundwater level data series in this study are cyclical. ARIMA models generate more accurate forecasts than the decomposition model. The forecasting of ARIMA models presents the characteristics of trend and seasonal variation.
A. Lehmann and M. Rode (1999) have studied long-term behavior and crosscorrelation analysis of water quality parameters of the Elbe River at Magdeburg, Germany. This study analyses weekly data samples from the Elbe River at Magdeburg between 1984 and 1996 to investigate the changes in metabolism and water quality in the Elbe River since the German reunification in 1990. Modeling water quality variables by Univariate time series models such as autoregressive component models and

ARIMA models reveals the improvement of water quality due to the reduction of wastewater emissions since 1990. The models are used to determine the long-term and seasonal behavior of important water quality parameters. A new procedure for testing the significance of a sample correlation coefficient is discussed. The cross-correlation analysis is applied to hydro-physical, biological, and chemical water quality variables of the Elbe River since 1990. Special emphasis is laid on the detection of spurious sample correlation coefficients.
D. P. Solomatine, C. J. Rojas, S. Velickov and J. C. Wüst (2000) have studied Chaos theory in predicting surge water levels in the North Sea. The problem of predicting surge water levels is important for ship guidance and navigation. The data collected in the coastal waters of the Netherlands (Hook of Holland) is analyzed with an objective of making such prediction. It was found that the correlation between data on surge, temperature, air pressure and wind is not sufficient to rely only on the inputoutput (connectionist) models like neural networks. It appeared that the surge time series in itself has enough information to make predictions. The applied linear prediction methods including autocorrelation and ARIMA models could not provide sufficient accuracy. Features of chaotic behavior were identified in surge, and methods of chaos theory were applied. The predictions are quite accurate (RMSE is 3.6 cm for 1 hour, and 6.1 cm for 3 hours). Possible techniques allowing for increase of the prediction accuracy and horizon (wavelet analysis, data mining techniques) were also identified.

Rodel (1997) studied the time series analysis of rising underground salt-water from the abandoned Werra potash-mining district in Germany. Time series analysis has
provided detailed information about the rates of discharge of salt-water disposed of under the Werra potash-mining district.

Hammett (1988) had studied the water use, land use, stream flow, and water quality characteristics of the Charlotte Harbor inflow area, USA-Florida. Florida is being subjected to increasing environmental stress by rapid population growth and development. So increase in freshwater demand is required. The three major rivers exist there flow into the harbor. Time series analysis was used in one of the rivers stations to forecast a decreasing trend after a long-term analysis. The increased population will require an additional 70 million gal/ day.

Berthouex and Box (1996) had studied time series models for forecasting wastewater treatment plant performance. The time series model has the form of an exponentially weighted moving average (EWMA). The interpretation of the model is that response of the system can be predicted by deviations from the EWMA smoothed values of the predictor variables.

Stegmann, Ehring, and Liem (1978) studied the application of time series analysis in water quality management and the prediction of the changes in water quality of the Oker River in Germany, on the basis of trends in dissolved oxygen and chloride levels. The environmental impacts of external factors were qualified by measuring the difference between measured water quality indicators and the extrapolated values.

Ellis, Ge , and Grasso (1990) studied the time series analysis of wastewater quality. In the treatment facilities, accurate time series forecasts must be available as
input data. The wastewater influent variables were analyzed using the application of ARMA and multivariate ARMA.

Paul A. Conrads (1998) has studied the effects of model output time averaging on the determination of the assimilative capacity of the Waccamaw River and Atlantic Intracoastal Waterway near Myrtle Beach, South Carolina. A branched Lagrangian transport model was calibrated and validated for the tidally influenced portions of the Waccamaw and Pee Dee Rivers, Bull Creek, and the Atlantic Intracoastal Waterway near Myrtle Beach, South Carolina. In determining the assimilative capacity of the Atlantic Intracoastal Waterway, 1-hour, 24-hour, 14-day, and 30-day averaging intervals were used. For each averaging interval, point-source loadings in the model were increased until the state dissolved-oxygen standard was violated. Results of the averaging intervals and point-source loadings for two locations were evaluated by comparing time series of dissolved-oxygen concentration at critical locations and longitudinal profiles of average dissolved-oxygen concentration for particular reaches of the system. The concentrations of the oxygen-consuming constituents that can be assimilated vary by 180 percent, depending upon the averaging interval used for interpreting the simulation model output
M. Stein and J. Lloret (1999) have studied forecasting of air and water temperatures using Fishery Statistical Methods to describe and forecast monthly mean air and bottom water temperatures from 3 sites in the Northwest Atlantic region, up to one year in advance. ARIMA (Auto-Regressive-Integrated-Moving- Average) models were developed that accounted for $92 \%$ of the total variability in the long-term time series of monthly means of air temperature and $80 \%$ for bottom water temperatures.

These models were then used to forecast conditions in 1999 with results showing good agreement between the predicted and observed values of both air and bottom water temperatures. Intervention analysis that models events as step-like features was also carried out. While this provided a better model fit to the observed data series, such events cannot be predicted. Since nearly all fitted interventions appeared during winter (December-March), prediction of temperatures during these months must be viewed with caution. Results showed that the use of ARIMA models yields better forecasts for highly variable time series than simple models based upon averages of previous monthly averages alone.

Hare, S.R. and R.C. Francis (1994) have studied the climate change and salmon production in the Northeast Pacific Ocean. Alaskan salmon stocks have exhibited enormous fluctuations in production during the century. They investigate their hypothesis that large-scale salmon-production variation was driven by the climatic processes in the Northeast Pacific Ocean. Using a time-series analytical techniques known as intervention analysis, they demonstrate that Alaskan salmonids alternate between high and low production regimes. The transition from high (low) regime to low (high) regime is called an intervention. To test for intervention, they first fitted the salmon time series to univariate autoregressive integrated moving average (ARIMA) models. On the basis of tentatively identified climatic regression, potential interventions were then identified and incorporated into the models, and the result was compared with the non-intervention models. A highly significant positive step intervention in the 1970s and a significant negative step intervention in the late 1940s were identified in the Alaska salmon stocks analyzed. We review the evidence for synchronous regime shift in
the 1940s and late 1970s that coincide with the shifts in salmon production. Potential mechanism in North Pacific climate processes to salmon production is identified.

Lu Guanghua, Tang Jie, Yuan Xing, and Zhao Yuanhui (2001) have studied Correlation for the structure and biodegradability of substituted benzenes in the Songhua river water. The biodegradability of 47 substituted benzenes was determined by BOD technique. The molecular weight $\left(M_{w}\right)$, the total surface area (TSA), the energy of the highest occupied molecular orbital ( $E_{\text {номо }}$ ), the heat of formation $\left(H_{f}\right)$, and the moment of dipole of 47 studied compounds were calculated by the quantum chemical method MOPAC6.0-AM1. The ionization constant $\left(p K_{a}\right)$ and n-octanol/water partition coefficient $(\log P)$ were obtained from Qsar software and Biobyte software, respectively. The quantitative structure-biodegradability relationship studies were developed by linear regression analysis. The correct prediction rate of obtained model is up to $85 \%$ for the testing set. It has been shown that the biodegradability of substituted benzenes in natural river water is mainly related to electronic parameter Еномо, $\mathrm{H}_{\mathrm{f}}$ and steric parameter $\mathrm{M}_{\mathrm{w}}$.
P.D. LaValle, V.C. Lakhan, and A.S. Trenhaile (2000) have studied the short term fluctuations of Lake Erie water levels and the $\mathrm{El} \mathrm{Nio} /$ Southern Oscillation. This study assesses the relationship between short term fluctuations of Lake Erie water levels and the El Niö/Southern Oscillation (ENSO) using data collected from May 1978 to May 1997. After standardizing the collected data, graphical and Box-Jenkins time series techniques are utilized to assess the temporal interrelationship of the Southern Oscillation Index and Lake Erie water level variables. The statistical results demonstrate that a first-order auto regressive model AR (1) provides the best fit for the data sets of
the analyzed variables. Both the graphical and statistical results suggest that short term Lake Erie water levels are fluctuating in response to the two ENSO phases, El Nio and La Nió. Negative values of the Southern Oscillation Index are related to higher lake levels while positive values are associated with lower lake levels.
N. Koning and JC Roos (1999) have studied the continued influence of organic pollution on the water quality of the turbid Modder River. The Modder River is a relatively small river, which drains an area of $7,960 \mathrm{~km} 2$, in the central region of the Free State Province, South Africa, and has a mean annual runoff of $184 \times 106 \mathrm{~m} 3$. Botshabelo is a city that was developed in the catchment area of the river and treated sewage is released into the Klein Modder River. This study determined seasonal and spatial patterns in the system as well as the continued influence that Botshabelo's treated sewage outflow has on the water quality of the river. Box-Jenkins plots were used in every sample point in this study, as well as, plots of the relations between different sample points. It was determined that the Modder River and Klein Modder River follow distinctive seasonal patterns in terms of algal growth and physical factors. There were periods when the waters of the Modder River and Klein Modder River are of acceptable quality. However, outflows from Botshabelo have a detrimental effect on the water quality in terms of nutrient concentrations and algal biomass. The inflow of the Klein Modder River into the Modder River caused an average of $112 \%$ increase in phosphate-phosphorus (PO4-P), $171 \%$ increase in nitrate-nitrogen (NO3-N) and a $50 \%$ increase in chlorophyll- $a$ concentration in the Modder River. The long-term detrimental effect of Botshabelo on the system can clearly be seen by comparing predicted nutrient increases with measured values.

## 2. KING TALAL DAM AND DATA COLLECTION

This chapter is divided into four parts. The first part is about the philosophy of King Talal Dam construction. The second part is about the water supplier of King Talal Dam. The third part is about water characteristics, and the last part is about data collection.

### 2.1 Philosophy of King Talal Dam Construction.

Since water is the main factor in agricultural production, and the main factor for extension in lands reclamation and since water resources are very limited in Jordan because of very low average of rainfall, it was necessary to think about ways of collecting surface water by dams to be constructed downstream rivers and valleys to be distributed to farms when required. In 1977 it was agreed to construct King Talal Dam as the first trial on Zarka River with a capacity of 56 MCM, which was increased to 86 MCM in 1988, 78 MCM of which were used for irrigation purposes (Salameh, 1996). The height of the dam was 108 m , the catchment area was $3157 \mathrm{Km}^{2}$ with a population of 2.439 million Capita (1996). This dam is considered the most important water project in Jordan. The water stored in it is used for irrigation of wide areas in the Ghor, the area irrigated by water coming from the dam mixed with water coming by King Abdullah channel is estimated by 100,000 donums. (RSS, 1988-2001).

Figure (1) shows the site of King Talal Dam, the borders of the catchment area and the branches that supply the lake behind the dam. The main branches

Figure (1): King Talal's catchment area and branches. (RSS reports. 1988 till 2001)

are Al Zarka River (main supplier), Al Dhlale Valley that pours in Al Zarka River, and Rumaimeen Valley that pours directly in the lake. The percent of water drained from Al Zarka River to the lake of the river is about $90 \%$, while for Rumaimeen Valley it is about $10 \%$. Population of the catchment area is about half of the kingdom population, which is distributed in Amman, Zarka, Sweileh, Rusaifeh, Baka'a and Jerash with several villages around these cities. (RSS, 1988-2001).

### 2.2 Water Suppliers for King Talal Reservoir:

The only two water suppliers for King Talal reservoir are Zarka River and Wadi Rumaimeen.

### 2.2.1 Zarka River

It is a river coming from the east directly to the lake of the dam (see figure 2 ). Zarka River is the main water branch of King Talal Reservoir. The water in Zarka River consists of the water in Wadi Dhuliel, the effluent from Al Samra, and Jerash Wastewater Treatment plant (see Figure 2). In addition to the treated and untreated effluents, which are drained from industrial factories and farms located on the river's banks. The underground water and rainfall affect water quality of this River.

### 2.2.2 Wadi Rumaimeen

It is a wadi coming from the south directly to the lake of the dam (see figure 2 ). This wadi was connected in 1988 to increase the capacity of the lake and to study the effect of AL Baka'a waste water plant on the quality of the

Figure (2): King Talal's reservoir inlets and outlets. (RSS reports. 1988 till 2001)
water in this valley and finally on the water in the lake and the springs water extending around this wadi that supply it with extra water.

### 2.3 Water Characteristics

There are three characteristics of water, these characteristics are: physical, chemical, and biological. According to these three characteristics, one can determine the usage of water. Water usage is generally divided into four categories: (1) domestic (water used for sanitary and general purposes, (2) industrial (no domestic purposes), (3) public service (water used for irrigation, fire fighting, and industrial system), (4) unaccounted for system losses and leakage. In this section, the water characteristics will be discussed for each of the five properties.

### 2.3.1 Total suspended solids (TSS):

TSS has physical characteristics; the term of the total suspended solids refers to non-filterable residue that is retained on a glass-fiber disk after filtration of a sample of water or wastewater. A measured portion of a sample is drown through a glass-fiber filter, retained in a funnel, applying a vacuum to the section flask under the filter with dump suspended solids adhering to the surface is transferred from the filtration apparatus to an aluminum or stainless steel planked as a support. After drying at $103^{\circ}$ $105^{\circ} \mathrm{C}$ in an oven the filtered dried suspended solids in milligrams divided by the volume of the sample by liters gives the total suspended solids expressed by milligrams per liters. (Viessman, 1985).

The total suspended solids fraction consists of the particulate matter with an approximate size range from 0.001 to $1 \mu \mathrm{~m}$. The dissolved solids consist of both organic
and inorganic molecules and ions that are present in true solution in water (Metcalf and eddy, 1991).

### 2.3.2 Biochemical oxygen demand (BOD):

BOD has chemical characteristics. Biochemical oxygen demand is the quantity of oxygen used in the aerobic stabilization of wastewater and polluted waters. The standard 5-days BOD value is commonly used to measure the amount of pollution in wastewater, to evaluate the efficiency of treatment by measuring oxygen demand remaining in the effluent, and to determine the amount of organic pollutant in surface waters (Viessman, 1985).

Laboratory analyses of wastewaters and polluted waters are considered using 300 ml of BOD bottle incubated at a room temperature of $20^{\circ} \mathrm{C}$. a measured portion is placed in the BOD bottle, then the bottle is filled with aerated dilution water containing phosphate buffer and inorganic nutrients. The sample is then diluted with distilled water, when the sample contains a large population of microorganisms (untreated wastewater, for example) seeding is not necessary. If required, the dilution water is seeded with a bacterial culture that has been acclimated with a bacterial matter or other materials that may be present in the wastewater. Readings are taken for five days or more, the biochemical oxygen demand exerted by a diluted wastewater progress approximately by first-order kinetics. Within 5 days, the oxidation of the carbonaceous organic matter is about 60-70 percent completion. (Metcalf and eddy, 1991).

It should be noted that the initial depletion of dissolved oxygen is the result of carbonaceous oxygen demand resulting from organic matter degradation. If presented in sufficient numbers, nitrifying bacteria exerts a secondary oxygen demand by lags several days behind the start of carbonaceous oxygen demand. (Viessman, 1985).

### 2.3.3 Chemical oxygen demand (COD):

COD has chemical characteristics, the chemical oxygen demand test is used to measure the content of organic matter of both wastewater and natural matter. The oxygen equivalent to the organic matter that can be oxidized is measured by using a strong chemical oxidized agent in an acidic medium. The test must be performed at an elevated temperature. A catalyst (silver sulfate) is required to aid the oxidation of certain classes of organic compounds. The COD test is also used to measure the organic matter in industrial and municipal wastes that contain compounds that are toxic to biological life. The value of COD is higher than BOD and that is because there are more compounds that can be chemically oxidized than biologically. For many types of wastes, it can be correlated COD with BOD, and this is useful because the value of COD can be determined in 3 hours, where as the value of BOD can be determined in 5 days. (Metcalf and Eddy, 1991).

### 2.3.4 Total phosphorus (T-P):

T-P has chemical characteristics. It is essential for the growth of algae and other biological organisms and that is because of noxious algal blooms that occur in surface waters, so there is presently much interest in controlling the amount of phosphorus compounds that enter surface waters in domestic and industrial waste discharges of natural run off. (Metcalf and Eddy, 1991)

### 2.3.5 Total nitrogen (T-N):

T-N has chemical characteristics, and it is essential for the growth of the biological characteristics; such as: Protista and plants. This happens because nitrogen is an essential factor in building the synthesis of protein.

In nature the nitrogen is presented in several ways, in water and wastewater, it is combined in proteinacous matter and urea. Nitrogen data will be required to evaluate the treatability of wastewater by biological processes, also the amount of nutrients should be controlled because of the algal growth. (Metcalf and eddy, 1991).

### 2.4 Data Collection:

The data collected in this thesis was collected from reports made by Jordan Valley Authority in cooperation with Royal Scientific Society. These reports contained monthly information covering the period from January 1988 till December 2000 (156 months) for two sights: Zarka River and Samra Waste water treatment plant. The information includes the monthly analysis of quality of the flow entering the dam through AL Zarka River and Samra Waste Water Treatment Plant effluent. The data contains the concentrations of total suspended solids in ( $\mathrm{mg} / \mathrm{l}$ ), biochemical oxygen demand in (mg/l), chemical oxygen demand ( $\mathrm{mg} / \mathrm{l}$ ), total phosphorus ( $\mathrm{mg} / \mathrm{l}$ ), total nitrogen ( $\mathrm{mg} / \mathrm{l}$ ), and finally the rate of flow (MCM/month).

Table 1: Data of water entering King Talal Reservoir from Zarka River

| YEAR | MONTH | Zarka River Flow MCM/month | TSS mg/l | $\begin{gathered} \text { BOD5 } \\ \text { mg/l } \end{gathered}$ | $\begin{aligned} & \text { COD } \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-N } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | January | 10.202 | 227.00 | 47.00 | 118.00 | 3.00 | 52.70 |
|  | February | 31.634 | 42.00 | 10.00 | 37.00 | 2.20 | 21.50 |
|  | March | 10.411 | 34.00 | 6.00 | 16.00 | 2.70 | 27.30 |
|  | April | 5.033 | 41.00 | 34.00 | 96.00 | 4.50 | 20.20 |
|  | May | 5.306 | 72.00 | 36.00 | 119.00 | 4.90 | 23.70 |
|  | June | 4.43 | 83.00 | 33.00 | 128.00 | 5.20 | 20.20 |
|  | July | 3.962 | 103.00 | 35.00 | 123.00 | 5.90 | 13.90 |
|  | August | 3.325 | 60.00 | 17.00 | 79.00 | 6.10 | 11.00 |
|  | September | 3.2 | 55.00 | 36.00 | 103.00 | 7.10 | 19.90 |
|  | October | 3.306 | 47.00 | 30.00 | 94.00 | 6.50 | 27.10 |
|  | November | 3.664 | 61.00 | 26.00 | 155.02 | 5.80 | 42.90 |
|  | December | 16.348 | 72.00 | 36.00 | 103.00 | 6.70 | 34.60 |
|  | Monthly Average | 8.402 | 72.34 | 24.40 | 76.91 | 4.15 | 27.64 |
|  | Total | 100.824 | 72.34 | 24.40 | 76.91 | 4.15 | 27.64 |
| 1989 | January | 7.247 | 83.00 | 12.00 | 51.00 | 3.10 | 23.21 |
|  | February | 5.101 | 27.00 | 11.50 | 63.99 | 5.05 | 30.02 |
|  | March | 6.825 | 30.00 | 22.00 | 58.50 | 5.70 | 35.75 |
|  | April | 4.829 | 36.50 | 14.50 | 64.50 | 5.67 | 36.93 |
|  | May | 3.87 | 59.00 | 37.00 | 127.51 | 6.50 | 33.57 |
|  | June | 2.878 | 55.00 | 25.00 | 107.50 | 5.89 | 20.21 |
|  | July | 3.298 | 60.49 | 25.00 | 103.98 | 8.10 | 27.10 |
|  | August | 3.15 | 54.01 | 18.50 | 97.51 | 6.54 | 22.13 |
|  | September | 2.967 | 52.00 | 16.00 | 72.00 | 11.80 | 28.24 |
|  | October | 2.946 | 64.99 | 16.00 | 87.49 | 8.36 | 32.96 |
|  | November | 3.555 | 74.99 | 18.00 | 82.49 | 8.97 | 45.46 |
|  | December | 4.953 | 45.50 | 21.00 | 91.50 | 8.60 | 51.94 |
|  | Monthly Average | 4.302 | 52.58 | 19.07 | 78.99 | 6.55 | 32.75 |
|  | Total | 51.619 | 52.58 | 19.07 | 78.99 | 6.55 | 32.75 |
| 1990 | January | 7.539 | 35.00 | 13.00 | 64.00 | 8.10 | 47.40 |
|  | February | 7.559 | 43.00 | 19.00 | 68.00 | 8.10 | 29.00 |
|  | March | 7.901 | 40.00 | 15.00 | 56.00 | 5.70 | 30.70 |
|  | April | 5.073 | 36.00 | 22.00 | 80.00 | 6.54 | 35.20 |
|  | May | 3.859 | 54.00 | 27.00 | 74.01 | 8.10 | 32.50 |
|  | June | 3.176 | 69.99 | 36.00 | 112.99 | 10.40 | 31.40 |
|  | July | 3.607 | 50.00 | 33.00 | 99.99 | 8.50 | 25.90 |
|  | August | 3.568 | 91.99 | 25.00 | 100.99 | 8.00 | 22.80 |
|  | September | 3.2 | 93.99 | 30.00 | 112.98 | 9.90 | 19.10 |
|  | October | 3.528 | 125.00 | 42.00 | 116.00 | 10.20 | 32.50 |
|  | November | 3.789 | 92.00 | 36.00 | 119.99 | 10.10 | 40.80 |
|  | December | 4.269 | 82.00 | 39.00 | 157.00 | 11.50 | 43.30 |
|  | Monthly Average | 4.756 | 60.75 | 25.33 | 89.46 | 8.39 | 33.51 |
|  | Total | 57.066 | 60.76 | 25.33 | 89.47 | 8.39 | 33.52 |
| 1991 | January | 8.355 | 79.00 | 47.00 | 132.00 | 10.50 | 39.40 |
|  | February | 6.941 | 20.00 | 29.00 | 125.00 | 8.13 | 36.10 |
|  | March | 8.773 | 47.00 | 11.00 | 54.00 | 3.70 | 35.00 |
|  | April | 4.061 | 51.00 | 28.00 | 78.00 | 7.45 | 29.00 |

Cont. Table 1:Data of water entering King Talal Reservoir from Zarka River.

| YEAR | MONTH | Zarka River Flow MCM/month | TSS mg/l | $\begin{gathered} \text { BOD5 } \\ \text { mg/l } \end{gathered}$ | $\begin{gathered} \text { COD } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 | May | 3.871 | 36.00 | 30.00 | 88.01 | 8.70 | 27.70 |
|  | June | 3.244 | 38.00 | 26.00 | 95.00 | 5.90 | 21.30 |
|  | July | 2.86 | 57.99 | 21.00 | 88.98 | 7.78 | 19.50 |
|  | August | 2.901 | 56.01 | 21.00 | 116.01 | 7.58 | 28.00 |
|  | September | 2.841 | 72.50 | 25.17 | 126.87 | 8.51 | 27.19 |
|  | October | 3.22 | 85.01 | 44.01 | 92.01 | 9.25 | 31.00 |
|  | November | 3.935 | 104.99 | 40.00 | 115.99 | 10.80 | 41.00 |
|  | December | 33.86 | 28.00 | 11.00 | 35.00 | 4.30 | 39.50 |
|  | Monthly Average | 7.073 | 45.36 | 22.06 | 74.23 | 6.30 | 35.27 |
|  | Total | 84.882 | 45.36 | 22.02 | 74.21 | 6.30 | 35.27 |
| 1992 | January | 24.598 | 59.00 | 23.00 | 51.00 | 3.40 | 17.00 |
|  | February | 68.544 | 65.00 | 19.00 | 59.00 | 2.05 | 19.00 |
|  | March | 29.245 | 71.00 | 15.00 | 67.00 | 0.70 | 21.00 |
|  | April | 15.179 | 59.00 | 21.00 | 60.00 | 3.70 | 28.80 |
|  | May | 10.266 | 89.00 | 25.00 | 73.00 | 4.60 | 27.00 |
|  | June | 8.693 | 118.00 | 28.00 | 85.00 | 5.50 | 25.20 |
|  | July | 6.909 | 102.99 | 33.00 | 85.99 | 6.40 | 21.25 |
|  | August | 6.596 | 86.98 | 36.99 | 86.98 | 7.30 | 17.30 |
|  | September | 6.187 | 90.00 | 24.00 | 93.00 | 6.84 | 26.20 |
|  | October | 6.537 | 92.99 | 11.00 | 98.99 | 6.37 | 35.00 |
|  | November | 8.206 | 75.00 | 13.00 | 77.00 | 6.29 | 26.80 |
|  | December | 14.78 | 56.00 | 15.00 | 54.00 | 6.20 | 18.50 |
|  | Monthly Average | 17.145 | 71.50 | 20.18 | 65.50 | 3.48 | 21.45 |
|  | Total | 205.737 | 71.50 | 20.15 | 65.51 | 3.48 | 21.45 |
| 1993 | January | 14.694 | 56.00 | 15.00 | 54.00 | 6.20 | 18.50 |
|  | February | 11.175 | 54.00 | 22.00 | 84.00 | 5.04 | 27.51 |
|  | March | 9.936 | 57.00 | 33.00 | 91.00 | 5.67 | 27.70 |
|  | April | 8.196 | 59.00 | 43.00 | 97.00 | 6.30 | 27.88 |
|  | May | 8.179 | 65.00 | 38.00 | 116.01 | 7.77 | 30.55 |
|  | June | 8.385 | 70.00 | 33.00 | 135.01 | 9.24 | 33.22 |
|  | July | 5.015 | 74.00 | 44.00 | 116.00 | 9.62 | 35.76 |
|  | August | 5.136 | 76.99 | 54.00 | 96.99 | 10.00 | 38.29 |
|  | September | 4.887 | 86.00 | 47.00 | 129.01 | 10.91 | 40.71 |
|  | October | 5.318 | 93.99 | 40.00 | 160.99 | 11.82 | 43.12 |
|  | November | 7.513 | 81.00 | 45.00 | 122.00 | 9.96 | 41.90 |
|  | December | 6.992 | 67.00 | 50.00 | 83.00 | 8.10 | 40.68 |
|  | Monthly Average | 7.952 | 66.64 | 35.21 | 100.35 | 7.79 | 31.59 |
|  | Total | 95.427 | 66.63 | 35.23 | 100.35 | 7.79 | 31.59 |
| 1994 | January | 9.827 | 47.00 | 45.00 | 86.00 | 7.16 | 38.60 |
|  | February | 8.346 | 26.00 | 39.00 | 88.00 | 6.21 | 36.52 |
|  | March | 8.184 | 60.00 | 28.00 | 90.00 | 6.10 | 40.62 |
|  | April | 5.201 | 37.00 | 58.00 | 127.00 | 6.81 | 37.51 |
|  | May | 4.561 | 63.99 | 42.00 | 127.99 | 12.61 | 49.55 |
|  | June | 4.017 | 67.00 | 52.00 | 123.00 | 13.96 | 48.20 |
|  | July | 3.344 | 82.99 | 60.99 | 120.98 | 13.37 | 48.81 |
|  | August | 3.803 | 80.00 | 63.00 | 149.01 | 12.99 | 47.46 |
|  | September | 3.395 | 93.99 | 47.00 | 115.99 | 14.71 | 50.20 |

Cont. Table 1:Data of water entering King Talal Reservoir from Zarka River

| YEAR | MONTH | Zarka River Flow MCM/month | $\begin{gathered} \text { TSS } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { BOD5 } \\ \text { mg/l } \end{gathered}$ | $\begin{aligned} & \text { COD } \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | October | 4.844 | 79.00 | 28.00 | 166.00 | 12.90 | 53.74 |
|  | November | 26.245 | 22.00 | 45.00 | 133.00 | 5.89 | 40.27 |
|  | December | 18.312 | 37.00 | 15.00 | 39.00 | 5.98 | 36.11 |
|  | Monthly Average | 8.34 | 44.58 | 38.90 | 104.20 | 7.91 | 41.21 |
|  | Total | 100.077 | 44.58 | 38.90 | 104.20 | 7.91 | 41.21 |
| 1996 | January | 5.989 | 62.95 | 92.00 | 101.02 | 8.01 | 52.10 |
|  | February | 8.115 | 60.01 | 32.04 | 83.06 | 7.02 | 37.09 |
|  | March | 7.693 | 50.96 | 41.99 | 88.00 | 8.71 | 47.06 |
|  | April | 6.653 | 41.94 | 55.01 | 94.99 | 8.27 | 47.35 |
|  | May | 6.568 | 33.95 | 61.97 | 103.99 | 10.66 | 50.24 |
|  | June | 5.197 | 72.93 | 50.03 | 113.91 | 11.55 | 47.91 |
|  | July | 5.07 | 111.05 | 54.04 | 140.04 | 13.41 | 69.23 |
|  | August | 5.117 | 86.97 | 45.93 | 154.97 | 14.27 | 39.48 |
|  | September | 4.785 | 84.01 | 51.62 | 174.92 | 3.34 | 50.37 |
|  | October | 4.766 | 71.97 | 45.95 | 147.08 | 12.38 | 58.54 |
|  | November | 5.165 | 70.09 | 18.97 | 143.08 | 11.62 | 49.76 |
|  | December | 6.413 | 79.06 | 36.96 | 104.01 | 10.92 | 55.98 |
|  | Monthly Average | 5.961 | 66.60 | 48.65 | 116.09 | 9.73 | 49.82 |
|  | Total | 71.532 | 66.54 | 48.64 | 116.17 | 9.80 | 49.75 |
| 1996 | January | 11.057 | 97.04 | 38.98 | 114.05 | 5.06 | 46.40 |
|  | February | 6.058 | 32.02 | 47.05 | 117.04 | 8.58 | 44.24 |
|  | March | 9.268 | 33.99 | 43.05 | 114.05 | 10.36 | 49.74 |
|  | April | 5.624 | 54.94 | 65.08 | 91.93 | 9.42 | 55.30 |
|  | May | 5.566 | 169.96 | 63.06 | 127.02 | 10.96 | 71.69 |
|  | June | 4.516 | 81.93 | 54.03 | 110.05 | 11.07 | 65.54 |
|  | July | 4.072 | 52.06 | 33.89 | 107.07 | 8.10 | 54.52 |
|  | August | 3.807 | 112.95 | 69.87 | 127.92 | 11.56 | 79.33 |
|  | September | 4.242 | 151.11 | 74.02 | 181.05 | 13.44 | 68.84 |
|  | October | 3.688 | 75.11 | 33.08 | 120.93 | 11.12 | 70.23 |
|  | November | 7.934 | 136.00 | 95.03 | 275.02 | 15.00 | 76.76 |
|  | December | 6.859 | 73.04 | 80.04 | 188.07 | 10.93 | 63.27 |
|  | Monthly Average | 6.058 | 87.32 | 57.94 | 142.46 | 10.07 | 60.09 |
|  | Total | 72.691 | 87.30 | 58.03 | 142.45 | 10.14 | 60.06 |
| 1997 | January | 18.864 | 127.01 | 125.00 | 231.98 | 12.72 | 66.05 |
|  | February | 11.551 | 70.04 | 41.04 | 72.03 | 8.57 | 43.98 |
|  | March | 8.071 | 187.96 | 85.00 | 223.02 | 8.67 | 55.88 |
|  | April | 5.275 | 73.93 | 68.06 | 115.07 | 6.26 | 37.54 |
|  | May | 5.45 | 55.96 | 44.04 | 79.08 | 7.89 | 48.26 |
|  | June | 4.826 | 98.01 | 77.08 | 156.03 | 9.12 | 63.41 |
|  | July | 4.863 | 92.95 | 56.96 | 104.05 | 9.05 | 49.15 |
|  | August | 4.839 | 75.02 | 36.99 | 93.20 | 11.57 | 61.38 |
|  | September | 4.674 | 65.90 | 40.01 | 90.07 | 12.84 | 57.55 |
|  | October | 4.609 | 70.95 | 31.89 | 116.08 | 14.10 | 62.05 |
|  | November | 6.854 | 106.07 | 46.98 | 147.07 | 10.50 | 60.84 |
|  | December | 10.711 | 43.97 | 135.00 | 302.96 | 12.32 | 52.56 |
|  | Monthly Average | 7.549 | 94.32 | 77.76 | 166.38 | 10.60 | 55.64 |
|  | Total | 90.587 | 94.26 | 77.78 | 166.39 | 10.56 | 55.65 |

Cont. Table 1:Data of water entering King Talal Reservoir from Zarka River

| YEAR | MONTH | Zarka River Flow MCM/month | TSS mg/l | $\begin{gathered} \text { BOD5 } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { COD } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-N } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1998 | January | 10.245 | 52.03 | 29.97 | 68.03 | 7.81 | 53.29 |
|  | February | 6.073 | 53.02 | 30.96 | 126.96 | 9.22 | 45.94 |
|  | March | 9.734 | 66.98 | 36.98 | 99.96 | 8.32 | 51.78 |
|  | April | 5.943 | 37.02 | 25.07 | 88.00 | 11.44 | 47.96 |
|  | May | 3.661 | 39.06 | 31.96 | 93.96 | 8.47 | 57.36 |
|  | June | 3.45 | 6.96 | 57.97 | 184.93 | 9.86 | 53.62 |
|  | July | 3.777 | 96.90 | 45.01 | 88.96 | 12.71 | 54.28 |
|  | August | 4.079 | 71.10 | 57.12 | 103.95 | 10.05 | 46.58 |
|  | September | 4.394 | 68.96 | 38.92 | 94.90 | 9.33 | 53.03 |
|  | October | 4.518 | 71.93 | 59.10 | 106.91 | 10.18 | 57.55 |
|  | November | 4.428 | 63.01 | 51.94 | 111.11 | 9.49 | 65.72 |
|  | December | 4.5 | 48.00 | 47.11 | 152.00 | 11.78 | 76.89 |
|  | Monthly Average | 5.4 | 56.67 | 40.19 | 104.63 | 9.63 | 54.63 |
|  | Total | 64.802 | 56.68 | 40.18 | 104.67 | 9.58 | 54.54 |
| 1999 | January | 7.318 | 40.04 | 40.99 | 203.06 | 12.16 | 66.41 |
|  | February | 9.728 | 39.99 | 24.98 | 86.04 | 8.53 | 66.51 |
|  | March | 5.202 | 33.06 | 62.09 | 104.00 | 8.84 | 68.63 |
|  | April | 4.201 | 74.98 | 64.03 | 194.00 | 12.62 | 77.60 |
|  | May | 4.798 | 67.94 | 61.90 | 255.94 | 12.71 | 66.07 |
|  | June | 3.961 | 89.88 | 77.00 | 268.11 | 14.64 | 62.11 |
|  | July | 3.14 | 80.89 | 42.04 | 230.89 | 11.46 | 50.32 |
|  | August | 3.317 | 106.12 | 56.07 | 177.87 | 9.04 | 61.80 |
|  | September | 3.529 | 79.91 | 30.04 | 207.14 | 14.45 | 54.41 |
|  | October | 4.37 | 97.03 | 30.89 | 102.97 | 13.73 | 59.50 |
|  | November | 4.68 | 64.10 | 35.90 | 176.07 | 16.03 | 72.22 |
|  | December | 5.551 | -- | 43.06 | 206.99 | 21.80 | 79.45 |
|  | Monthly Average | 4.983 | 63.21 | 45.15 | 174.59 | 12.84 | 66.43 |
|  | Total | 59.795 | 63.18 | 45.20 | 174.55 | 12.76 | 66.43 |
| 2000 | January | 12.228 | 44.00 | 35.98 | 180.98 | 12.43 | 63.13 |
|  | February | 6.552 | 58.00 | 39.07 | 219.02 | 8.55 | 67.00 |
|  | March | 7.031 | 33.00 | 22.05 | 164.98 | 9.53 | 62.58 |
|  | April | 4.31 | 54.06 | 57.08 | 229.93 | 12.76 | 77.49 |
|  | May | 4.226 | 78.09 | 57.97 | 136.06 | 12.07 | 79.51 |
|  | June | 3.971 | 67.99 | 45.08 | 184.08 | 11.08 | 57.92 |
|  | July | 3.651 | 87.92 | 30.13 | 205.97 | 19.99 | 64.64 |
|  | August | 3.275 | 105.95 | 37.86 | 258.93 | 15.57 | 59.85 |
|  | September | 3.355 | 85.84 | 34.87 | 298.96 | 12.22 | 64.68 |
|  | October | 7.113 | 85.06 | 58.06 | 162.94 | 13.07 | 73.67 |
|  | November | 4.384 | 57.03 | 18.93 | 135.95 | 14.14 | 69.34 |
|  | December | 8.792 | 68.02 | 40.04 | 226.00 | 20.02 | 79.85 |
|  | Monthly Average | 5.741 | 63.75 | 39.54 | 195.26 | 13.41 | 68.63 |
|  | Total | 68.888 | 63.76 | 39.47 | 195.26 | 13.36 | 68.66 |

Table 2: Data of Samra Wastewater Treatment Plant Effluent

| YEAR | MONTH | Samra WWTP flow MCM/month | TSS mg/l | $\begin{gathered} \text { BOD5 } \\ \text { mg/l } \end{gathered}$ | $\begin{gathered} \text { COD } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | January | 1.88 | 184.0 | 127.0 | 338.9 | 13.0 | 123.2 |
|  | February | 2.30 | 168.0 | 144.0 | 391.1 | 13.0 | 99.2 |
|  | March | 2.18 | 157.0 | 126.0 | 343.0 | 13.0 | 84.2 |
|  | April | 1.79 | 180.0 | 142.0 | 366.0 | 18.0 | 76.2 |
|  | May | 1.68 | 204.1 | 173.0 | 351.1 | 18.5 | 99.2 |
|  | June | 1.87 | 210.0 | 114.0 | 372.0 | 8.0 | 95.2 |
|  | July | 2.32 | 190.0 | 112.0 | 348.9 | 19.0 | 68.2 |
|  | August | 2.07 | 209.0 | 126.0 | 339.1 | 15.0 | 79.2 |
|  | September | 2.15 | 195.0 | 157.0 | 355.9 | 17.9 | 68.2 |
|  | October | 2.17 | 191.0 | 162.0 | 358.0 | 16.5 | 89.3 |
|  | November | 1.99 | 193.0 | 132.0 | 298.0 | 17.4 | 79.2 |
|  | December | 2.50 | 195.0 | 135.0 | 354.9 | 17.0 | 88.2 |
|  | Monthly Average | 2.07 | 189.2 | 137.1 | 351.8 | 15.6 | 86.9 |
|  | Total | 24.88 | 2270.3 | 1644.8 | 4221.1 | 186.8 | 1043.4 |
| 1989 | January | 2.27 | 126.0 | 102.0 | 344.0 | 17.3 | 95.2 |
|  | February | 1.96 | 105.0 | 124.0 | 420.1 | 17.3 | 102.2 |
|  | March | 2.33 | 145.0 | 145.0 | 455.0 | 17.6 | 106.2 |
|  | April | 2.23 | 172.0 | 118.0 | 340.0 | 17.1 | 98.2 |
|  | May | 2.20 | 196.0 | 62.0 | 324.0 | 18.4 | 98.2 |
|  | June | 2.06 | 137.0 | 31.0 | 275.0 | 18.9 | 90.5 |
|  | July | 2.26 | 130.0 | 69.0 | 276.0 | 16.0 | 93.9 |
|  | August | 2.54 | 144.0 | 64.0 | 263.0 | 13.8 | 95.5 |
|  | September | 2.08 | 170.0 | 93.1 | 293.0 | 16.0 | 91.6 |
|  | October | 2.19 | 203.0 | 64.0 | 350.0 | 13.0 | 103.1 |
|  | November | 2.25 | 221.0 | 107.0 | 335.0 | 21.3 | 109.3 |
|  | December | 2.40 | 172.0 | 95.0 | 332.0 | 8.6 | 112.0 |
|  | Monthly Average | 2.23 | 160.4 | 89.5 | 333.2 | 16.2 | 99.8 |
|  | Total | 26.76 | 1924.9 | 1074.4 | 3998.5 | 194.2 | 1197.4 |
| 1990 | January | 2.45 | 137.0 | 123.0 | 346.0 | 15.0 | 108.0 |
|  | February | 2.29 | 109.0 | 154.0 | 375.0 | 12.0 | 100.0 |
|  | March | 2.39 | 124.0 | 134.0 | 361.0 | 15.8 | 89.0 |
|  | April | 2.81 | 160.0 | 110.0 | 345.0 | 17.2 | 94.0 |
|  | May | 2.46 | 189.0 | 52.0 | 230.0 | 19.2 | 92.0 |
|  | June | 2.41 | 173.0 | 98.0 | 289.0 | 19.0 | 93.5 |
|  | July | 2.74 | 147.0 | 74.0 | 282.0 | 16.4 | 102.0 |
|  | August | 2.46 | 203.0 | 88.0 | 348.0 | 18.5 | 86.8 |
|  | September | 2.41 | 198.0 | 94.0 | 256.0 | 22.6 | 98.8 |
|  | October | 2.40 | 211.0 | 112.0 | 255.0 | 22.8 | 97.1 |
|  | November | 2.47 | 242.0 | 111.0 | 352.0 | 20.6 | 102.0 |
|  | December | 2.36 | 232.0 | 112.0 | 416.0 | 24.0 | 111.0 |
|  | Monthly Average | 2.47 | 176.9 | 104.6 | 320.7 | 18.6 | 97.8 |
|  | Total | 29.64 | 2122.4 | 1254.8 | 3848.7 | 222.8 | 1173.7 |
| 1991 | January | 2.67 | 246.8 | 146.8 | 412.3 | 32.8 | 123.1 |
|  | February | 2.46 | 56.5 | 81.9 | 353.2 | 23.0 | 102.0 |
|  | March | 2.70 | 152.8 | 35.8 | 175.5 | 12.0 | 113.8 |

Cont. Table 2: Data of Samra Wastewater Treatment Plant Effluent

| YEAR | MONTH | Samra WWTP flow MCM/month | $\begin{gathered} \text { TSS } \\ \text { mg/l } \end{gathered}$ | $\begin{aligned} & \text { BOD5 } \\ & \text { mg/l } \end{aligned}$ | $\begin{gathered} \text { COD } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathbf{l} \end{gathered}$ | $\begin{gathered} \text { T-N } \\ \text { mg/l } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 | April | 2.36 | 87.8 | 48.2 | 134.3 | 12.8 | 49.9 |
|  | May | 2.50 | 55.7 | 46.4 | 136.0 | 13.4 | 42.8 |
|  | June | 2.64 | 46.8 | 32.0 | 116.9 | 7.3 | 26.2 |
|  | July | 2.41 | 68.7 | 24.9 | 105.5 | 9.2 | 23.1 |
|  | August | 2.53 | 64.3 | 24.1 | 133.2 | 8.7 | 32.2 |
|  | September | 2.43 | 84.7 | 29.4 | 148.2 | 9.9 | 31.8 |
|  | October | 2.40 | 114.2 | 59.1 | 123.6 | 12.4 | 41.7 |
|  | November | 2.39 | 172.6 | 65.8 | 190.7 | 17.8 | 67.4 |
|  | December | 2.85 | 333.2 | 130.9 | 416.4 | 51.2 | 470.0 |
|  | Monthly Average | 2.53 | 126.9 | 61.6 | 207.7 | 18.0 | 98.7 |
|  | Total | 30.34 | 1523.1 | 739.2 | 2491.8 | 216.0 | 1184.3 |
| 1992 | January | 3.24 | 70.0 | 134.0 | 323.0 | 14.9 | 70.0 |
|  | February | 3.15 | 78.0 | 149.0 | 328.0 | 13.0 | 82.1 |
|  | March | 3.29 | 73.0 | 177.0 | 289.0 | 13.9 | 51.0 |
|  | April | 3.26 | 144.0 | 117.0 | 367.0 | 19.6 | 84.4 |
|  | May | 3.35 | 218.0 | 123.0 | 424.0 | 21.0 | 87.2 |
|  | June | 3.21 | 221.0 | 106.0 | 391.0 | 21.3 | 76.8 |
|  | July | 3.36 | 205.0 | 99.0 | 338.0 | 20.9 | 73.1 |
|  | August | 3.35 | 169.0 | 83.0 | 291.0 | 17.9 | 77.1 |
|  | September | 3.25 | 146.0 | 84.0 | 250.0 | 19.0 | 75.2 |
|  | October | 3.40 | 177.0 | 87.0 | 303.0 | 18.6 | 81.5 |
|  | November | 3.16 | 167.0 | 135.0 | 320.0 | 17.8 | 78.0 |
|  | December | 3.36 | 135.0 | 112.0 | 372.0 | 16.9 | 88.5 |
|  | Monthly Average | 3.28 | 150.2 | 117.2 | 333.0 | 17.9 | 77.1 |
|  | Total | 39.38 | 1803.0 | 1406.0 | 3996.0 | 214.8 | 924.9 |
| 1993 | January | 3.45 | 98.0 | 169.0 | 404.0 | 17.1 | 80.0 |
|  | February | 3.10 | 78.0 | 200.0 | 448.0 | 15.0 | 79.0 |
|  | March | 3.41 | 91.0 | 248.0 | 501.0 | 18.6 | 87.0 |
|  | April | 3.26 | 140.0 | 188.0 | 500.0 | 21.3 | 93.6 |
|  | May | 3.47 | 187.0 | 134.0 | 440.0 | 23.0 | 93.4 |
|  | June | 3.23 | 182.0 | 92.0 | 339.0 | 22.6 | 80.6 |
|  | July | 3.34 | 183.0 | 118.0 | 325.0 | 21.0 | 75.4 |
|  | August | 3.36 | 167.0 | 117.0 | 341.0 | 23.7 | 75.6 |
|  | September | 3.26 | 267.0 | 106.0 | 288.0 | 19.6 | 73.4 |
|  | October | 3.38 | 174.0 | 96.0 | 324.0 | 19.9 | 75.0 |
|  | November | 3.32 | 146.0 | 133.0 | 313.0 | 18.8 | 79.0 |
|  | December | 3.46 | 107.0 | 134.0 | 398.0 | 20.0 | 90.0 |
|  | Monthly Average | 3.34 | 151.7 | 144.6 | 385.1 | 20.0 | 81.8 |
|  | Total | 40.03 | 1820.0 | 1735.0 | 4621.0 | 240.6 | 982.0 |
| 1994 | January | 3.39 | 79.0 | 177.0 | 428.0 | 20.7 | 94.4 |
|  | February | 2.93 | 66.0 | 162.0 | 423.0 | 17.8 | 90.0 |
|  | March | 3.19 | 111.0 | 102.0 | 398.0 | 18.8 | 76.4 |
|  | April | 3.07 | 213.0 | 117.0 | 390.0 | 19.7 | 78.1 |
|  | May | 3.18 | 164.0 | 85.0 | 398.0 | 20.5 | 77.2 |
|  | June | 3.07 | 193.0 | 90.0 | 331.0 | 21.3 | 80.6 |
|  | July | 3.42 | 169.0 | 41.0 | 311.0 | 20.7 | 79.1 |

Cont. Table 2: Data of Samra Wastewater Treatment Plant Effluent

| YEAR | MONTH | Samra WWTP flow MCM/month | $\begin{gathered} \text { TSS } \\ \text { mg/l } \end{gathered}$ | $\begin{aligned} & \text { BOD5 } \\ & \text { mg/l } \end{aligned}$ | $\begin{gathered} \text { COD } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | T-P mg/l | $\begin{aligned} & \mathrm{T}-\mathrm{N} \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1994 | August | 3.41 | 144.0 | 40.0 | 324.0 | 20.0 | 71.1 |
|  | September | 3.50 | 143.0 | 68.0 | 289.0 | 19.8 | 76.5 |
|  | October | 3.73 | 149.0 | 39.0 | 379.0 | 19.6 | 76.8 |
|  | November | 3.69 | 134.0 | 93.0 | 410.0 | 17.6 | 76.4 |
|  | December | 4.02 | 103.0 | 117.0 | 292.0 | 15.5 | 76.1 |
|  | Monthly Average | 3.38 | 139.0 | 94.2 | 364.4 | 19.3 | 79.4 |
|  | Total | 40.61 | 1668.0 | 1131.0 | 4373.0 | 232.0 | 952.7 |
| 1995 | January | 4.11 | 62.1 | 134.9 | 322.9 | 15.6 | 81.1 |
|  | February | 3.70 | 68.9 | 104.9 | 410.1 | 15.7 | 93.0 |
|  | March | 4.13 | 98.1 | 194.9 | 409.9 | 17.9 | 90.1 |
|  | April | 3.83 | 136.0 | 120.1 | 402.9 | 19.6 | 89.0 |
|  | May | 3.67 | 157.0 | 107.9 | 340.1 | 19.1 | 85.0 |
|  | June | 3.49 | 158.1 | 56.0 | 447.1 | 18.4 | 39.3 |
|  | July | 3.43 | 153.1 | 44.9 | 313.1 | 19.2 | 79.0 |
|  | August | 3.44 | 167.0 | 63.9 | 340.0 | 20.0 | 78.1 |
|  | September | 3.47 | 150.0 | 45.9 | 242.9 | 19.6 | 72.1 |
|  | October | 3.45 | 145.9 | 67.0 | 291.0 | 18.9 | 82.1 |
|  | November | 3.58 | 139.0 | 55.9 | 360.0 | 17.9 | 86.4 |
|  | December | 3.58 | 75.1 | 137.9 | 415.1 | 17.0 | 89.1 |
|  | Monthly Average | 3.66 | 125.9 | 94.5 | 357.9 | 18.2 | 80.4 |
|  | Total | 43.86 | 1510.3 | 1134.2 | 4295.2 | 218.9 | 964.4 |
| 1996 | January | 4.27 | 70.1 | 150.0 | 462.0 | 15.5 | 99.1 |
|  | February | 3.98 | 72.1 | 193.1 | 423.0 | 13.8 | 92.9 |
|  | March | 4.29 | 74.1 | 161.9 | 468.1 | 16.3 | 98.1 |
|  | April | 3.99 | 79.0 | 193.1 | 522.9 | 18.6 | 103.1 |
|  | May | 3.83 | 109.9 | 216.1 | 539.9 | 18.8 | 108.0 |
|  | June | 3.64 | 120.0 | 132.9 | 414.1 | 18.9 | 112.0 |
|  | July | 3.59 | 145.0 | 54.0 | 341.0 | 18.4 | 106.1 |
|  | August | 3.61 | 140.9 | 140.0 | 424.0 | 17.5 | 100.1 |
|  | September | 3.62 | 164.0 | 138.0 | 401.2 | 18.8 | 104.1 |
|  | October | 4.29 | 194.1 | 103.9 | 364.0 | 20.0 | 107.9 |
|  | November | 3.73 | 106.9 | 266.1 | 694.0 | 20.6 | 110.9 |
|  | December | 3.75 | 102.9 | 330.9 | 790.1 | 21.1 | 113.1 |
|  | Monthly Average | 3.88 | 114.9 | 173.3 | 487.0 | 18.2 | 104.6 |
|  | Total | 46.60 | 1379.0 | 2080.1 | 5844.5 | 218.2 | 1255.4 |
| 1997 | January | 3.72 | 243.1 | 182.0 | 583.8 | 19.1 | 109.0 |
|  | February | 3.38 | 197.0 | 365.9 | 621.9 | 16.0 | 104.0 |
|  | March | 3.74 | 96.0 | 249.9 | 604.9 | 16.0 | 102.1 |
|  | April | 3.14 | 114.0 | 308.1 | 584.9 | 17.8 | 107.0 |
|  | May | 3.87 | 129.9 | 205.0 | 492.9 | 18.1 | 112.1 |
|  | June | 3.79 | 185.1 | 156.1 | 452.1 | 19.0 | 101.9 |
|  | July | 4.77 | 139.0 | 88.0 | 250.1 | 17.0 | 70.0 |
|  | August | 4.28 | 140.0 | 114.1 | 225.1 | 18.0 | 76.0 |
|  | September | 4.02 | 153.0 | 98.0 | 227.1 | 19.4 | 97.0 |
|  | October | 4.27 | 144.9 | 82.1 | 257.9 | 20.4 | 80.9 |
|  | November | 4.17 | 94.1 | 79.9 | 271.0 | 20.6 | 71.1 |

Cont. Table 2: Data of Samra Wastewater Treatment Plant Effluent

| YEAR | MONTH | $\begin{aligned} & \text { Samra WWTP } \\ & \text { flow } \\ & \text { MCM/month } \end{aligned}$ | TSS mg/l | $\begin{aligned} & \text { BOD5 } \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | $\begin{aligned} & \text { COD } \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | T-N mg/l |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1997 | December | 4.70 | 64.0 | 186.1 | 265.0 | 18.9 | 90.0 |
|  | Monthly Average | 3.99 | 141.7 | 176.3 | 403.1 | 18.4 | 93.4 |
|  | Total | 47.83 | 1700.1 | 2115.2 | 4836.7 | 220.4 | 1121.1 |
| 1998 | January | 4.87 | 61.0 | 229.0 | 324.9 | 17.9 | 94.7 |
|  | February | 3.83 | 77.1 | 176.9 | 355.9 | 18.6 | 98.0 |
|  | March | 4.72 | 111.9 | 177.0 | 382.0 | 19.3 | 109.4 |
|  | April | 4.30 | 78.0 | 149.9 | 274.0 | 18.4 | 93.4 |
|  | May | 4.58 | 105.9 | 161.0 | 318.0 | 19.0 | 93.5 |
|  | June | 4.46 | 121.0 | 138.0 | 285.9 | 22.9 | 99.0 |
|  | July | 4.63 | 128.1 | 94.0 | 252.9 | 21.0 | 93.1 |
|  | August | 4.36 | 145.0 | 89.0 | 208.1 | 17.2 | 128.9 |
|  | September | 4.92 | 134.1 | 64.0 | 329.9 | 27.0 | 130.0 |
|  | October | 4.80 | 140.9 | 89.1 | 305.0 | 14.4 | 95.5 |
|  | November | 4.40 | 119.0 | 67.9 | 265.9 | 14.1 | 100.1 |
|  | December | 4.65 | 94.9 | 111.9 | 384.9 | 19.8 | 107.0 |
|  | Monthly Average | 4.54 | 109.7 | 129.0 | 307.3 | 19.1 | 103.6 |
|  | Total | 54.51 | 1316.9 | 1547.8 | 3687.5 | 229.4 | 1242.6 |
| 1999 | January | 4.61 | 47.9 | 164.9 | 399.0 | 21.3 | 108.9 |
|  | February | 4.07 | 76.1 | 107.0 | 318.9 | 18.4 | 105.1 |
|  | March | 4.56 | 56.0 | 126.4 | 478.1 | 14.9 | 106.9 |
|  | April | 4.24 | 96.0 | 91.1 | 521.1 | 25.0 | 118.0 |
|  | May | 4.35 | 111.1 | 80.0 | 327.0 | 20.0 | 106.3 |
|  | June | 4.17 | 124.9 | 82.0 | 247.9 | 22.1 | 99.3 |
|  | July | 4.44 | 105.0 | 114.9 | 454.9 | 24.1 | 93.5 |
|  | August | 4.36 | 142.0 | 75.0 | 253.0 | 20.6 | 97.9 |
|  | September | 4.33 | 112.0 | 70.9 | 382.1 | 14.6 | 97.0 |
|  | October | 4.66 | 129.0 | 86.1 | 258.9 | 20.0 | 100.5 |
|  | November | 4.45 | 115.1 | 102.9 | 351.1 | 23.6 | 115.1 |
|  | December | 4.70 | 75.9 | 209.9 | 364.7 | 20.0 | 110.8 |
|  | Monthly Average | 4.41 | 99.3 | 109.3 | 363.1 | 20.4 | 104.9 |
|  | Total | 52.93 | 1191.1 | 1311.2 | 4356.8 | 244.5 | 1259.2 |
| 2000 | January | 4.99 | 67.0 | 121.9 | 427.1 | 16.2 | 103.1 |
|  | February | 4.22 | 54.9 | 114.1 | 339.0 | 18.0 | 91.1 |
|  | March | 4.55 | 49.0 | 152.1 | 386.1 | -- | -- |
|  | April | 4.66 | 117.1 | 110.0 | 349.9 | 21.1 | 116.2 |
|  | May | 4.71 | 111.1 | 139.1 | 499.0 | -- | -- |
|  | June | 4.40 | 138.1 | -- | 303.9 | 20.9 | 102.7 |
|  | July | 4.29 | 151.1 | 97.0 | 354.9 | -- | -- |
|  | August | 4.33 | 184.9 | 54.0 | 256.9 | 21.9 | 102.3 |
|  | September | 4.53 | 144.0 | 81.0 | 336.9 | -- | -- |
|  | October | 4.59 | 132.9 | 62.1 | 269.1 | 18.3 | 112.2 |
|  | November | 4.50 | 115.9 | 108.0 | 416.9 | -- | -- |
|  | December | 4.98 | 109.0 | 137.1 | 453.0 | 22.1 | 81.3 |
|  | Monthly Average | 4.56 | 114.6 | 106.9 | 366.1 | 19.8 | 101.3 |
|  | Total | 54.75 | 1375.0 | 1283.3 | 4393.0 | 237.4 | 1215.2 |

## 3. STATISTICAL ANALYSIS

This Chapter is divided into five sections. The first section about exploration of data, the second section about equations needed for forecasting, the third section about the statistical softwares, the fourth section about the methodology, and the fifth section is about the analysis of data for the variables in king Talal Reservoir and it's relation with AL Samra WWTP.

### 3.1. Exploration of Data:

The first step of time series analysis is to plot the data in a scatter diagram. Graphic presentation of statistical data gives a pictorial effect. The collected data will generally be complex; it will be very difficult to understand the importance of collected data. Graphic presentation makes the data easy to be understood and grasped. Also it shows if there is any trend that maybe present and the direction in which the trend may change. (Pillai and Bagavalthi, 1997).

The plot of time series contains fluctuations, which means that it is hard to bring order into seemingly hazard movement of the data through time. Nevertheless, in making simplifying assumptions, we can identify, explain and measure fluctuations that appear in the time series. The fluctuations that appear in the plot are due to four basic types of variations: secular trend, seasonal, cyclic and irregular variations (A. Sakakini, 2001). The variation in the data will lead to outlier. In general, an outlier is an observation that is far from the rest of the data.

### 3.2. Equations Needed in Forecasting:

Through applying mathematical models to the past data, one can analyze the data through mathematical models so that future forecasting of the data can be estimated. The mathematical models used in this thesis were:

### 3.2.1 Normality of data

### 3.2.1.1. Weibull distribution:

Is one of the most efficient formulas for computing plotting positions for unspecified distributions, the mathematical formula of Weibull Distribution is:

$$
\begin{equation*}
P=\frac{m}{n+1} \tag{1}
\end{equation*}
$$

Where:
P: The estimated probability of an X value
m: The Rank
n : The Number of years of record

The technique in all cases is to arrange the data in increasing or decreasing order of magnitude and to assign order number $m$ to the rank values. (Viessman and Lewis, 1996).

### 3.2.1.2. coefficient of variation process (COV)

The coefficient of variation is often used to describe the relative amount of variation in a population. The sample estimate for the coefficient of variation is defined as:

$$
\begin{equation*}
\mathrm{COV}=\frac{\mathrm{S}}{\frac{\mathrm{X}}{-}} \tag{2}
\end{equation*}
$$

Where:
COV: The coefficient of variation (dimensionless)

## S : Standard deviation <br> X : The mean

The coefficient of variable (COV) test is the simplest test available to determine whether the data can be characterized by normal distribution or not (McBean and Rovers, 1998).

### 3.2.1.3. Kurtosis process:

The fourth moment about the mean is the kurtosis, which is a measure of the peakness of the distribution. The sample estimate of the Kurtosis is obtained from the following equation:

$$
\begin{align*}
& K=\frac{n^{2} \sum_{i=1}^{n}\left(X_{i}-\overline{X^{2}}\right)}{(n-1)(n-2)(n-3)}  \tag{3}\\
& C_{k}=\frac{K}{S^{4}}-3 \tag{4}
\end{align*}
$$

Where:
$\mathrm{C}_{\mathrm{k}}$ : Coefficient of Kurtosis.

K : The fourth moment about the mean.
n : The number of data.
$\mathrm{X}_{\mathrm{i}}$ : Observed data.

X : Mean.

S : Standard deviation. (McBean and Rovers, 1998)

Kurtosis concerns the relative concentration of values in the center of the distribution as compared to the tails. In term of this property, three types of distributions can be defined: leptokurtic, mesokurtic and platykutric. Leptokurtic distribution is characterized by a prominent peak and by a relatively large proportion of values falling in the tail. A mesokurtic distribution is one, which the values are mainly located in the center of the distribution. A platykutric distribution is characterized that the peak is relatively flat and very few values appear in the tails (Levine, Ramsey and Smidt, 2001)

### 3.2.1.4. Shapiro-Wilk test

The Shapiro-Wilk "W' test is another statistical goodness of fit test that performs for a small sample (3-50). The steps for calculating W are:

## A- Order the sample data

B- Compute a weighted sum (b) of the differences between the most extreme observations

$$
\begin{equation*}
b=\sum_{i=1}^{n} a_{n-i+1}\left(X_{n-i+1}-X_{i}\right) \tag{5}
\end{equation*}
$$

C- Divide the weighted sum by a multiple of the standard deviation, and square the results to get the Shapiro-Wilk statistics Wi.

$$
\begin{equation*}
W=\left\{\frac{b^{2}}{S(n-1)^{0.5}}\right\} \tag{6}
\end{equation*}
$$

D- Compare the computed value of W with five percent critical value for a specific sample size. (McBean and Rovers, 1998).

### 3.2.2 Order of autoregressive

The value of autoregressive is indicated how the joint distribution of some P consecutive $y_{t}$ 's and the identical distribution of the (independently distributed) $u_{t}$ 's together with:

$$
\begin{equation*}
\mathrm{U}_{\mathrm{t}}=\sum_{\mathrm{r}=0}^{\mathrm{p}} \beta_{\mathrm{r}} \mathrm{y}_{\mathrm{t}-\mathrm{r}} \tag{7}
\end{equation*}
$$

Where:
$\mathrm{U}_{\mathrm{t}}$ : Random process with mean equals to zero and variance equals to one.
$\beta_{\mathrm{r}}$ : Unknown parameter with $\beta_{0}=1$ and $\left|\beta_{\mathrm{r}}\right|<1$.
$y_{t-r}$ : observation (data) at time equals to t-r.

For first order case, where $\mathrm{p}=1$

$$
\mathrm{U}_{\mathrm{t}}=\beta_{0} \mathrm{y}_{\mathrm{t}}+\beta_{1} \mathrm{y}_{\mathrm{t}-1}
$$

For $\beta_{0}=1$ and $\rho=\beta_{\mathrm{r}}$ the equation will be

$$
\begin{equation*}
y_{t}=\rho y_{t-1}+U_{t} \tag{8}
\end{equation*}
$$

For second order case, where $\mathrm{P}=2$, replacing $\mathrm{y}_{\mathrm{t}-1}$ in equation (7) and replacing t
with $t-1$ (that is, $y_{t-1}=\rho y_{t-2}+u_{t-1}$ ) we obtain:

$$
\begin{equation*}
y_{t}=U_{t}+\rho U_{t-1}+\rho^{2} y_{t-2} \tag{9}
\end{equation*}
$$

So in general:

$$
\begin{equation*}
y_{t}=U_{t}+\rho U_{t-1}+\ldots .+\rho^{p-1} U_{(t-(p-1))}+\rho^{p} y_{t-p} \tag{10}
\end{equation*}
$$

Let $\lambda$ be the lag operator, so the relation between the observation and the lag will be :

$$
\begin{equation*}
\lambda y_{t}=y_{t-1} \tag{11}
\end{equation*}
$$

So equation (7) can be written as:

$$
\begin{equation*}
\mathrm{U}_{\mathrm{t}}=\sum_{\mathrm{r}=0}^{\mathrm{p}} \beta_{\mathrm{rr}} \lambda^{\mathrm{r}} \mathrm{y}_{\mathrm{t}-\mathrm{r}} \tag{12}
\end{equation*}
$$

(Anderson, 1971)

### 3.2.3 The moving average model

The general formula for the moving average model is given by the equation:
$y_{t}=\sum_{j=0}^{q} \alpha_{q} v_{t-j}$

Where:
$y_{t}$ : Current observation.
$\alpha_{q}$ : unknown parameter, $\alpha_{0}=1$ and $\left|\alpha_{q}\right|<1$.
$v_{t}$ : is a sequence of independent random variables with $\xi v_{t}=0$ and $\xi v_{t}^{2}=\sigma^{2}$.
When adding the lag parameter $(\lambda)$ into equation (13), so the equation can be written as:

$$
\begin{equation*}
y_{t}=\sum_{j=0}^{q} \alpha_{q} \lambda^{j} v_{t-j} \tag{14}
\end{equation*}
$$

### 3.2.4 Autoregressive process with moving average residuals: -

This model is summarized by the following equation

$$
\begin{equation*}
\sum_{\mathrm{j}=0}^{\mathrm{q}} \alpha_{\mathrm{q}} \mathrm{v}_{\mathrm{t}-\mathrm{j}}=\sum_{\mathrm{r}=0}^{\mathrm{p}} \beta_{\mathrm{r}} \mathrm{y}_{\mathrm{t}-\mathrm{r}} \tag{15}
\end{equation*}
$$

Equation (14) is valued when lag parameter is not used, when using lag parameter the equation will be

$$
\sum_{\mathrm{j}=0}^{\mathrm{q}} \alpha_{\mathrm{q}} \lambda^{\mathrm{j}} \mathrm{v}_{\mathrm{t}-\mathrm{j}}=\sum_{\mathrm{r}=0}^{\mathrm{p}} \beta_{\mathrm{r}} \lambda^{\mathrm{r}} \mathrm{y}_{\mathrm{t}-\mathrm{r}}
$$

(16) (Anderson, 1971)

After analysis of equation (15), the final equation that is used in the ARMA modeling is:

$$
\begin{equation*}
Y_{t}=\rho_{1} Y_{t-1}+\ldots .+\rho_{\mathrm{p}} Y_{t-p}+U_{t}+\alpha_{q} U_{t}+\ldots .+\alpha_{q} U_{t-q} \tag{17}
\end{equation*}
$$

The residual sum of squares can be calculated at every point on a suitable grid of parameter values, and the values that give the minimum sum of squares may then be assessed. (Chatfield, 1984)

### 3.2.5 Integrated models:

In practice, most time series are non-stationary. In order to fit a stationary model, it is necessary to remove non-stationary sources of variation. If the observed time series is non-stationary in mean then we can difference the series. This is done by replacing $Y_{t}$ in equation (17) by $\quad{ }^{d}{ }_{X t}$. And representing or roplacing ${ }^{d}{ }_{X t}$ with $W_{t}$. So the general AZoregressive integrated moving average process (abbreviated ARIMA process) is of the form

$$
\begin{equation*}
\mathrm{W}_{\mathrm{t}}=\rho_{1} \mathrm{~W}_{\mathrm{t}-1}+\ldots .+\rho_{\mathrm{p}} \mathrm{~W}_{\mathrm{t}-\mathrm{p}}+\mathrm{U}_{\mathrm{t}}+\alpha_{\mathrm{q}} \mathrm{U}_{\mathrm{t}}+\ldots .+\alpha_{\mathrm{q}} \mathrm{U}_{\mathrm{t}-\mathrm{q}} \tag{18}
\end{equation*}
$$

The above ARIMA process describing the $d^{\text {th }}$ differences of the data is said to be of order (p, d, q). (Chatfield, 1984)

### 3.2.6 Linear regression

The mechanics of using the regression line to predict Y from X are simple enough once the line has been fitted through the scatter plot. The equation that defines the regression line is shown in the following equation:

$$
\begin{equation*}
Y^{\prime}=a+b_{y} X \tag{19}
\end{equation*}
$$

Where
$Y^{\prime}$ : The predicted value of $Y$
a : The regression constant
$b_{y}$ : The regression coefficient
X : Observed value

In words, the predicted value of the variable ( $\mathrm{Y}^{\prime}$ ) for any value of the variable X is computed by multiplying X by the regression coefficient $\left(\mathrm{b}_{\mathrm{y}}\right)$ and adding the regression constant (a). The equation that estimate the regression constant (a) and coefficient $\left(b_{y}\right)$ is called the least square equation. (Diekhoff, 1996).

### 3.2.7 Quadratic regression

If the relationship was clearly non-linear. In practice you may find it necessary to fit several types of curve to the data and choose the one, which gives the best fit, so if a quadratic relationship is thought to be the appropriate, the regression curve is given by:

$$
\begin{equation*}
Y=a_{0}+a_{1} X+a_{2} x^{2} \tag{20}
\end{equation*}
$$

Then the quantities $a_{0}, a_{1}$ and $a_{2}$ must be estimated from the data. A general method of estimating the parameter is by the method of least squares. (Chatfield, 1978).

### 3.2.8 Exponential growth regression

If the relationship was clearly non-linear. In practice you may find it necessary to fit several types of curve to the data and choose the one, which gives the best fit, so if an exponential growth relationship is thought to be the appropriate, the regression curve is given by:

$$
\begin{equation*}
\mathrm{Y}=\mathrm{a}_{0} \mathrm{e}^{\mathrm{x}} \tag{21}
\end{equation*}
$$

Then the quantity $\mathrm{a}_{0}$ must be estimated from the data. A general method of estimating the parameter is by the method of least squares. (Chatfield, 1978).

### 3.2.9 Single exponential smoothing

The model should only be used in its basic form for non-seasonal time series showing no trend, so if the data showed trend or seasonality then the first step to make single exponential smoothing regression is to eliminate this trend. The single exponential smoothing equation is written as :

$$
\begin{equation*}
\hat{\mathrm{Y}}_{(\mathrm{N}+1)}=\theta \hat{\mathrm{Y}}_{(\mathrm{N})}+\theta(1-\theta) \hat{\mathrm{Y}}_{(\mathrm{N}-1)}+\theta\left(1-\theta^{2}\right) \hat{\mathrm{Y}}_{(\mathrm{N}-2)}+\ldots \tag{22}
\end{equation*}
$$

Where

$$
\begin{array}{ll}
\hat{\mathrm{Y}}_{(\mathbb{N}+1)} & : \text { Estimated value } \\
\hat{\mathrm{Y}}_{(\mathrm{N})} \hat{\mathrm{Y}}_{(\mathrm{N}-1)} \hat{\mathrm{Y}}_{(\mathrm{N}-2)} & : \text { Past values } \\
\theta & : \text { Smoothing constant }, 0<\theta<1
\end{array}
$$

The value of the smoothing constant, $\theta$, depends on the properties of the given time series. Values between 0.1 and 0.3 is commonly used (Chatfield, 1984).

### 3.2.10 Testing differences between means

The mechanism of testing the difference between two means is assigning randomly the data to two samples, the first sample is the experimental group, which is affected with the treatment and the second one is the controlled group, which get nothing special. So the difference between the mean is the difference between the average of the first and second sample, mathematically:

$$
\begin{equation*}
\text { Difference in mean }=-X_{1}-X_{2} \tag{23}
\end{equation*}
$$

Where
$\overline{\mathrm{X}} 1=$ The mean of the experimental group.
$\bar{X}_{2}=$ The mean of the controlled group.
The difference in mean is a good test to see if one must accept or reject the null theory. (Chase, 1976).

### 3.2.11 Cross and distance correlation

For cross and distance correlation, the population correlation coefficient $\rho_{\mathrm{xy}}$ between the two random variables x and y is equal to:

$$
\begin{equation*}
\rho_{\mathrm{xy}}=\frac{\sigma_{\mathrm{xy}}}{\sigma_{\mathrm{x}} \sigma_{\mathrm{y}}} \tag{24}
\end{equation*}
$$

Where
$\sigma$ : is the variance

### 3.3 Statistical software:

Many software were established to satisfy statistical purposes, some of these software are: Minitab13, S-Plus 2000, SPSS, Statistica, NCSS 2001, Matlab, and many other software. In this thesis two packages were used: Minitab13 and S-Plus 2000.

Minitab statistical software is designed to help in solving statistical problems. Minitab is a general-purpose statistical system that can be run in either interactive or batch mode, and on mainframe computers, minicomputers or microcomputers. Minitab is easy to learn and easy to use. It is an ideal companion for the student or the worker faced with the need to analyze data and to formulate problems involving random problems. Minitab is a package used for simulation, regression, statistics identifications, fitting, checking, forecasting using ARIMA models, and other applications. (Miller, 1988).

The S-Plus programming provides an interactive computing environment for programming, graphics, and statistical analysis. A variety of data can be stored, manipulated, plotted, and analyzed using build up functions. S plus can easily accommodate and manipulate data stored in ways suitable for direct manipulation as vectors and matrices, as well as the "observations and variables" model used in most statistical packages. This flexibility allows performing a wide range of tasks. S-Plus can be used in statistical analysis, graphics tools, matrices and lists. (Spector, 1994).

### 3.4. Methodology:

### 3.4.1 Outliers and missing data

3.4.1.1 The missing data is determined and calculated by the average of the data of the same month.
3.4.1.2 A scatter diagram was plotted using Excel Charts; the diagram was plotted between the variable vs. the time.
3.4.1.3 A box diagram was plotted for the original data by seasonal period and Residuals by seasonal period using the software "Minitab 13"; the outliers are shown in those two diagrams as points.
3.4.1.4 The outliers ware determined in the two cases and checked if they were real outliers or not. The real outliers were not changed; the others were adjusted to the average monthly value.
3.4.1.5 A new scatter diagram was plotted for the new adjusted data using Excel software.

### 3.4.2 Normality: -

To determine if the data was normal or not in this thesis, four methods were used.

### 3.4.2.1. Weibull distribution model histogram:

A) The average monthly value for the variable was calculated.
B) Weibull Distribution Model Histogram was drawn using "Minitab 13".
C) The histogram is determined if skewed to left or right or not.

### 3.4.2.2. coefficient of variance $(\mathrm{COV})$ :

A) The data is divided into four quarters.
B) The coefficient of variable issues found for each quarter Excel sheet.
C) The value of the coefficient of variable is compared with [1], if it is less than
$|1|$ then the data is not skewed; otherwise it is skewed to left or right.

### 3.4.2.3. Kurtosis coefficient:

A) The data were divided into four quarters.
B) The Kurtosis Coefficient was found for each quarter using Excel sheet.
C) The value of the Kurtosis Coefficient was compared with $|1|$, if it is less than $|1|$ then the data is called Mesokurtic, if not, the data is called Leptokurtic for the positive sign and Platykutric for the negative sign.

### 3.4.2.4. Shapiro-Wilk test:

A) The data were divided into four quarters.
B) The Shapiro-Wilk Test was found for each quarter using Excel sheet.
C) The value of the Shapiro-Wilk Test was compared with the value in the Appendix (3), if it is greater than it, the data is normally distributed, if it is less than it, the data is skewed.

### 3.4.3 Order of (AR): -

The value of AR (p) is determined through drawing the Autocorrelation function for the variable, taking into consideration that the value of (p) should not exceed (1) in surface water forecasting, because the small rivers (as Zarka River) the water characteristics do not need more than few days to dilute, so the correlation does not
exceed one month (Viessman and Lewis, 1996). The figures of autocorrelation function is drawn using "Minitab 13" software.

### 3.4.4 Order of moving average: -

3.4.4.1 The moving average graph for the variables is drawn using Minitab 13 software using different values of (q).
3.4.4.2 The value of $(\mathrm{q})$ is determined when the graph indicates that the trend was minimized and following graph shows no difference with the previous one.

### 3.4.5 Order of "I'":

3.4.5.1 Four figures containing the original data, detrended, seasonally adjusted data and seasonally adjusted and detrended data using "Minitab 13" software.
3.4.5.2 If the figures show big difference then it needs differentiations, and if they do not show big difference, then there is no need for differentiation.
3.4.5.3 A diagnostic model diagrams for the ARIMA model is drawn using "Minitab 13 " software with $I=0,2$, because the seasonality in Jordan is affected by just the summer and winter seasons.
3.4.5.4 The ARIMA that has less residual between the two diagnostic models is considered to be the model, which we will use.

### 3.4.6 Forecasting future values: -

The method of forecasting was divided into two parts; the deterministic and the stochastic forecasting:

### 3.4.6.1 deterministic forecasting:

Four different methods were used in this section, these methods are: linear regression, quadratic regression, exponential growth regression and single exponential smoothing models. The acceptance of a model is determined if the model has an error less than $10 \%$ of the real data.

### 3.4.6.2 stochastic forecasting:

Three forecasting methods were used in this section, these methods are: autoregressive model, moving average model and the ARIMA model. If the error in the stochastic forecasting is less than $10 \%$, then the model will be the best one even if there is lower error in the deterministic model.

### 3.4.7 Results of forecasting: -

The results of the error in each model are shown in the percentage of error table, the calculations in this table were based on the difference in mean between the real data and the forecasted one, the stochastic forecasting is determined to be the best model if it satisfies the maximum acceptance error, which equals to $10 \%$, otherwise the lowest error in the in the deterministic model will be the best model.

### 3.4.8 Cross and distance correlation: -

3.4.8.2 The cross and distance correlation was drawn using Minitab 13 software
3.4.8.3 Analyses were made for each figure.

### 3.5. Analysis

In this section, the data that was collected will be analyzed; six variables will be analyzed in King Talal Reservoir. These variables are: Total suspended solids (TSS), biochemical oxygen demand $\left(\mathrm{BOD}_{5}\right)$, chemical oxygen demand (COD), Total Phosphorus (T-P), and finally the total Nitrogen (T-N). The method that each variable will be analyzed will be the same as discussed in the methodology. The best forecasting method would be determined in the end of each analysis.

### 3.5.1. Total suspended solids (TSS) variable:

The consequences that were used to analyze the TSS variable were as follows:

### 3.5.1.1. detection of missing data and outliers:

From the table (1) it is observed that the contains one missing data in December 1999, the observation will be estimated to be the average of the observations of the same month, which is December in this case. The new calculated value is $58.30 \mathrm{mg} / \mathrm{l}$. To detect the outliers, data should be drawn in a scatter diagram (Figure 3) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have approximately five outliers and they are in the following months: January 1988, May 1996, September 1996, March 1997, and finally, June 1998. It was observed that the rainfall in 1996 was high, and it is known that when the rainfall is high then the TSS will get higher (Appendix (1)). So the real data are on May and September 1996, the other three data were assumed to be outliers due to human error. They should be adjusted to a new value since they may greatly influence any statistical calculations and yield biased result. The way that outliers were adjusted was the same as the missing data treated and it was equals to the average monthly value.


Figure (3): Original Data of TSS

Figure (4) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are five outliers in both the original data and the residual data in the seasonal condition and that they are the same outliers in the scatter diagram of the original data. Also figure (4) shows the variation in the data for the same month, it can be observe the variation was the highest on January, and was the lowest on April.


Figure (4) Outliers for Seasonal Analysis for TSS Variable

After adjustment the outliers, the new adjusted data are plotted in Figure (4), the figure shows that there are still outliers but these outliers cannot be omitted because they are real data so it can influence the statistics results. While comparing the old data (Figure 3) with the new adjusted data (Figure 5) it can be observed that two figures are quite the same and they have the same trend, so the effect of the outliers on the data was so little.


Figure (5): The New Adjusted Data for TSS mg/l

### 3.5.1.2. normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the TSS variable; the calculated values were as follows

| Month | Jan. | Feb. | Mar. | Apr. | May | Jun. | Jul. | Aug. | Sep. | Oct. | Nov. | Dec. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| TSS mg/l | 65.2 | 45.4 | 46.4 | 50.5 | 68.0 | 76.0 | 81.0 | 81.9 | 83.0 | 81.5 | 77.5 | 61.9 |

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from figure (6) that the data of TSS is quite normal and there is a little
skewness to the left and bulked to the right, but in general the graph gives an indication that the data is normal.


Figure (6) : Weibull Distribution Model Histogram

## B- coefficient of variation (COV), preliminary test:

The data were divided into four quarters; each quarter consists of 39 data. Table (3) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the TSS variable.

Table (3): The coefficient of variable for TSS

|  | MEAN | VARIANCE | ST. DE. | C.O.V. |
| :--- | :--- | :---: | :---: | :---: | :---: |
| TSS $\quad\left(1^{\text {st }}\right.$ Quarter $)$ | 64.1 | 1244.5 | 35.3 | 0.55 |
| TSS $\quad\left(2^{\text {nd }}\right.$ Quarter $)$ | 67.6 | 442.7 | 21.0 | 0.31 |
| TSS $\quad\left(3^{\text {rd }}\right.$ Quarter $)$ | 80.5 | 1332.0 | 36.5 | 0.45 |
| TSS $\quad\left(4^{\text {th }}\right.$ Quarter $)$ | 65.7 | 535.5 | 23.1 | 0.35 |

It can be shown from table (3) that the value of the coefficient of variation for each quarter is less than 1 , which means that each quarter of the data has a little skewed (either to right or left), so the total data of the TSS variable has less skewness than each of the four TSS quarters, it can be concluded that the TSS variable does not have skewness

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K , which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (4) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of K, and the Kurtosis coefficient.

Table (4): The Kurtosis Coefficient for TSS

|  |  |  |  |  | Kurtosis |
| :--- | :--- | :---: | :---: | :---: | :---: | :---: |
| Coeff. |  |  |  |  |  |

From table (4) one can observe that the data in the first quarter was leptokurtic, in the second quarter it was normally distributed (mesokurtic), in the third quarter it was fairly leptokurtic and in the fourth quarter is was normally distributed.

The total data of the TSS variable can be assumed to be as fairly normally distributed (mesokurtic).

## D- Shapiro-Wilk test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of $\left(a_{n-1+1}\right)$, was taken for 20 data since the value of $n-1+i$ was equal to 20 , the value of $\left(a_{n-1+I}\right)$ was taken from Appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

Table (5): Shapiro-Wilk Test for the Data of TSS's $1^{\text {st }}$ quarter

| No | $\begin{aligned} & \mathrm{TSS} \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | Ordering TSS <br> (1) | Inverse order TSS (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{x} \mathrm{a}_{(\mathrm{n}-1 \mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 65.17 | 20.00 | 125.00 | 105.00 | 0.3989 | 41.88 |
| 2 | 42.00 | 27.00 | 103.00 | 76.00 | 0.2755 | 20.94 |
| 3 | 34.00 | 30.00 | 93.99 | 63.98 | 0.2380 | 15.23 |
| 4 | 41.00 | 34.00 | 92.00 | 58.00 | 0.2104 | 12.20 |
| 5 | 72.00 | 35.00 | 91.99 | 56.99 | 0.1880 | 10.71 |
| 6 | 83.00 | 36.00 | 83.00 | 47.00 | 0.1689 | 7.94 |
| 7 | 103.00 | 36.50 | 83.00 | 46.50 | 0.1520 | 7.07 |
| 8 | 60.00 | 40.00 | 82.00 | 42.00 | 0.1366 | 5.74 |
| 9 | 55.00 | 41.00 | 79.00 | 38.00 | 0.1225 | 4.65 |
| 10 | 47.00 | 42.00 | 74.99 | 32.99 | 0.1092 | 3.60 |
| 11 | 61.00 | 43.00 | 72.00 | 29.00 | 0.0967 | 2.80 |
| 12 | 72.00 | 45.50 | 72.00 | 26.50 | 0.0848 | 2.25 |
| 13 | 83.00 | 47.00 | 69.99 | 23.00 | 0.0733 | 1.69 |
| 14 | 27.00 | 47.00 | 65.00 | 18.00 | 0.0622 | 1.12 |
| 15 | 30.00 | 50.00 | 64.99 | 15.00 | 0.0515 | 0.77 |
| 16 | 36.50 | 52.00 | 61.00 | 9.00 | 0.0409 | 0.37 |
| 17 | 59.00 | 54.00 | 60.49 | 6.49 | 0.0305 | 0.20 |
| 18 | 55.00 | 54.01 | 60.00 | 5.99 | 0.0203 | 0.12 |
| 19 | 60.49 | 55.00 | 59.00 | 4.00 | 0.0101 | 0.04 |
| 20 | 54.01 | 55.00 | 55.00 | 0.00 |  | $\mathrm{b}=139.32$ |
| 21 | 52.00 | 59.00 | 55.00 | -4.00 |  | $\mathrm{S}=35.277$ |
| 22 | 64.99 | 60.00 | 54.01 | -5.99 |  |  |
| 23 | 74.99 | 60.49 | 54.00 | -6.49 |  |  |
| 24 | 45.50 | 61.00 | 52.00 | -9.00 |  | $\mathrm{W}=0.410<0.939$ |
| 25 | 35.00 | 64.99 | 50.00 | -15.00 |  |  |
| 26 | 43.00 | 65.17 | 47.00 | -18.17 |  | $=$ Didn't Satisfied |
| 27 | 40.00 | 69.99 | 47.00 | -23.00 |  |  |
| 28 | 36.00 | 72.00 | 45.50 | -26.50 |  |  |
| 29 | 54.00 | 72.00 | 43.00 | -29.00 |  |  |
| 30 | 69.99 | 74.99 | 42.00 | -32.99 |  |  |
| 31 | 50.00 | 79.00 | 41.00 | -38.00 |  |  |
| 32 | 91.99 | 82.00 | 40.00 | -42.00 |  |  |
| 33 | 93.99 | 83.00 | 36.50 | -46.50 |  |  |
| 34 | 125.00 | 83.00 | 36.00 | -47.00 |  |  |
| 35 | 92.00 | 91.99 | 35.00 | -56.99 |  |  |
| 36 | 82.00 | 92.00 | 34.00 | -58.00 |  |  |
| 37 | 79.00 | 93.99 | 30.00 | -63.98 |  |  |
| 38 | 20.00 | 103.00 | 27.00 | -76.00 |  |  |
| 39 | 47.00 | 125.00 | 20.00 | -105.00 |  |  |

Table (6): Shapiro-Wilk Test for the Data of TSS's $2^{\text {nd }}$ quarter

| No | $\begin{gathered} \mathrm{TSS} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering TSS <br> (1) | Inverse order TSS (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 51.00 | 26.00 | 118.00 | 92.00 | 0.3989 | 36.70 |
| 41 | 36.00 | 28.00 | 104.99 | 76.99 | 0.2755 | 21.21 |
| 42 | 38.00 | 36.00 | 102.99 | 66.99 | 0.2380 | 15.94 |
| 43 | 57.99 | 37.00 | 93.99 | 56.99 | 0.2104 | 11.99 |
| 44 | 56.01 | 38.00 | 92.99 | 54.99 | 0.1880 | 10.34 |
| 45 | 72.50 | 47.00 | 90.00 | 43.00 | 0.1689 | 7.26 |
| 46 | 85.01 | 51.00 | 89.00 | 38.00 | 0.1520 | 5.78 |
| 47 | 104.99 | 54.00 | 86.98 | 32.98 | 0.1366 | 4.50 |
| 48 | 28.00 | 56.00 | 86.00 | 30.00 | 0.1225 | 3.68 |
| 49 | 59.00 | 56.00 | 85.01 | 29.01 | 0.1092 | 3.17 |
| 50 | 65.00 | 56.01 | 81.00 | 24.99 | 0.0967 | 2.42 |
| 51 | 71.00 | 57.00 | 76.99 | 20.00 | 0.0848 | 1.70 |
| 52 | 59.00 | 57.99 | 75.00 | 17.01 | 0.0733 | 1.25 |
| 53 | 89.00 | 59.00 | 74.00 | 15.00 | 0.0622 | 0.93 |
| 54 | 118.00 | 59.00 | 72.50 | 13.50 | 0.0515 | 0.70 |
| 55 | 102.99 | 59.00 | 71.00 | 12.00 | 0.0409 | 0.49 |
| 56 | 86.98 | 60.00 | 70.00 | 10.00 | 0.0305 | 0.31 |
| 57 | 90.00 | 63.99 | 67.00 | 3.01 | 0.0203 | 0.06 |
| 58 | 92.99 | 65.00 | 67.00 | 2.00 | 0.0101 | 0.02 |
| 59 | 75.00 | 65.00 | 65.00 | 0.00 |  | $\mathrm{b}=128.43$ |
| 60 | 56.00 | 67.00 | 65.00 | -2.00 |  | $\mathrm{S}=21.041$ |
| 61 | 56.00 | 67.00 | 63.99 | -3.01 |  |  |
| 62 | 54.00 | 70.00 | 60.00 | -10.00 |  |  |
| 63 | 57.00 | 71.00 | 59.00 | -12.00 |  | $\mathrm{W}=0.980<0.939$ |
| 64 | 59.00 | 72.50 | 59.00 | -13.50 |  |  |
| 65 | 65.00 | 74.00 | 59.00 | -15.00 |  | $\underline{=}$ Satisfied |
| 66 | 70.00 | 75.00 | 57.99 | -17.01 |  |  |
| 67 | 74.00 | 76.99 | 57.00 | -20.00 |  |  |
| 68 | 76.99 | 81.00 | 56.01 | -24.99 |  |  |
| 69 | 86.00 | 85.01 | 56.00 | -29.01 |  |  |
| 70 | 93.99 | 86.00 | 56.00 | -30.00 |  |  |
| 71 | 81.00 | 86.98 | 54.00 | -32.98 |  |  |
| 72 | 67.00 | 89.00 | 51.00 | -38.00 |  |  |
| 73 | 47.00 | 90.00 | 47.00 | -43.00 |  |  |
| 74 | 26.00 | 92.99 | 38.00 | -54.99 |  |  |
| 75 | 60.00 | 93.99 | 37.00 | -56.99 |  |  |
| 76 | 37.00 | 102.99 | 36.00 | -66.99 |  |  |
| 77 | 63.99 | 104.99 | 28.00 | -76.99 |  |  |
| 78 | 67.00 | 118.00 | 26.00 | -92.00 |  |  |

Table (7): Shapiro-Wilk Test for the Data of TSS's $3{ }^{\text {rd }}$ quarter

| No | $\begin{gathered} \mathrm{TSS} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering TSS <br> (1) | Inverse order TSS <br> (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 79 | 82.99 | 22.00 | 187.96 | 165.96 | 0.3989 | 66.20 |
| 80 | 80.00 | 32.02 | 169.96 | 137.94 | 0.2755 | 38.00 |
| 81 | 93.99 | 33.95 | 151.11 | 117.16 | 0.2380 | 27.88 |
| 82 | 79.00 | 33.99 | 136.00 | 102.01 | 0.2104 | 21.46 |
| 83 | 22.00 | 37.00 | 127.01 | 90.02 | 0.1880 | 16.92 |
| 84 | 37.00 | 41.94 | 112.95 | 71.01 | 0.1689 | 11.99 |
| 85 | 62.95 | 50.96 | 111.05 | 60.09 | 0.1520 | 9.13 |
| 86 | 60.01 | 52.06 | 98.01 | 45.95 | 0.1366 | 6.28 |
| 87 | 50.96 | 54.94 | 97.04 | 42.10 | 0.1225 | 5.16 |
| 88 | 41.94 | 55.96 | 93.99 | 38.03 | 0.1092 | 4.15 |
| 89 | 33.95 | 60.01 | 92.95 | 32.93 | 0.0967 | 3.18 |
| 90 | 72.93 | 62.95 | 86.97 | 24.02 | 0.0848 | 2.04 |
| 91 | 111.05 | 65.90 | 84.01 | 18.12 | 0.0733 | 1.33 |
| 92 | 86.97 | 70.04 | 82.99 | 12.95 | 0.0622 | 0.81 |
| 93 | 84.01 | 70.09 | 81.93 | 11.84 | 0.0515 | 0.61 |
| 94 | 71.97 | 71.97 | 80.00 | 8.04 | 0.0409 | 0.33 |
| 95 | 70.09 | 72.93 | 79.06 | 6.13 | 0.0305 | 0.19 |
| 96 | 79.06 | 73.04 | 79.00 | 5.96 | 0.0203 | 0.12 |
| 97 | 97.04 | 73.93 | 75.11 | 1.17 | 0.0101 | 0.01 |
| 98 | 32.02 | 75.02 | 75.02 | 0.00 |  | $b=215.80$ |
| 99 | 33.99 | 75.11 | 73.93 | -1.17 |  | $\mathrm{S}=36.50$ |
| 100 | 54.94 | 79.00 | 73.04 | -5.96 |  |  |
| 101 | 169.96 | 79.06 | 72.93 | -6.13 |  |  |
| 102 | 81.93 | 80.00 | 71.97 | -8.04 |  | $\mathrm{W}=0.920<0.939$ |
| 103 | 52.06 | 81.93 | 70.09 | -11.84 |  |  |
| 104 | 112.95 | 82.99 | 70.04 | -12.95 | - | - Didn't Satisfied |
| 105 | 151.11 | 84.01 | 65.90 | -18.12 |  |  |
| 106 | 75.11 | 86.97 | 62.95 | -24.02 |  |  |
| 107 | 136.00 | 92.95 | 60.01 | -32.93 |  |  |
| 108 | 73.04 | 93.99 | 55.96 | -38.03 |  |  |
| 109 | 127.01 | 97.04 | 54.94 | -42.10 |  |  |
| 110 | 70.04 | 98.01 | 52.06 | -45.95 |  |  |
| 111 | 187.96 | 111.05 | 50.96 | -60.09 |  |  |
| 112 | 73.93 | 112.95 | 41.94 | -71.01 |  |  |
| 113 | 55.96 | 127.01 | 37.00 | -90.02 |  |  |
| 114 | 98.01 | 136.00 | 33.99 | -102.01 |  |  |
| 115 | 92.95 | 151.11 | 33.95 | -117.16 |  |  |
| 116 | 75.02 | 169.96 | 32.02 | -137.94 |  |  |
| 117 | 65.90 | 187.96 | 22.00 | -165.96 |  |  |

Table (8): Shapiro-Wilk Test for the Data of TSS's $4^{\text {th }}$ quarter

| No | TSS $\mathrm{mg} / \mathrm{l}$ | Ordering TSS <br> (1) | Inverse order TSS <br> (2) | 2-1 | $a(n-1+i)$ | $(2-1) \mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 70.95 | 6.96 | 106.12 | 99.16 | 0.3989 | 39.56 |
| 119 | 106.07 | 33.00 | 106.07 | 73.07 | 0.2755 | 20.13 |
| 120 | 43.97 | 33.06 | 105.95 | 72.89 | 0.2380 | 17.35 |
| 121 | 52.03 | 37.02 | 104.59 | 67.57 | 0.2104 | 14.22 |
| 122 | 53.02 | 39.06 | 97.03 | 57.96 | 0.1880 | 10.90 |
| 123 | 66.98 | 39.99 | 96.90 | 56.91 | 0.1689 | 9.61 |
| 124 | 37.02 | 40.04 | 89.88 | 49.84 | 0.1520 | 7.58 |
| 125 | 39.06 | 43.97 | 87.92 | 43.95 | 0.1366 | 6.00 |
| 126 | 6.96 | 44.00 | 85.84 | 41.84 | 0.1225 | 5.13 |
| 127 | 96.90 | 48.00 | 85.06 | 37.06 | 0.1092 | 4.05 |
| 128 | 71.10 | 52.03 | 80.89 | 28.87 | 0.0967 | 2.79 |
| 129 | 68.96 | 53.02 | 79.91 | 26.89 | 0.0848 | 2.28 |
| 130 | 71.93 | 54.06 | 78.09 | 24.03 | 0.0733 | 1.76 |
| 131 | 63.01 | 57.03 | 74.98 | 17.96 | 0.0622 | 1.12 |
| 132 | 48.00 | 58.00 | 71.93 | 13.94 | 0.0515 | 0.72 |
| 133 | 40.04 | 63.01 | 71.10 | 8.09 | 0.0409 | 0.33 |
| 134 | 39.99 | 64.10 | 70.95 | 6.85 | 0.0305 | 0.21 |
| 135 | 33.06 | 66.98 | 68.96 | 1.98 | 0.0203 | 0.04 |
| 136 | 74.98 | 67.94 | 68.02 | 0.07 | 0.0101 | 0.00 |
| 137 | 67.94 | 67.99 | 67.99 | 0.00 |  | $\mathrm{b}=143.76$ |
| 138 | 89.88 | 68.02 | 67.94 | -0.07 |  | $\mathrm{S}=23.14$ |
| 139 | 80.89 | 68.96 | 66.98 | -1.98 |  |  |
| 140 | 106.12 | 70.95 | 64.10 | -6.85 |  |  |
| 141 | 79.91 | 71.10 | 63.01 | -8.09 |  | $\mathrm{W}=1.016<0.939$ |
| 142 | 97.03 | 71.93 | 58.00 | -13.94 |  |  |
| 143 | 64.10 | 74.98 | 57.03 | -17.96 |  |  |
| 144 | 104.59 | 78.09 | 54.06 | -24.03 |  | $=$ Satisfied |
| 145 | 44.00 | 79.91 | 53.02 | -26.89 |  |  |
| 146 | 58.00 | 80.89 | 52.03 | -28.87 |  |  |
| 147 | 33.00 | 85.06 | 48.00 | -37.06 |  |  |
| 148 | 54.06 | 85.84 | 44.00 | -41.84 |  |  |
| 149 | 78.09 | 87.92 | 43.97 | -43.95 |  |  |
| 150 | 67.99 | 89.88 | 40.04 | -49.84 |  |  |
| 151 | 87.92 | 96.90 | 39.99 | -56.91 |  |  |
| 152 | 105.95 | 97.03 | 39.06 | -57.96 |  |  |
| 153 | 85.84 | 104.59 | 37.02 | -67.57 |  |  |
| 154 | 85.06 | 105.95 | 33.06 | -72.89 |  |  |
| 155 | 57.03 | 106.07 | 33.00 | -73.07 |  |  |
| 156 | 68.02 | 106.12 | 6.96 | -99.16 |  |  |

From the Tables shown previously, it has been shown that the data in the first quarter is nonnormal, the second quarter has a normal data, the third has a quite nonnormal one, and the fourth one has a normal distribution. It can be safely say that TSS variable is normally distributed.

### 3.5.1.3. order of (AR)

For water quality like King Talal Dam, the value of $A R$, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of TSS does not need more than 1 month till it sediments (Viessman and Lewis, 1996). From Figure (7) it can be seen that the value of AR is equal to 1 , so the value of p will be 1 for the TSS variable.


Figure (7) Autocorrelation Function for TSS Variable

### 3.5.1.4. order of moving average (MA)

After finding the value of AR , which was 1 , the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (8) shows the change between the real data of the variable TSS and it's moving average with different length of p .


Figure (8) Moving Average of TSS with Different Values of (p)

The moving average can be determined from Figure (8) when the difference between the previous length of p and the followed one have a small difference and that occurred when the value of p was 4 (as shown in Figure (8)), so the TSS variable has a value of MA(4).

### 3.5.1.5. order of (I)

The last coefficient of ARIMA's parameters is the integrated model (I), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (9) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that they are almost the same, which means that the detrended and seasonally effects are almost negligible.


Figure (9): Component Analysis for TSS mg/l

Two season; summer and winter can affect seasonality in Jordan, so if the data has no seasonality effect, then the value of $d=0$ and if we have seasonality effect then the value of $d=2$. Figures (10) and (11) provides ARIMA model diagnostics for ARIMA $=(1,0,4)$ and $(1,2,4)$. It is seen from the two graphs that the residual in Figure (10) is less than Figure (11) so the coefficients of ARIMA that will be used are $(1,0,4)$


Figure (9): ARIMA (1,0,4) Model Diagnostics for TSS


Figure (10): ARIMA $(1,2,4)$ Model Diagnostics for TSS

### 3.5.1.6. forecasting future values

The following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last $10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (12).


Figure (12): Trend Analysis for TSS mg/l
It can be observed from the above figure and equation of the linear trend that the data is increasing slowly. Table (9) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (9): The values of the predicted and actual data by linear regression for TSS variable

| Row   <br> 1  Period (months) | Forecasted $(\mathrm{mg} / \mathrm{l})$ | Actual $(\mathrm{mg} / \mathrm{l})$ <br> 2 | 141 |
| :---: | :---: | :---: | :---: |
|  | 75.96 | 79.91 |  |
| 3 | 142 | 76.07 | 97.03 |
| 4 | 143 | 76.19 | 64.10 |
| 5 | 144 | 76.31 | 104.59 |
| 6 | 145 | 76.43 | 44.00 |
| 7 | 146 | 76.55 | 58.00 |
| 8 | 147 | 76.66 | 33.00 |
| 9 | 148 | 76.78 | 54.06 |
| 10 | 149 | 76.90 | 78.09 |
| 11 | 150 | 77.02 | 67.99 |
| 12 | 151 | 77.14 | 87.92 |
| 13 | 152 | 77.26 | 105.95 |
| 14 | 153 | 77.37 | 85.84 |
| 15 | 154 | 77.49 | 85.06 |
| 16 | 155 | 77.61 | 57.03 |
| Comparing the actual values with the predicted ones, one can conclude, after |  |  |  |

calculating the predication error, which equals to $5.0 \%$, that the linear trend model has satisfied the forecasting for the TSS variable.

## B3- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (13).


Figure (13): Trend Analysis for TSS mg/l

It can be observed from the above figure and the equation of the quadratic trend that the data is concaved down or in another way that there is an increase in the beginning and then a decrease in the end. Table (10) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (10): The values of the predicted and actual data by quadratic regression for TSS variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 68.91 | 79.91 |
| 2 | 142 | 68.73 | 97.03 |
| 3 | 143 | 68.54 | 64.10 |
| 4 | 144 | 68.35 | 104.59 |
| 5 | 145 | 68.16 | 44.00 |
| 6 | 146 | 67.96 | 58.00 |
| 7 | 147 | 67.76 | 33.00 |
| 8 | 148 | 67.55 | 54.06 |
| 9 | 149 | 67.34 | 78.09 |
| 10 | 150 | 67.12 | 67.99 |
| 11 | 151 | 66.90 | 87.92 |
| 12 | 152 | 66.68 | 105.95 |
| 13 | 153 | 66.45 | 85.84 |
| 14 | 154 | 66.22 | 85.06 |
| 15 | 155 | 65.98 | 57.03 |
| 16 | 156 | 65.74 | 68.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $8.6 \%$, that the quadratic trend model has satisfied the forecasting for the TSS variable.

## B3- exponential growth regression model

The regression of the additive exponential growth trend model is shown in Figure (14).


Figure (14): Trend Analysis for TSS mg/l

It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing trend. Table (11) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (11): The values of the predicted and actual data by exponential growth regression for TSS variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 71.01 | 79.91 |
| 2 | 142 | 71.14 | 97.03 |
| 3 | 143 | 71.26 | 64.10 |
| 4 | 144 | 71.38 | 104.59 |
| 5 | 145 | 71.51 | 44.00 |
| 6 | 146 | 71.63 | 58.00 |
| 7 | 147 | 71.76 | 33.00 |
| 8 | 148 | 71.88 | 54.06 |
| 9 | 149 | 72.01 | 78.09 |
| 10 | 150 | 72.13 | 67.99 |
| 11 | 151 | 72.26 | 87.92 |
| 12 | 152 | 72.38 | 105.95 |
| 13 | 153 | 72.51 | 85.84 |
| 14 | 154 | 72.64 | 85.06 |
| 15 | 155 | 72.76 | 57.03 |
| 16 | 156 | 72.89 | 68.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $1.7 \%$, that the exponential growth trend model has satisfied the forecasting for the TSS variable.

## B3- single exponential smoothing model

The regression of the additive single exponential smoothing trend model is shown in Figure (15).


Figure (15): Single Exponential Smoothing for TSS mg/l

Table (12) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (12) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (12): Forecasted, lower, upper and actual values for single exponential smoothing for TSS variable

| $\underline{\text { Row }}$ | $\frac{\text { Period }}{(\mathrm{month})}$ | Forecast <br> $\mathrm{mg} / \mathrm{l}$ | Lower <br> $\mathrm{mg} / \mathrm{l}$ | $\frac{\text { Upper }}{\mathrm{mg} / \mathrm{l}}$ | $\underline{\text { Actual }}$ <br> 1 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 141 | 83.49 | 35.72 | 131.26 | 79.91 |
| 2 | 142 | 83.49 | 35.72 | 131.26 | 97.03 |
| 3 | 143 | 83.49 | 35.72 | 131.26 | 64.10 |
| 4 | 144 | 83.49 | 35.72 | 131.26 | 104.59 |
| 5 | 145 | 83.49 | 35.72 | 131.26 | 44.00 |
| 6 | 146 | 83.49 | 35.72 | 131.26 | 58.00 |
| 7 | 147 | 83.49 | 35.72 | 131.26 | 33.00 |
| 8 | 148 | 83.49 | 35.72 | 131.26 | 54.06 |
| 9 | 149 | 83.49 | 35.72 | 131.26 | 78.09 |
| 10 | 150 | 83.49 | 35.72 | 131.26 | 67.99 |
| 11 | 151 | 83.49 | 35.72 | 131.26 | 87.92 |
| 12 | 152 | 83.49 | 35.72 | 131.26 | 105.95 |
| 13 | 153 | 83.49 | 35.72 | 131.26 | 85.84 |
| 14 | 154 | 83.49 | 35.72 | 131.26 | 85.06 |
| 15 | 155 | 83.49 | 35.72 | 131.26 | 57.03 |
| 16 | 156 | 83.49 | 35.72 | 131.26 | 68.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $12.4 \%$, that the simple exponential smoothing trend model has not satisfied the forecasting for the TSS variable.

## B- stochastic forecasting

## B3- auto regression model

Table (13) shows the $\operatorname{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (13) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (13): Forecasted, lower, upper and actual values for $\operatorname{AR}(1)$ for TSS variable

| Row | $\frac{\text { Period }}{(\mathrm{month})}$ | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |  | $\frac{\text { Lower }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Upper }}{\underline{\mathrm{mg} / / \mathrm{l}}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $5.4 \%$, that the $\operatorname{AR}(1)$ trend model has satisfied the forecasting for the TSS variable.

## B3- moving average regression model

The regression of the additive MA(4) trend model is shown in Figure(16).


Figure (16): Moving Average Trend for TSS mg/l
Table (14) shows the $\mathrm{MA}(4)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (14) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (14): Forecasted, lower, upper and actual values for MA(4) for TSS variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | $\underline{\mathrm{mg} / \mathrm{l}}$ | mg/l | $\underline{\mathrm{mg} / \mathrm{l}}$ | mg/l |
| 1 | 141 | 82.32 | 27.60 | 137.03 | 79.91 |
| 2 | 142 | 82.32 | 27.60 | 137.03 | 97.03 |
| 3 | 143 | 82.32 | 27.60 | 137.03 | 64.10 |
| 4 | 144 | 82.32 | 27.60 | 137.03 | 104.59 |
| 5 | 145 | 82.32 | 27.60 | 137.03 | 44.00 |
| 6 | 146 | 82.32 | 27.60 | 137.03 | 58.00 |
| 7 | 147 | 82.32 | 27.60 | 137.03 | 33.00 |
| 8 | 148 | 82.32 | 27.60 | 137.03 | 54.06 |
| 9 | 149 | 82.32 | 27.60 | 137.03 | 78.09 |
| 10 | 150 | 82.32 | 27.60 | 137.03 | 67.99 |
| 11 | 151 | 82.32 | 27.60 | 137.03 | 87.92 |
| 12 | 152 | 82.32 | 27.60 | 137.03 | 105.95 |
| 13 | 153 | 82.32 | 27.60 | 137.03 | 85.84 |
| 14 | 154 | 82.32 | 27.60 | 137.03 | 85.06 |
| 15 | 155 | 82.32 | 27.60 | 137.03 | 57.03 |
| 16 | 156 | 82.32 | 27.60 | 137.03 | 68.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $11.1 \%$, that the $\mathrm{MA}(5)$ trend model has not satisfied the forecasting for the TSS variable.

## B3- ARMA modeling

Table (15) shows the $\operatorname{ARMA}(1,4)$ prediction values for the next $10 \%$ of the predicted and the real data, which equals to 16 observations. In Table (15) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (15): Forecasted, lower, upper and actual values for $\operatorname{ARIMA}(1,0,4)$ for $\operatorname{TSS}$ variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | mg/l | mg/l | mg/l | mg/l |
| 1 | 141 | 67.77 | 66.95 | 68.58 | 79.91 |
| 2 | 142 | 67.82 | 66.33 | 69.31 | 97.03 |
| 3 | 143 | 67.72 | 65.76 | 69.69 | 64.10 |
| 4 | 144 | 67.61 | 65.46 | 69.75 | 104.59 |
| 5 | 145 | 67.60 | 65.46 | 69.75 | 44.00 |
| 6 | 146 | 67.60 | 65.46 | 69.75 | 58.00 |
| 7 | 147 | 67.60 | 65.46 | 69.75 | 33.00 |
| 8 | 148 | 67.60 | 65.46 | 69.75 | 54.06 |
| 9 | 149 | 67.60 | 65.46 | 69.75 | 78.09 |
| 10 | 150 | 67.60 | 65.46 | 69.75 | 67.99 |
| 11 | 151 | 67.60 | 65.46 | 69.75 | 87.92 |
| 12 | 152 | 67.60 | 65.46 | 69.75 | 105.95 |
| 13 | 153 | 67.60 | 65.46 | 69.75 | 85.84 |
| 14 | 154 | 67.60 | 65.46 | 69.75 | 85.06 |
| 15 | 155 | 67.60 | 65.46 | 69.75 | 57.03 |
| 16 | 156 | 67.60 | 65.46 | 69.75 | 68.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $8.2 \%$, that the $\operatorname{ARIMA}(1,0,4)$ trend model has satisfied the forecasting for the TSS variable.

### 3.5.1.7 Results of Prediction

The results of error are summarized in the following Table (16), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (16) : Percentage of error of each model for TSS variable

| Model | Percentage of <br> Mean Error |
| :--- | :---: |
| Linear Method | $5.0 \%$ |
| Quadratic Method | $8.6 \%$ |
| Exponential Growth Method | $1.7 \%$ |
| Simple Exponential Smoothing | $12.4 \%$ |
| Auto Regression, AR(1) | $5.4 \%$ |
| Moving Average, MA(4) | $11.1 \%$ |
| ARIMA (1,0,4) | $8.2 \%$ |

The previous Table (16) shows that all the methods have satisfied the $10 \%$ acceptable prediction limits, except in MA(4) and in simple exponential smoothing. When finding the best model that gave the least error it will be the exponential growth method, this is for deterministic forecasting. But one should take into consideration that we deal with stochastic method, so AR (1) will be the best method for forecasting.

### 3.5.2 Biochemical oxygen demand $\left(\mathrm{BOD}_{5}\right)$ variable:

The consequences that were used to analyze the $\mathrm{BOD}_{5}$ variable were as follows:

### 3.5.2.1 detection of missing data and outliers:

From the table (1) it is observed that the data do not contain any missing data, so the second step is to find the outliers, data should be drawn in a scatter diagram (Figure 17) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have approximately three outliers and they are in the following months: January 1995, January 1997, and December 1997. It was observed that the rainfall in January 1995 was low, and it is known that when the rainfall is low then the $\mathrm{BOD}_{5}$ will get high, in January and December 1997 the rainfall was high so the $\mathrm{BOD}_{5}$ should be low (Appendix (1)). So the real data is on January 1995, the other three data were assumed to be outliers due to human error, and they should be adjusted to a new value since they may greatly influence any statistical calculations and yield biased result. The way that outliers were adjusted was the same as the missing data treated and it was equals to the average monthly value.


Figure (17): Original Data of $\mathrm{BOD}_{5} \mathrm{mg} / \mathrm{l}$

Figure (18) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are five outliers in both the original data and the residual data in the seasonal condition. Also figure (18) shows the variation in the data for the same month, it can be observe that the variation was the highest on December, and was the lowest on February. Another two outliers were found in the seasonal drawings, they are in November 1996, and March 1997. There was high rainfall in March 1997 so it should be adjusted (Appendix (1)), while in November 1996 it was a real data.

## Seasonal Analysis for BOD5 mg/l



Figure (18) Outliers for Seasonal Analysis for BOD5 Variable

After adjustment the outliers, the new adjusted data are plotted in Figure (17), the figure shows that their still outliers but these outliers cannot be omitted because they are real data so it can influence the statistics results. While comparing the old data (Figure 17) with the new adjusted data (Figure 19) it can be observed that two figures are quite the same and they have the same trend, so the effect of the outliers on the data was so little.


Figure (19): The New Adjusted Data of BOD5 mg/l

### 3.5.2.2 normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the $\mathrm{BOD}_{5}$ variable; the calculated values were as follows

Month Jan. Feb. Mar. Apr. May Jun. Jul. Aug. Sep. Oct. Nov. Dec.<br>$\begin{array}{lllllllllllll}\text { BOD5 } & 36.7 & 28.1 & 28.0 & 42.7 & 42.8 & 45.7 & 39.5 & 41.5 & 38.1 & 36.2 & 37.8 & 36.2\end{array}$

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from figure (20) that the data of $\mathrm{BOD}_{5}$ is quite normal and there is a little skewness to the left and bulked to the right, but in general the graph gives an indication that the data is normal.


Figure (20) : Weibull Distribution Model Histogram

## B- coefficient of variation (COV), preliminary test:

The data were divided into four quarters; each quarter consists of 39 data. Table (17) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the $\mathrm{BOD}_{5}$ variable.

Table (17): The coefficient of variable for $\mathrm{BOD}_{5}$

|  | MEAN | VARIANCE | ST. DE. <br> $(\mathrm{S})$ | C.O.V. |  |
| :--- | :--- | :---: | :---: | :---: | :---: |
| $\mathrm{BOD}_{5}$ | $\left(1^{\text {st }}\right.$ Quarter $)$ | 25.8 | 116.6 | 10.8 | 0.4 |
| $\mathrm{BOD}_{5} \quad\left(2^{\text {nd }}\right.$ Quarter $)$ | 31.7 | 164.9 | 12.8 | 0.4 |  |
| $\mathrm{BOD}_{5} \quad\left(3^{\text {rd }}\right.$ Quarter $)$ | 54.3 | 479.1 | 21.9 | 0.4 |  |
| $\mathrm{BOD}_{5} \quad\left(4^{\text {th }}\right.$ Quarter $)$ | 45.4 | 398.8 | 20.0 | 0.4 |  |

It can be shown from the table that the value of the coefficient of variation for each quarter is less than 1 , which means that each quarter of the data has a little skewed (either to right or left), so the total data of the $\mathrm{BOD}_{5}$ variable has less skewness than
each of the four $\mathrm{BOD}_{5}$ quarters, it can be concluded that the $\mathrm{BOD}_{5}$ variable does not have skewness.

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K, which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (18) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of K, and the Kurtosis coefficient.

Table (18): The Kurtosis Coefficient for $\mathrm{BOD}_{5}$

|  | MEAN | VARIANCE | ST. DE. <br> $(\mathrm{S})$ | K | Kurtosis <br> Coeff. <br> $\mathrm{C}_{\mathrm{K}}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| BOD5 (1st Quarter ) | 25.8 | 116.6 | 10.8 | 30758.3 | -0.7 |
| BOD5 (2nd Quarter ) | 31.7 | 164.9 | 12.8 | 60771.7 | -0.8 |
| BOD5 (3rd Quarter ) | 54.3 | 479.1 | 21.9 | 1114502.6 | 1.9 |
| BOD5 (4th Quarter ) | 45.4 | 398.8 | 20.0 | 2021622.8 | 9.7 |

From table (18) one can observe that the data in the first and second quarters were normally distributed (mesokurtic), in the third quarter it was fairly leptokurtic and in the fourth quarter is was leptokurtic. The total data of the BOD5 variable can be assumed to be as fairly normally distributed (mesokurtic).

## D- Shapiro-Wilk test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of $\left(a_{n-1+1}\right)$, was taken for 20 data since the value of $n-1+\mathrm{i}$ was equal to 20 , the value of $\left(\mathrm{a}_{\mathrm{n}-1+\mathrm{I}}\right)$ was taken from appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

From the Tables (19), (20), (21), and (22) it has been shown that the data in the first quarter is nonnormal, the second quarter has a normal data, the third has a quite not normal one, and the fourth one has a normal distribution. It can be safely say that $\mathrm{BOD}_{5}$ variable is normally distributed.

Table (19): Shapiro-Wilk Test for the Data of $\mathrm{BOD}_{5}$ 's $1^{\text {st }}$ quarter

| No | $\begin{gathered} \mathrm{BOD}_{5} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering $\mathrm{BOD}_{5}$ <br> (1) | Inverse order $\mathrm{BOD}_{5}$ <br> (2) | 2-1 | A(n-1+i) | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 47.00 | 6.00 | 47.00 | 41.00 | 0.3989 | 16.35 |
| 2 | 10.00 | 10.00 | 47.00 | 37.00 | 0.2755 | 10.19 |
| 3 | 6.00 | 11.00 | 42.00 | 31.00 | 0.2380 | 7.38 |
| 4 | 34.00 | 11.50 | 39.00 | 27.50 | 0.2104 | 5.79 |
| 5 | 36.00 | 12.00 | 37.00 | 25.00 | 0.1880 | 4.70 |
| 6 | 33.00 | 13.00 | 36.00 | 23.00 | 0.1689 | 3.88 |
| 7 | 35.00 | 14.50 | 36.00 | 21.50 | 0.1520 | 3.27 |
| 8 | 17.00 | 15.00 | 36.00 | 21.00 | 0.1366 | 2.87 |
| 9 | 36.00 | 16.00 | 36.00 | 20.00 | 0.1225 | 2.45 |
| 10 | 30.00 | 16.00 | 36.00 | 20.00 | 0.1092 | 2.18 |
| 11 | 26.00 | 17.00 | 35.00 | 18.00 | 0.0967 | 1.74 |
| 12 | 36.00 | 18.00 | 34.00 | 16.00 | 0.0848 | 1.36 |
| 13 | 12.00 | 18.50 | 33.00 | 14.50 | 0.0733 | 1.06 |
| 14 | 11.50 | 19.00 | 33.00 | 14.00 | 0.0622 | 0.87 |
| 15 | 22.00 | 21.00 | 30.00 | 9.00 | 0.0515 | 0.46 |
| 16 | 14.50 | 22.00 | 30.00 | 7.99 | 0.0409 | 0.33 |
| 17 | 37.00 | 22.00 | 29.00 | 7.00 | 0.0305 | 0.21 |
| 18 | 25.00 | 25.00 | 27.00 | 2.01 | 0.0203 | 0.04 |
| 19 | 25.00 | 25.00 | 26.00 | 1.00 | 0.0101 | 0.01 |
| 20 | 18.50 | 25.00 | 25.00 | 0.00 |  | $\mathrm{b}=65.15$ |
| 21 | 16.00 | 26.00 | 25.00 | -1.00 |  | $\mathrm{S}=10.80$ |
| 22 | 16.00 | 27.00 | 25.00 | -2.01 |  |  |
| 23 | 18.00 | 29.00 | 22.00 | -7.00 |  |  |
| 24 | 21.00 | 30.00 | 22.00 | -7.99 |  | $\mathrm{W}=0.958>0.939$ |
| 25 | 13.00 | 30.00 | 21.00 | -9.00 |  |  |
| 26 | 19.00 | 33.00 | 19.00 | -14.00 |  | Satisfied |
| 27 | 15.00 | 33.00 | 18.50 | -14.50 |  |  |
| 28 | 22.00 | 34.00 | 18.00 | -16.00 |  |  |
| 29 | 27.00 | 35.00 | 17.00 | -18.00 |  |  |
| 30 | 36.00 | 36.00 | 16.00 | -20.00 |  |  |
| 31 | 33.00 | 36.00 | 16.00 | -20.00 |  |  |
| 32 | 25.00 | 36.00 | 15.00 | -21.00 |  |  |
| 33 | 30.00 | 36.00 | 14.50 | -21.50 |  |  |
| 34 | 42.00 | 36.00 | 13.00 | -23.00 |  |  |
| 35 | 36.00 | 37.00 | 12.00 | -25.00 |  |  |
| 36 | 39.00 | 39.00 | 11.50 | -27.50 |  |  |
| 37 | 47.00 | 42.00 | 11.00 | -31.00 |  |  |
| 38 | 29.00 | 47.00 | 10.00 | -37.00 |  |  |
| 39 | 11.00 | 47.00 | 6.00 | -41.00 |  |  |

Table (20): Shapiro-Wilk Test for the Data of $\mathrm{BOD}_{5}$ 's $2^{\text {nd }}$ quarter

| No | $\begin{gathered} \mathrm{BOD}_{5} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering $\mathrm{BOD}_{5}$ <br> (1) | Inverse order $\mathrm{BOD}_{5}$ (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 28.00 | 11.00 | 58.00 | 47.00 | 0.3989 | 18.75 |
| 41 | 30.00 | 11.00 | 54.00 | 43.00 | 0.2755 | 11.85 |
| 42 | 26.00 | 13.00 | 52.00 | 39.00 | 0.2380 | 9.28 |
| 43 | 21.00 | 15.00 | 50.00 | 35.00 | 0.2104 | 7.36 |
| 44 | 21.00 | 15.00 | 47.00 | 32.00 | 0.1880 | 6.02 |
| 45 | 25.17 | 15.00 | 45.00 | 30.00 | 0.1689 | 5.07 |
| 46 | 44.01 | 19.00 | 45.00 | 26.00 | 0.1520 | 3.95 |
| 47 | 40.00 | 21.00 | 44.01 | 23.01 | 0.1366 | 3.14 |
| 48 | 11.00 | 21.00 | 44.00 | 23.00 | 0.1225 | 2.82 |
| 49 | 23.00 | 21.00 | 43.00 | 22.00 | 0.1092 | 2.40 |
| 50 | 19.00 | 22.00 | 42.00 | 19.99 | 0.0967 | 1.93 |
| 51 | 15.00 | 23.00 | 40.00 | 17.00 | 0.0848 | 1.44 |
| 52 | 21.00 | 24.00 | 40.00 | 16.00 | 0.0733 | 1.17 |
| 53 | 25.00 | 25.00 | 39.00 | 14.00 | 0.0622 | 0.87 |
| 54 | 28.00 | 25.17 | 38.00 | 12.83 | 0.0515 | 0.66 |
| 55 | 33.00 | 26.00 | 36.99 | 10.99 | 0.0409 | 0.45 |
| 56 | 36.99 | 28.00 | 33.00 | 5.00 | 0.0305 | 0.15 |
| 57 | 24.00 | 28.00 | 33.00 | 5.00 | 0.0203 | 0.10 |
| 58 | 11.00 | 28.00 | 33.00 | 5.00 | 0.0101 | 0.05 |
| 59 | 13.00 | 30.00 | 30.00 | 0.00 |  | $\mathrm{b}=77.47$ |
| 60 | 15.00 | 33.00 | 28.00 | -5.00 |  | $\mathrm{S}=12.84$ |
| 61 | 15.00 | 33.00 | 28.00 | -5.00 |  |  |
| 62 | 22.00 | 33.00 | 28.00 | -5.00 |  |  |
| 63 | 33.00 | 36.99 | 26.00 | -10.99 |  | $\mathrm{W}=0.958>0.939$ |
| 64 | 43.00 | 38.00 | 25.17 | -12.83 |  |  |
| 65 | 38.00 | 39.00 | 25.00 | -14.00 |  | Satisfied |
| 66 | 33.00 | 40.00 | 24.00 | -16.00 |  |  |
| 67 | 44.00 | 40.00 | 23.00 | -17.00 |  |  |
| 68 | 54.00 | 42.00 | 22.00 | -19.99 |  |  |
| 69 | 47.00 | 43.00 | 21.00 | -22.00 |  |  |
| 70 | 40.00 | 44.00 | 21.00 | -23.00 |  |  |
| 71 | 45.00 | 44.01 | 21.00 | -23.01 |  |  |
| 72 | 50.00 | 45.00 | 19.00 | -26.00 |  |  |
| 73 | 45.00 | 45.00 | 15.00 | -30.00 |  |  |
| 74 | 39.00 | 47.00 | 15.00 | -32.00 |  |  |
| 75 | 28.00 | 50.00 | 15.00 | -35.00 |  |  |
| 76 | 58.00 | 52.00 | 13.00 | -39.00 |  |  |
| 77 | 42.00 | 54.00 | 11.00 | -43.00 |  |  |
| 78 | 52.00 | 58.00 | 11.00 | -47.00 |  |  |

Table (21): Shapiro-Wilk Test for the Data of $\mathrm{BOD}_{5}$ 's $3^{\text {rd }}$ quarter

|  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| No | $\mathrm{BOD}_{5}$ <br> $\mathrm{mg} / \mathrm{l}$ | Ordering <br> $\mathrm{BOD}_{5}$ <br> $(1)$ | Inverse order <br> $\mathrm{BOD}_{5}$ <br> $(2)$ | $2-1$ |  |  |
| $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | $(2-1) \mathrm{x} \mathrm{a}_{(\mathrm{n}-1+\mathrm{i})}$ |  |  |  |  |  |
| 79 | 60.99 | 15.00 | 125.00 | 110.00 | 0.3989 | 43.88 |
| 80 | 63.00 | 18.97 | 95.03 | 76.06 | 0.2755 | 20.95 |
| 81 | 47.00 | 28.00 | 92.00 | 64.00 | 0.2380 | 15.23 |
| 82 | 28.00 | 32.04 | 85.00 | 52.96 | 0.2104 | 11.14 |
| 83 | 45.00 | 33.08 | 80.04 | 46.96 | 0.1880 | 8.83 |
| 84 | 15.00 | 33.89 | 77.08 | 43.19 | 0.1689 | 7.30 |
| 85 | 92.00 | 36.96 | 74.02 | 37.07 | 0.1520 | 5.63 |
| 86 | 32.04 | 36.99 | 69.87 | 32.88 | 0.1366 | 4.49 |
| 87 | 41.99 | 38.98 | 68.06 | 29.08 | 0.1225 | 3.56 |
| 88 | 55.01 | 40.01 | 65.08 | 25.07 | 0.1092 | 2.74 |
| 89 | 61.97 | 41.04 | 63.06 | 22.03 | 0.0967 | 2.13 |
| 90 | 50.03 | 41.99 | 63.00 | 21.02 | 0.0848 | 1.78 |
| 91 | 54.04 | 43.05 | 61.97 | 18.92 | 0.0733 | 1.39 |
| 92 | 45.93 | 44.04 | 60.99 | 16.95 | 0.0622 | 1.05 |
| 93 | 51.62 | 45.00 | 56.96 | 11.96 | 0.0515 | 0.62 |
| 94 | 45.95 | 45.93 | 55.01 | 9.09 | 0.0409 | 0.37 |
| 95 | 18.97 | 45.95 | 54.04 | 8.09 | 0.0305 | 0.25 |
| 96 | 36.96 | 47.00 | 54.03 | 7.03 | 0.0203 | 0.14 |
| 97 | 38.98 | 47.05 | 51.62 | 4.57 | 0.0101 | 0.05 |
| 98 | 47.05 | 50.03 | 50.03 | 0.00 |  | $\mathrm{b}=131.53$ <br> 99 43.05 |

Table (22): Shapiro-Wilk Test for the Data of $\mathrm{BOD}_{5}$ 's $4^{\text {th }}$ quarter

| No | $\begin{gathered} \mathrm{BOD}_{5} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering $\mathrm{BOD}_{5}$ <br> (1) | Inverse order $\mathrm{BOD}_{5}$ <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 31.89 | 18.93 | 135.00 | 116.07 | 0.3989 | 46.30 |
| 119 | 46.98 | 22.05 | 77.00 | 54.96 | 0.2755 | 15.14 |
| 120 | 135.00 | 24.98 | 64.03 | 39.05 | 0.2380 | 9.29 |
| 121 | 29.97 | 25.07 | 62.09 | 37.02 | 0.2104 | 7.79 |
| 122 | 30.96 | 29.97 | 61.90 | 31.93 | 0.1880 | 6.00 |
| 123 | 36.98 | 30.04 | 59.10 | 29.06 | 0.1689 | 4.91 |
| 124 | 25.07 | 30.13 | 58.06 | 27.93 | 0.1520 | 4.25 |
| 125 | 31.96 | 30.89 | 57.97 | 27.08 | 0.1366 | 3.70 |
| 126 | 57.97 | 30.96 | 57.97 | 27.01 | 0.1225 | 3.31 |
| 127 | 45.01 | 31.89 | 57.12 | 25.23 | 0.1092 | 2.76 |
| 128 | 57.12 | 31.96 | 57.08 | 25.12 | 0.0967 | 2.43 |
| 129 | 38.92 | 34.87 | 56.07 | 21.20 | 0.0848 | 1.80 |
| 130 | 59.10 | 35.90 | 51.94 | 16.04 | 0.0733 | 1.18 |
| 131 | 51.94 | 35.98 | 47.11 | 11.13 | 0.0622 | 0.69 |
| 132 | 47.11 | 36.98 | 46.98 | 10.00 | 0.0515 | 0.51 |
| 133 | 40.99 | 37.86 | 45.08 | 7.21 | 0.0409 | 0.30 |
| 134 | 24.98 | 38.92 | 45.01 | 6.09 | 0.0305 | 0.19 |
| 135 | 62.09 | 39.07 | 43.06 | 3.98 | 0.0203 | 0.08 |
| 136 | 64.03 | 40.04 | 42.04 | 2.00 | 0.0101 | 0.02 |
| 137 | 61.90 | 40.99 | 40.99 | 0.00 |  | b=110.64 |
| 138 | 77.00 | 42.04 | 40.04 | -2.00 |  | $\mathrm{S}=19.97$ |
| 139 | 42.04 | 43.06 | 39.07 | -3.98 |  |  |
| 140 | 56.07 | 45.01 | 38.92 | -6.09 |  |  |
| 141 | 30.04 | 45.08 | 37.86 | -7.21 |  | $\mathrm{W}=0.808<0.939$ |
| 142 | 30.89 | 46.98 | 36.98 | -10.00 |  |  |
| 143 | 35.90 | 47.11 | 35.98 | -11.13 |  |  |
| 144 | 43.06 | 51.94 | 35.90 | -16.04 |  | Did not Satisfied |
| 145 | 35.98 | 56.07 | 34.87 | -21.20 |  |  |
| 146 | 39.07 | 57.08 | 31.96 | -25.12 |  |  |
| 147 | 22.05 | 57.12 | 31.89 | -25.23 |  |  |
| 148 | 57.08 | 57.97 | 30.96 | -27.01 |  |  |
| 149 | 57.97 | 57.97 | 30.89 | -27.08 |  |  |
| 150 | 45.08 | 58.06 | 30.13 | -27.93 |  |  |
| 151 | 30.13 | 59.10 | 30.04 | -29.06 |  |  |
| 152 | 37.86 | 61.90 | 29.97 | -31.93 |  |  |
| 153 | 34.87 | 62.09 | 25.07 | -37.02 |  |  |
| 154 | 58.06 | 64.03 | 24.98 | -39.05 |  |  |
| 155 | 18.93 | 77.00 | 22.05 | -54.96 |  |  |
| 156 | 40.04 | 135.00 | 18.93 | -116.07 |  |  |

### 3.5.2.3 order of (AR)

For water quality like King Talal Dam, the value of AR, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of $\mathrm{BOD}_{5}$ does not need more than 1 month till it analyze (Viessman and Lewis, 1996). From Figure (21) it can be seen that the value of AR is more than 1 , but the value of p that will be used is 1 for the $\mathrm{BOD}_{5}$ variable.


Figure (21) Autocorrelation Function for $\mathrm{BOD}_{5}$ Variable

### 3.5.2.4 order of moving average (MA)

After finding the value of AR, which was 1 , the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (22) shows the change between the real data of the variable $\mathrm{BOD}_{5}$ and it's moving average with different lengths of p .

| Mbing Average | MbingAcrage |  |
| :---: | :---: | :---: |
|  |  |  |
| MbingAverage | MoingArage |  |
| Mouing Average | MbingAerage |  |

Figure (22) Moving Average of $\mathrm{BOD}_{5}$ with Different Values of (p)

The moving average can be determined from Figure (22) when the difference between the previous length of $p$ and the followed one have a small difference and that occurred when the value of p was 3 ( as shown in Figure (22) ), so the BOD5 variable has a value of MA(3).

### 3.5.2.5 order of (I)

The last coefficient of ARIMA's parameters is the integrated model ( I ), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (23) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that they are almost the same, which means that the detrended and seasonally effects are almost negligible.


Figure (23): Component Analysis for BOD5 mg/l

Two season; summer and winter can affect seasonality in Jordan, so if the data has no seasonality effect, then the value of $\mathrm{d}=0$ and if we have seasonality effect then the value of $\mathrm{d}=2$. Figure (24) provides ARIMA model diagnostics for ARIMA $=$ $(1,0,3)$, but for ARIMA $=(1,2,3)$ the program could not draw it since it gives singular matrix. It can be concluded that ARIMA $(1,0,3)$ gives the best regression.


Figure (24): ARIMA (1,0,3) Model Diagnostics for $\mathrm{BOD}_{5}$

### 3.5.2.6 forecasting future values

The following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last $10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (25).


Figure(25): Trend Analysis for BOD5 mg/l
It can be observed from the above figure and equation of the linear trend that the data is increasing slowly. Table (23) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (23): The values of the predicted and actual data by linear regression for $\mathrm{BOD}_{5}$ variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 60.19 | 30.04 |
| 2 | 142 | 60.49 | 30.89 |
| 3 | 143 | 60.78 | 35.90 |
| 4 | 144 | 61.08 | 43.06 |
| 5 | 145 | 61.37 | 35.98 |
| 6 | 146 | 61.67 | 39.07 |
| 7 | 147 | 61.96 | 22.05 |
| 8 | 148 | 62.26 | 57.08 |
| 9 | 149 | 62.55 | 57.97 |
| 10 | 150 | 62.85 | 45.08 |
| 11 | 151 | 63.14 | 30.13 |
| 12 | 152 | 63.44 | 37.86 |


| 13 | 153 | 63.73 | 34.87 |
| :--- | :--- | :--- | :--- |
| 14 | 154 | 64.03 | 58.06 |
| 15 | 155 | 64.32 | 18.93 |
| 16 | 156 | 64.62 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $38.2 \%$, that the linear trend model did not satisfy the forecasting for the BOD5 variable.

## A2- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (26).


It can be observed from the above figure and the equation of the quadratic trend that the data is increasing upward slowly. Table (24) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (24): The values of the predicted and actual data by quadratic regression for $\mathrm{BOD}_{5}$ variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 59.93 | 30.04 |
| 2 | 142 | 60.21 | 30.89 |
| 3 | 143 | 60.49 | 35.90 |
| 4 | 144 | 60.78 | 43.06 |


| 5 | 145 | 61.06 | 35.98 |
| :---: | :---: | :---: | :---: |
| 6 | 146 | 61.34 | 39.07 |
| 7 | 147 | 61.63 | 22.05 |
| 8 | 148 | 61.91 | 57.08 |
| 9 | 149 | 62.19 | 57.97 |
| 10 | 150 | 62.47 | 45.08 |
| 11 | 151 | 62.76 | 30.13 |
| 12 | 152 | 63.04 | 37.86 |
| 13 | 153 | 63.32 | 34.87 |
| 14 | 154 | 63.60 | 58.06 |
| 15 | 155 | 63.88 | 18.93 |
| 16 | 156 | 64.16 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $37.9 \%$, that the quadratic trend model did not satisfy the forecasting for the BOD5 variable.

## A3- exponential growth regression model

The regression of the additive exponential growth trend model is shown in Figure (27).

Figure(27): Trend Analysis for BOD5 mg/l
Growth Curve Model $\mathrm{Yt}=19.4234^{\star}\left(1.00811^{* *} \mathrm{t}\right)$


It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing trend. Table (25) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (25): The values of the predicted and actual data by exponential growth regression for $\mathrm{BOD}_{5}$ variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 71.01 | 30.04 |
| 2 | 142 | 71.14 | 30.89 |
| 3 | 143 | 71.26 | 35.90 |
| 4 | 144 | 71.38 | 43.06 |
| 5 | 145 | 71.51 | 35.98 |
| 6 | 146 | 71.63 | 39.07 |
| 7 | 147 | 71.76 | 22.05 |
| 8 | 148 | 71.88 | 57.08 |
| 9 | 149 | 72.01 | 57.97 |
| 10 | 150 | 72.13 | 45.08 |
| 11 | 151 | 72.26 | 30.13 |
| 12 | 152 | 72.38 | 37.86 |
| 13 | 153 | 72.51 | 34.87 |
| 14 | 154 | 72.64 | 58.06 |
| 15 | 155 | 72.76 | 18.93 |
| 16 | 156 | 72.89 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $46.4 \%$, that the exponential growth trend model did not satisfy the forecasting for the BOD5 variable.

## A4- single exponential smoothing model

The regression of the additive single exponential smoothing trend model is shown in Figure (28).


Figure(28): Single Exponential Smoothing for BOD5 mg/l

Table (26) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (26) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (26): Forecasted, lower, upper and actual values by single exponential smoothing for $\mathrm{BOD}_{5}$ variable

| $\underline{\text { Row }}$ | Period <br> $(\mathrm{month})$ | $\frac{\text { Forecast }}{\mathrm{mg} / \mathrm{l}}$ | $\frac{\text { Lower }}{\mathrm{mg} / \mathrm{l}}$ | Upper <br> $\mathrm{mg} / \mathrm{l}$ | $\underline{\text { Actual }}$ <br> 1 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 141 | 53.91 | 23.34 | 84.48 | 30.04 |
| 2 | 142 | 53.91 | 23.34 | 84.48 | 30.89 |
| 3 | 143 | 53.91 | 23.34 | 84.48 | 35.90 |
| 4 | 144 | 53.91 | 23.34 | 84.48 | 43.06 |
| 5 | 145 | 53.91 | 23.34 | 84.48 | 35.98 |
| 6 | 146 | 53.91 | 23.34 | 84.48 | 39.07 |
| 7 | 147 | 53.91 | 23.34 | 84.48 | 22.05 |
| 8 | 148 | 53.91 | 23.34 | 84.48 | 57.08 |
| 9 | 149 | 53.91 | 23.34 | 84.48 | 57.97 |
| 10 | 150 | 53.91 | 23.34 | 84.48 | 45.08 |
| 11 | 151 | 53.91 | 23.34 | 84.48 | 30.13 |
| 12 | 152 | 53.91 | 23.34 | 84.48 | 37.86 |
| 13 | 153 | 53.91 | 23.34 | 84.48 | 34.87 |
| 14 | 154 | 53.91 | 23.34 | 84.48 | 58.06 |


| 15 | 155 | 53.91 | 23.34 | 84.48 | 18.93 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 16 | 156 | 53.91 | 23.34 | 84.48 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $28.5 \%$, that the simple exponential smoothing trend model has not satisfied the forecasting for the BOD5 variable.

## B- stochastic forecasting

## B1- auto regression model

Table (27) shows the $\mathrm{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (27) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (27): Forecasted, lower, upper and actual values by $\mathrm{AR}(1)$ for $\mathrm{BOD}_{5}$ variable

| Row | Period (month) | Forecast mg/l | $\frac{\text { Lower }}{\mathrm{mg} / \mathrm{l}}$ | Upper mg/l | $\frac{\text { Actual }}{\mathrm{mg} / \mathrm{l}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 82.95 | 36.96 | 128.93 | 30.04 |
| 2 | 142 | 73.92 | 24.57 | 123.28 | 30.89 |
| 3 | 143 | 70.41 | 20.57 | 120.25 | 35.90 |
| 4 | 144 | 69.04 | 19.13 | 118.96 | 43.06 |
| 5 | 145 | 68.51 | 18.59 | 118.44 | 35.98 |
| 6 | 146 | 68.30 | 18.38 | 118.23 | 39.07 |
| 7 | 147 | 68.22 | 18.30 | 118.15 | 22.05 |
| 8 | 148 | 68.19 | 18.26 | 118.12 | 57.08 |
| 9 | 149 | 68.18 | 18.25 | 118.11 | 57.97 |
| 10 | 150 | 68.18 | 18.25 | 118.10 | 45.08 |
| 11 | 151 | 68.17 | 18.25 | 118.10 | 30.13 |
| 12 | 152 | 68.17 | 18.25 | 118.10 | 37.86 |
| 13 | 153 | 68.17 | 18.24 | 118.10 | 34.87 |
| 14 | 154 | 68.17 | 18.24 | 118.10 | 58.06 |
| 15 | 155 | 68.17 | 18.24 | 118.10 | 18.93 |
| 16 | 156 | 68.17 | 18.24 | 118.10 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $44.7 \%$, that the $\operatorname{AR}(1)$ trend model has not satisfied the forecasting for the BOD5 variable.

## B2- moving average regression model

The regression of the additive MA (3) trend model is shown in Figure (29).


Figure (29): Moving Average Trend for BOD5 mg/l

Table (28) shows the MA(3) prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (28) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (28): Forecasted, lower, upper and actual values by MA (3) for $\mathrm{BOD}_{5}$ variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | $\underline{\mathrm{mg} / \mathrm{l}}$ | mg/l | $\mathrm{mg} / \mathrm{l}$ | mg/l |
| 1 | 141 | 58.37 | 22.55 | 94.19 | 30.04 |
| 2 | 142 | 58.37 | 22.55 | 94.19 | 30.89 |
| 3 | 143 | 58.37 | 22.55 | 94.19 | 35.90 |
| 4 | 144 | 58.37 | 22.55 | 94.19 | 43.06 |
| 5 | 145 | 58.37 | 22.55 | 94.19 | 35.98 |
| 6 | 146 | 58.37 | 22.55 | 94.19 | 39.07 |
| 7 | 147 | 58.37 | 22.55 | 94.19 | 22.05 |


| 8 | 148 | 58.37 | 22.55 | 94.19 | 57.08 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 9 | 149 | 58.37 | 22.55 | 94.19 | 57.97 |
| 10 | 150 | 58.37 | 22.55 | 94.19 | 45.08 |
| 11 | 151 | 58.37 | 22.55 | 94.19 | 30.13 |
| 12 | 152 | 58.37 | 22.55 | 94.19 | 37.86 |
| 13 | 153 | 58.37 | 22.55 | 94.19 | 34.87 |
| 14 | 154 | 58.37 | 22.55 | 94.19 | 58.06 |
| 15 | 155 | 58.37 | 22.55 | 94.19 | 18.93 |
| 16 | 156 | 58.37 | 22.55 | 94.19 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $33.9 \%$, that the $\mathrm{MA}(3)$ trend model has not satisfied the forecasting for the BOD5 variable.

## B3- ARMA modeling

Table (29) shows the $\operatorname{ARMA}(1,3)$ prediction values for the next $10 \%$ of the predicted and the real data, which equals to 16 observations. In Table (29) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (29): Forecasted, lower, upper and actual values by ARMA $(1,3)$ for $\mathrm{BOD}_{5}$ variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | mg/l | mg/l | mg/l | mg/l |
| 1 | 141 | 50.48 | 17.07 | 83.89 | 30.04 |
| 2 | 142 | 51.34 | 17.49 | 85.18 | 30.89 |
| 3 | 143 | 50.82 | 16.23 | 85.41 | 35.90 |
| 4 | 144 | 50.56 | 15.67 | 85.45 | 43.06 |
| 5 | 145 | 50.30 | 15.13 | 85.47 | 35.98 |
| 6 | 146 | 50.05 | 14.61 | 85.48 | 39.07 |
| 7 | 147 | 49.80 | 14.10 | 85.49 | 22.05 |
| 8 | 148 | 49.55 | 13.61 | 85.49 | 57.08 |
| 9 | 149 | 49.31 | 13.14 | 85.49 | 57.97 |
| 10 | 150 | 49.08 | 12.68 | 85.48 | 45.08 |
| 11 | 151 | 48.85 | 12.24 | 85.46 | 30.13 |
| 12 | 152 | 48.62 | 11.81 | 85.44 | 37.86 |
| 13 | 153 | 48.40 | 11.39 | 85.41 | 34.87 |
| 14 | 154 | 48.19 | 10.99 | 85.39 | 58.06 |
| 15 | 155 | 47.97 | 10.59 | 85.35 | 18.93 |
| 16 | 156 | 47.76 | 10.21 | 85.31 | 40.04 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $16.1 \%$, that the $\operatorname{ARMA}(1,3)$ trend model has satisfied the forecasting for the BOD5 variable.

### 3.5.2.7 results of prediction

The results of error are summarized in the following Table (30), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (30) : Percentage of error of each model for $\mathrm{BOD}_{5}$ variable

| Model | Percentage of <br> Mean Error |
| :--- | :---: |
| Linear Method | $38.2 \%$ |
| Quadratic Method | $37.9 \%$ |
| Exponential Growth Method | $46.4 \%$ |
| Simple Exponential Smoothing | $28.5 \%$ |
| Auto Regression, AR(1) | $44.7 \%$ |
| Moving Average, MA(3) | $33.9 \%$ |
| ARMA (1, 3) | $16.1 \%$ |

The previous Table (30) shows that the methods have not satisfied the $10 \%$ acceptable prediction limits. When finding the best model that gave the least error it will be ARMA $(1,3)$.

### 3.5.3 Chemical oxygen demand (COD) variable:

The consequences that were used to analyze the COD variable were as follows:

### 3.5.3.1 detection of missing data and outliers:

From the table(1) it is observed that the data do not contain any missing data, so the second step is to find the outliers, data should be drawn in a scatter diagram (Figure 30) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have approximately five outliers and they are in the following months: March 1988, November 1996, May 1999, June 1999, and August 2000. It was observed that the rainfall in December was high, and it is known that when the rainfall is high then the COD will get low, in November 1996 the rainfall was high, May and June 1999 the rainfall was low so the COD should be high and finally in August 2000 the rainfall was low so the COD should be high (Appendix (1)). So the real data are in March 1988, May 1999, June 1999, and August 2000, the other one data (November 1996) was assumed to be outliers due to human error, and it should be adjusted to a new value since it may greatly influence any statistical calculations and yield biased results. The way that outliers were adjusted was the same as the missing data treated and it was equals to the average monthly value.


Figure (30) Original Data of COD (mg/l)

Figure (31) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are ten outliers in both the original data and the residual data in the seasonal condition. Also Figure (31) shows the variation in the data for the same month, it can be observe that the variation was the highest on December, and was the lowest on October. Another six outliers were found in the seasonal drawings, they are in March 1997, April 1999, July 1999, February 2000, April 2000, and July 2000. There was high rainfall in March 1997 and February 2000 (Appendix (1)) so they should be adjusted, while in the other months they were real data.


Figure (31) Outliers Seasonal Analysis for COD Variable

After adjustment the outliers, the new adjusted data are plotted in Figure (32), the figure shows that their still outliers but these outliers cannot be omitted because they are real data so it can influence the statistics results. While comparing the old data (Figure 30) with the new adjusted data (Figure 32) it can be observed that two figures are quite the same and they have the same trend, so the effect of the outliers on the data was so little.


Figure (32): The New Adjusted Data for COD(mg/l)

### 3.5.3.2 normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the COD variable; the calculated values were as follows

| Month | Jan. | Feb. | Mar. | Apr. | May | Jun. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| COD mg/l | 111.9 | 84.2 | 83.6 | 109.0 | 117.0 | 138.7 |
| Month | Jul. | Aug. | Sep. | Oct. | Nov. | Dec. |
| COD mg/l | 124.3 | 126.4 | 138.5 | 121.0 | 127.1 | 134.0 |

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from figure (33) that the data of COD is quite normal and there is a little skewness to the left and bulked to the right, but in general the graph gives an indication that the data is normal.


Figure (33) : Weibull Distribution Model Histogram

## B- coefficient of variation (COV), preliminary est:

The data were divided into four quarters; each quarter consists of 39 data. Table (31) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the COD variable.

Table (31): The coefficient of variable for COD

|  | MEAN |  | ST. DE. | C.O.V. |
| :--- | :--- | :---: | :---: | :---: | :---: |
| COD $\quad\left(1^{\text {st }}\right.$ Quarter $)$ | 93.6 | 977.8 | 31.3 | 0.3 |
| COD $\quad\left(2^{\text {nd }}\right.$ Quarter $)$ | 93.7 | 725.1 | 26.9 | 0.3 |
| COD $\quad\left(3^{\text {rd }}\right.$ Quarter $)$ | 128.5 | 2122.9 | 46.1 | 0.4 |
| COD $\quad\left(4^{\text {th }}\right.$ Quarter $)$ | 166.7 | 4045.4 | 63.6 | 0.4 |

It can be shown from the table that the value of the coefficient of variation for each quarter is less than 1 , which means that each quarter of the data has a little skewed (either to right or left), so the total data of the COD variable has less
skewness than each of the four COD quarters, it can be concluded that the COD variable does not have skewness.

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K , which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (32) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of $K$, and the Kurtosis coefficient.

Table (32): The Kurtosis Coefficient for COD

|  | MEAN | VARIANCE | ST. DE. <br> $(\mathrm{S})$ | K | Kurtosis <br> Coeff. <br> $\mathrm{C}_{\mathrm{K}}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| COD ( 1st Quarter ) | 93.6 | 977.8 | 31.3 | 2939949.6 | 0.1 |
| COD ( 2nd Quarter ) | 93.7 | 725.1 | 26.9 | 45856.6 | -2.9 |
| COD ( 3rd Quarter ) | 128.5 | 2122.9 | 46.1 | 23428503.4 | 2.2 |
| COD ( 4th Quarter ) | 166.7 | 4045.4 | 63.6 | 39728843.7 | -0.6 |

From table (32) one can observe that the data in the first and fourth quarters were normally distributed (mesokurtic), in the second quarter it was fairly platykutric and in the third quarter is was fairly leptokurtic. The total data of the COD variable can be assumed to be as fairly normally distributed (mesokurtic).

## D- Shapiro-Wilk Test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of ( $a_{n-1+1}$ ), was taken for 20 data since the value of $n-1+i$ was equal to 20 , the value of $\left(a_{n-1+1}\right)$ was taken from appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

From the Tables (33), (34), (35), and (36) it has been shown that the data in the first, the second, and the fourth quarters are normal, the third has a quite not normal one. It can be safely say that COD variable is normally distributed.

Table (33): Shapiro-Wilk Test for the Data of COD's 1 ${ }^{\text {st }}$ quarter


Table (34): Shapiro-Wilk Test for the Data of COD's $2^{\text {nd }}$ quarter

| No | $\begin{aligned} & \text { COD } \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | Ordering COD <br> (1) | Inverse Order COD <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{I})$ | I*J |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 78.00 | 35.00 | 160.99 | 125.99 | 0.3989 | 50.26 |  |
| 41 | 88.01 | 51.00 | 135.01 | 84.00 | 0.2755 | 23.14 |  |
| 42 | 95.00 | 54.00 | 129.01 | 75.00 | 0.2380 | 17.85 |  |
| 43 | 88.98 | 54.00 | 127.99 | 73.99 | 0.2104 | 15.57 |  |
| 44 | 116.01 | 59.00 | 127.00 | 68.00 | 0.1880 | 12.78 |  |
| 45 | 126.87 | 60.00 | 126.87 | 66.87 | 0.1689 | 11.29 |  |
| 46 | 92.01 | 67.00 | 123.00 | 56.00 | 0.1520 | 8.51 |  |
| 47 | 115.99 | 73.00 | 122.00 | 49.00 | 0.1366 | 6.69 |  |
| 48 | 35.00 | 77.00 | 116.01 | 39.02 | 0.1225 | 4.78 |  |
| 49 | 51.00 | 78.00 | 116.01 | 38.01 | 0.1092 | 4.15 |  |
| 50 | 59.00 | 83.00 | 116.00 | 33.00 | 0.0967 | 3.19 |  |
| 51 | 67.00 | 84.00 | 115.99 | 31.98 | 0.0848 | 2.71 |  |
| 52 | 60.00 | 85.00 | 98.99 | 13.99 | 0.0733 | 1.03 |  |
| 53 | 73.00 | 85.99 | 97.00 | 11.00 | 0.0622 | 0.68 |  |
| 54 | 85.00 | 86.00 | 96.99 | 11.00 | 0.0515 | 0.57 |  |
| 55 | 85.99 | 86.98 | 95.00 | 8.02 | 0.0409 | 0.33 |  |
| 56 | 86.98 | 88.00 | 93.00 | 4.99 | 0.0305 | 0.15 |  |
| 57 | 93.00 | 88.01 | 92.01 | 4.00 | 0.0203 | 0.08 |  |
| 58 | 98.99 | 88.98 | 91.00 | 2.01 | 0.0101 | 0.02 |  |
| 59 | 77.00 | 90.00 | 90.00 | 0.00 |  | $\mathrm{b}=$ | 163.79 |
| 60 | 54.00 | 91.00 | 88.98 | -2.01 |  | $\mathrm{S}=$ | 26.93 |
| 61 | 54.00 | 92.01 | 88.01 | -4.00 |  |  |  |
| 62 | 84.00 | 93.00 | 88.00 | -4.99 |  | W=0 | -0.939 |
| 63 | 91.00 | 95.00 | 86.98 | -8.02 |  | W=0. | > 0.939 |
| 64 | 97.00 | 96.99 | 86.00 | -11.00 |  |  |  |
| 65 | 116.01 | 97.00 | 85.99 | -11.00 |  |  | Satisfied |
| 66 | 135.01 | 98.99 | 85.00 | -13.99 |  |  | Satisfied |
| 67 | 116.00 | 115.99 | 84.00 | -31.98 |  |  |  |
| 68 | 96.99 | 116.00 | 83.00 | -33.00 |  |  |  |
| 69 | 129.01 | 116.01 | 78.00 | -38.01 |  |  |  |
| 70 | 160.99 | 116.01 | 77.00 | -39.02 |  |  |  |
| 71 | 122.00 | 122.00 | 73.00 | -49.00 |  |  |  |
| 72 | 83.00 | 123.00 | 67.00 | -56.00 |  |  |  |
| 73 | 86.00 | 126.87 | 60.00 | -66.87 |  |  |  |
| 74 | 88.00 | 127.00 | 59.00 | -68.00 |  |  |  |
| 75 | 90.00 | 127.99 | 54.00 | -73.99 |  |  |  |
| 76 | 127.00 | 129.01 | 54.00 | -75.00 |  |  |  |
| 77 | 127.99 | 135.01 | 51.00 | -84.00 |  |  |  |
| 78 | 123.00 | 160.99 | 35.00 | -125.99 |  |  |  |

Table (35): Shapiro-Wilk Test for the Data of COD's $3^{\text {rd }}$ quarter

| No | $\begin{aligned} & \mathrm{COD} \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | Ordering COD <br> (1) | Inverse Order COD (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{I})$ | I* J |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 79 | 120.98 | 39.00 | 275.02 | 236.02 | 0.3989 | 94.15 |
| 80 | 149.01 | 72.03 | 231.98 | 159.95 | 0.2755 | 44.07 |
| 81 | 115.99 | 79.08 | 223.02 | 143.94 | 0.2380 | 34.26 |
| 82 | 166.00 | 83.06 | 188.07 | 105.02 | 0.2104 | 22.10 |
| 83 | 133.00 | 88.00 | 181.05 | 93.04 | 0.1880 | 17.49 |
| 84 | 39.00 | 90.07 | 174.92 | 84.85 | 0.1689 | 14.33 |
| 85 | 101.02 | 91.93 | 166.00 | 74.07 | 0.1520 | 11.26 |
| 86 | 83.06 | 93.20 | 156.03 | 62.83 | 0.1366 | 8.58 |
| 87 | 88.00 | 94.99 | 154.97 | 59.98 | 0.1225 | 7.35 |
| 88 | 94.99 | 101.02 | 149.01 | 47.99 | 0.1092 | 5.24 |
| 89 | 103.99 | 103.99 | 147.08 | 43.09 | 0.0967 | 4.17 |
| 90 | 113.91 | 104.01 | 143.08 | 39.07 | 0.0848 | 3.31 |
| 91 | 140.04 | 104.05 | 140.04 | 35.99 | 0.0733 | 2.64 |
| 92 | 154.97 | 107.07 | 133.00 | 25.93 | 0.0622 | 1.61 |
| 93 | 174.92 | 110.05 | 127.92 | 17.87 | 0.0515 | 0.92 |
| 94 | 147.08 | 113.91 | 127.02 | 13.11 | 0.0409 | 0.54 |
| 95 | 143.08 | 114.05 | 120.98 | 6.94 | 0.0305 | 0.21 |
| 96 | 104.01 | 114.05 | 120.93 | 6.88 | 0.0203 | 0.14 |
| 97 | 114.05 | 115.07 | 117.04 | 1.96 | 0.0101 | 0.02 |
| 98 | 117.04 | 115.99 | 115.99 | 0.00 |  | $\mathrm{b}=\quad 272.38$ |
| 99 | 114.05 | 117.04 | 115.07 | -1.96 |  | $\mathrm{S}=\quad 46.07$ |
| 100 | 91.93 | 120.93 | 114.05 | -6.88 |  |  |
| 101 | 127.02 | 120.98 | 114.05 | -6.94 |  | $\mathrm{W}=0.920<0.939$ |
| 102 | 110.05 | 127.02 | 113.91 | -13.11 |  | W $=0.920<0.939$ |
| 103 | 107.07 | 127.92 | 110.05 | -17.87 |  |  |
| 104 | 127.92 | 133.00 | 107.07 | -25.93 |  | Satisfied |
| 105 | 181.05 | 140.04 | 104.05 | -35.99 |  |  |
| 106 | 120.93 | 143.08 | 104.01 | -39.07 |  |  |
| 107 | 275.02 | 147.08 | 103.99 | -43.09 |  |  |
| 108 | 188.07 | 149.01 | 101.02 | -47.99 |  |  |
| 109 | 231.98 | 154.97 | 94.99 | -59.98 |  |  |
| 110 | 72.03 | 156.03 | 93.20 | -62.83 |  |  |
| 111 | 223.02 | 166.00 | 91.93 | -74.07 |  |  |
| 112 | 115.07 | 174.92 | 90.07 | -84.85 |  |  |
| 113 | 79.08 | 181.05 | 88.00 | -93.04 |  |  |
| 114 | 156.03 | 188.07 | 83.06 | -105.02 |  |  |
| 115 | 104.05 | 223.02 | 79.08 | -143.94 |  |  |
| 116 | 93.20 | 231.98 | 72.03 | -159.95 |  |  |
| 117 | 90.07 | 275.02 | 39.00 | -236.02 |  |  |

Table (36): Shapiro-Wilk Test for the Data of COD's $4^{\text {th }}$ quarter

| No | $\begin{aligned} & \mathrm{COD} \\ & \mathrm{mg} / \mathrm{l} \end{aligned}$ | Ordering COD <br> (1) | Inverse order COD (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{I})$ | I*J |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 116.08 | 68.03 | 302.96 | 234.93 | 0.3989 | 93.71 |  |
| 119 | 147.07 | 86.04 | 298.96 | 212.92 | 0.2755 | 58.66 |  |
| 120 | 302.96 | 88.00 | 268.11 | 180.11 | 0.2380 | 42.87 |  |
| 121 | 68.03 | 88.96 | 258.93 | 169.97 | 0.2104 | 35.76 |  |
| 122 | 126.96 | 93.96 | 255.94 | 161.98 | 0.1880 | 30.45 |  |
| 123 | 99.96 | 94.90 | 230.89 | 135.99 | 0.1689 | 22.97 |  |
| 124 | 88.00 | 99.96 | 229.93 | 129.97 | 0.1520 | 19.76 |  |
| 125 | 93.96 | 102.97 | 226.00 | 123.03 | 0.1366 | 16.81 |  |
| 126 | 184.93 | 103.95 | 219.02 | 115.07 | 0.1225 | 14.10 |  |
| 127 | 88.96 | 104.00 | 207.14 | 103.14 | 0.1092 | 11.26 |  |
| 128 | 103.95 | 106.91 | 206.99 | 100.08 | 0.0967 | 9.68 |  |
| 129 | 94.90 | 111.11 | 205.97 | 94.86 | 0.0848 | 8.04 |  |
| 130 | 106.91 | 116.08 | 203.06 | 86.98 | 0.0733 | 6.38 |  |
| 131 | 111.11 | 126.96 | 194.00 | 67.05 | 0.0622 | 4.17 |  |
| 132 | 152.00 | 135.95 | 184.93 | 48.98 | 0.0515 | 2.52 |  |
| 133 | 203.06 | 136.06 | 184.08 | 48.02 | 0.0409 | 1.96 |  |
| 134 | 86.04 | 147.07 | 180.98 | 33.91 | 0.0305 | 1.03 |  |
| 135 | 104.00 | 152.00 | 177.87 | 25.87 | 0.0203 | 0.53 |  |
| 136 | 194.00 | 162.94 | 176.07 | 13.13 | 0.0101 | 0.13 |  |
| 137 | 255.94 | 164.98 | 164.98 | 0.00 |  | $\mathrm{b}=$ | 380.79 |
| 138 | 268.11 | 176.07 | 162.94 | -13.13 |  | $\mathrm{S}=$ | 63.60 |
| 139 | 230.89 | 177.87 | 152.00 | -25.87 |  | $\mathrm{W}=0.943>0.939$ |  |
| 140 | 177.87 | 180.98 | 147.07 | -33.91 |  |  |  |
| 141 | 207.14 | 184.08 | 136.06 | -48.02 |  |  |  |
| 142 | 102.97 | 184.93 | 135.95 | -48.98 |  |  |  |
| 143 | 176.07 | 194.00 | 126.96 | -67.05 |  | Satisfied |  |
| 144 | 206.99 | 203.06 | 116.08 | -86.98 |  |  |  |
| 145 | 180.98 | 205.97 | 111.11 | -94.86 |  |  |  |
| 146 | 219.02 | 206.99 | 106.91 | -100.08 |  |  |  |
| 147 | 164.98 | 207.14 | 104.00 | -103.14 |  |  |  |
| 148 | 229.93 | 219.02 | 103.95 | -115.07 |  |  |  |
| 149 | 136.06 | 226.00 | 102.97 | -123.03 |  |  |  |
| 150 | 184.08 | 229.93 | 99.96 | -129.97 |  |  |  |
| 151 | 205.97 | 230.89 | 94.90 | -135.99 |  |  |  |
| 152 | 258.93 | 255.94 | 93.96 | -161.98 |  |  |  |
| 153 | 298.96 | 258.93 | 88.96 | -169.97 |  |  |  |
| 154 | 162.94 | 268.11 | 88.00 | -180.11 |  |  |  |
| 155 | 135.95 | 298.96 | 86.04 | -212.92 |  |  |  |
| 156 | 226.00 | 302.96 | 68.03 | -234.93 |  |  |  |

### 3.5.3.3 order of (AR)

For water quality like King Talal Dam, the value of $A R$, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of COD does not need more than 1 month till it analyze (Viessman and Lewis, 1996). From Figure (34) it can be seen that the value of AR is 1 , the value of $p$ that will be used is 1 for the COD variable.


Figure (34) Autocorrelation Function for COD Variable

### 3.5.3.4 order of moving average (MA)

After finding the value of AR, which was 1 , the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (35) shows the change between the real data of the variable COD and it's moving average with different lengths of p .


Figure (35) Moving Average of COD with Different Values of (p)

The moving average can be determined from Figure (35) when the difference between the previous length of p and the followed one have a small difference and that occurred when the value of p was 4 (as shown in Figure (35)), so the COD variable has a value of MA(4).

### 3.5.3.5 order of (I)

The last coefficient of ARIMA's parameters is the integrated model ( I ), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (36) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that there are a difference between the original figure and the detrended one but in the seasonal case they are almost the same, which means that the detrended effect could take into consideration.

Figure (36): Component Analysis for COD mg/l


Two season; summer and winter can affect seasonality in Jordan, so if the data has no trend effect, then the value of $d=0$ and if we have trend effect then the value of $\mathrm{d}=2$. Figures (37) and (38) provides ARIMA model diagnostics for ARIMA $=(1,0,4)$ and for ARIMA $=(1,2,4)$. The two figures indicate the same results. It can be concluded that the data has trend effect so $\operatorname{ARIMA}(1,2,4)$ should be used.


Figure (37): ARIMA (1,0,4) Diagnostics for COD


Figure (38): ARIMA (1,2,4) Diagnostics for COD

### 3.5.3.6 forecasting future values

The following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last $10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (39).


Figure (39): Trend Analysis for COD mg/l

It can be observed from the above figure and equation of the linear trend that the data is increasing. Table (37) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (37): The values of the predicted and actual data by linear regression for COD variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 150.81 | 207.14 |
| 2 | 142 | 151.36 | 102.97 |
| 3 | 143 | 151.90 | 176.07 |
| 4 | 144 | 152.45 | 206.99 |
| 5 | 145 | 153.00 | 180.98 |
| 6 | 146 | 153.54 | 84.18 |
| 7 | 147 | 154.09 | 164.98 |
| 8 | 148 | 154.64 | 229.93 |
| 9 | 149 | 155.18 | 136.06 |
| 10 | 150 | 155.73 | 184.08 |
| 11 | 151 | 156.27 | 205.97 |
| 12 | 152 | 156.82 | 258.93 |
| 13 | 153 | 157.37 | 298.96 |
| 14 | 154 | 157.91 | 162.94 |
| 15 | 155 | 158.46 | 135.95 |
| 16 | 156 | 159.00 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $19.5 \%$, that the linear trend model did not satisfy the forecasting for the COD variable.

## A2- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (40).


It can be observed from the above figure and the equation of the quadratic trend that the data is increasing upward. Table (38) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (38): The values of the predicted and actual data by quadratic regression for COD variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 168.35 | 207.14 |
| 2 | 142 | 169.64 | 102.97 |
| 3 | 143 | 170.95 | 176.07 |
| 4 | 144 | 172.26 | 206.99 |
| 5 | 145 | 173.58 | 180.98 |
| 6 | 146 | 174.92 | 84.18 |
| 7 | 147 | 176.26 | 164.98 |
| 8 | 148 | 177.62 | 229.93 |
| 9 | 149 | 178.98 | 136.06 |
| 10 | 150 | 180.36 | 184.08 |
| 11 | 151 | 181.75 | 205.97 |
| 12 | 152 | 183.15 | 258.93 |
| 13 | 153 | 184.55 | 298.96 |
| 14 | 154 | 185.97 | 162.94 |
| 15 | 155 | 187.40 | 135.95 |
| 16 | 156 | 188.84 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $3.8 \%$, that the quadratic trend model has satisfied the forecasting for the COD variable.

## A3- exponential growth regression model

The regression of the additive exponential growth trend model is shown in Figure (41).


Figure (41): Trend Analysis for COD mg/l

It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing trend. Table (39) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (39): The values of the predicted and actual data by exponential growth regression for COD variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 71.01 | 207.14 |
| 2 | 142 | 71.14 | 102.97 |
| 3 | 143 | 71.26 | 176.07 |


| 4 | 144 | 71.38 | 206.99 |
| :--- | :---: | :---: | :---: |
| 5 | 145 | 71.51 | 180.98 |
| 6 | 146 | 71.63 | 84.18 |
| 7 | 147 | 71.76 | 164.98 |
| 8 | 148 | 71.88 | 229.93 |
| 9 | 149 | 72.01 | 136.06 |
| 10 | 150 | 72.13 | 184.08 |
| 11 | 151 | 72.26 | 205.97 |
| 12 | 152 | 72.38 | 258.93 |
| 13 | 153 | 72.51 | 298.96 |
| 14 | 154 | 72.64 | 162.94 |
| 15 | 155 | 72.76 | 135.95 |
| 16 | 156 | 72.89 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $23.8 \%$, that the exponential growth trend model did not satisfy the forecasting for the COD variable.

## A4- single exponential smoothing model

The regression of the additive single exponential smoothing trend model is shown in Figure (42).


Figure (42): Single Exponential Smoothing for COD mg/l

Table (40) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (40) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (40): Forecasted, lower, upper and actual values by single exponential smoothing for COD variable

| Row | $\frac{\text { Period }}{(\text { month })}$ | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Lower }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Upper }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Actual }}{\underline{\mathrm{m}} / \mathrm{l}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 199.22 | 125.62 | 272.82 | 207.14 |
| 2 | 142 | 199.22 | 125.62 | 272.82 | 102.97 |
| 3 | 143 | 199.22 | 125.62 | 272.82 | 176.07 |
| 4 | 144 | 199.22 | 125.62 | 272.82 | 206.99 |
| 5 | 145 | 199.22 | 125.62 | 272.82 | 180.98 |
| 6 | 146 | 199.22 | 125.62 | 272.82 | 84.18 |
| 7 | 147 | 199.22 | 125.62 | 272.82 | 164.98 |
| 8 | 148 | 199.22 | 125.62 | 272.82 | 229.93 |
| 9 | 149 | 199.22 | 125.62 | 272.82 | 136.06 |
| 10 | 150 | 199.22 | 125.62 | 272.82 | 184.08 |
| 11 | 151 | 199.22 | 125.62 | 272.82 | 205.97 |
| 12 | 152 | 199.22 | 125.62 | 272.82 | 258.93 |
| 13 | 153 | 199.22 | 125.62 | 272.82 | 298.96 |
| 14 | 154 | 199.22 | 125.62 | 272.82 | 162.94 |
| 15 | 155 | 199.22 | 125.62 | 272.82 | 135.95 |
| 16 | 156 | 199.22 | 125.62 | 272.82 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $7.1 \%$, that the simple exponential smoothing trend model has satisfied the forecasting for the COD variable.

## B- stochastic forecasting

## B1- auto regression model

Table (41) shows the $\mathrm{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (41) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (41): Forecasted, lower, upper and actual values by AR(1) for COD variable

| $\underline{\text { Row }}$ | Period <br> $(\mathrm{month})$ | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |  | $\frac{\text { Lower }}{\mathrm{mg} / \mathrm{l}}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 151.95 |  | $\frac{\text { Upper }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Actual }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |
| 2 | 142 | 137.93 | 37.08 | 240.66 | 207.14 |
| 3 | 143 | 130.35 | 26.21 | 234.78 | 102.97 |
| 4 | 144 | 126.40 | 21.16 | 231.33 | 206.97 |
| 5 | 145 | 124.02 | 18.67 | 229.38 | 180.98 |
| 6 | 146 | 122.82 | 17.39 | 228.26 | 84.18 |
| 7 | 147 | 122.17 | 16.71 | 227.63 | 164.98 |
| 8 | 148 | 121.82 | 16.35 | 227.29 | 229.93 |
| 9 | 149 | 121.63 | 16.16 | 227.10 | 136.06 |
| 10 | 150 | 121.53 | 16.06 | 227.00 | 184.08 |
| 11 | 151 | 121.47 | 16.00 | 226.94 | 205.97 |
| 12 | 152 | 121.44 | 15.97 | 226.91 | 258.93 |
| 13 | 153 | 121.43 | 15.96 | 226.90 | 298.96 |
| 14 | 154 | 121.42 | 15.95 | 226.89 | 162.94 |
| 15 | 155 | 121.41 | 15.94 | 226.88 | 135.95 |
| 16 | 156 | 121.41 | 15.94 | 226.88 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $47.4 \%$, that the $\mathrm{AR}(1)$ trend model has not satisfied the forecasting for the COD variable.

## B2- moving average regression model

The regression of the additive MA(4) trend model is shown in Figure(43).


Table (42) shows the MA(4) prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (42) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (42): Forecasted, lower, upper and actual values by MA(4) for COD variable

| Row | Period (month) | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Lower }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Upper }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Actual }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 235.22 | 147.50 | 322.94 | 207.14 |
| 2 | 142 | 235.22 | 147.50 | 322.94 | 102.97 |
| 3 | 143 | 235.22 | 147.50 | 322.94 | 176.07 |
| 4 | 144 | 235.22 | 147.50 | 322.94 | 206.99 |
| 5 | 145 | 235.22 | 147.50 | 322.94 | 180.98 |
| 6 | 146 | 235.22 | 147.50 | 322.94 | 84.18 |
| 7 | 147 | 235.22 | 147.50 | 322.94 | 164.98 |
| 8 | 148 | 235.22 | 147.50 | 322.94 | 229.93 |
| 9 | 149 | 235.22 | 147.50 | 322.94 | 136.06 |
| 10 | 150 | 235.22 | 147.50 | 322.94 | 184.08 |
| 11 | 151 | 235.22 | 147.50 | 322.94 | 205.97 |
| 12 | 152 | 235.22 | 147.50 | 322.94 | 258.93 |
| 13 | 153 | 235.22 | 147.50 | 322.94 | 298.9 |


| 14 | 154 | 235.22 | 147.50 | 322.94 | 162.94 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 15 | 155 | 235.22 | 147.50 | 322.94 | 135.95 |
| 16 | 156 | 235.22 | 147.50 | 322.94 | 226.00 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $21.3 \%$, that the $\mathrm{MA}(4)$ trend model did not satisfied the forecasting for the COD variable.

## B3- ARIMA modeling

Table (43) shows the $\operatorname{ARIMA}(1,2,4)$ prediction values for the next $10 \%$ of the predicted and the real data, which equals to 16 observations. In Table (43) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values

Table (43): Forecasted, lower, upper and actual values by ARIMA(1,2,4) for COD variable

| Row | Period <br> $(\mathrm{month})$ | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |  | $\frac{\text { Lower }}{\mathrm{mg} / \mathrm{l}}$ |  | $\frac{\text { Upper }}{\mathrm{mg} / \mathrm{l}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |

Comparing the actual values with the predicted ones, one can conclude, after calcula ting the predication error, which equals to $20.7 \%$, that the $\operatorname{ARIMA}(1,2,4)$ trend model did not satisfy the forecasting for the COD variable.

### 3.5.3.7 results of prediction

The results of error are summarized in the following Table (44), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (44) : Percentage of error of each model for COD variable
Percentage of
Model Mean Error
Linear Method $19.5 \%$
Quadratic Method 3.8 \%
Exponential Growth Method 23.8 \%
Simple Exponential Smoothing $7.1 \%$
Auto Regression, AR(1) $47.4 \%$
Moving Average, MA(4) 21.3 \%
ARIMA (1,2,4) $20.7 \%$

The previous Table (44) shows that the methods, which have satisfied the $10 \%$ acceptable prediction limits, are the quadratic and simple exponential smoothing methods. The best model that gave the least error is the quadratic method.

### 3.5.4 Total phosphorus (T-P) variable:

The consequences that were used to analyze the T-P variable were as follows:

### 3.5.4. detection of missing data and outliers:

From the table (1) it is observed that the data do not contain any missing data, so the second step is to find the outliers, data should be drawn in a scatter diagram (Figure 44) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have approximately two outliers and they are in the following months: March 1992, and December 1999. It was observed that the rainfall in March 1992 was high, and it is known that when the rainfall is high then the T-P will get low, in December 1999 the rainfall was high so the T-P should be low (Appendix (1)). So the real data is on March 1992, the other data was assumed to be an outlier due to human error, and it should be adjusted to a new value since it may greatly influence any statistical calculations and yield biased results. The way that the outlier was adjusted was the same as the missing data treated and it was equal to the average monthly value.


Figure (44):Original Data of T-P mg/l

Figure (45) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are two outliers in both the original data and the residual data in the seasonal condition. Also figure (45) shows the variation in the data for the same month, it can be observe that the variation was the highest on December, and was the lowest on October. Another two outliers were found in the seasonal drawings, they are March 1992, and December 1999, these two outliers were observed in the original data, so no adjustment will be made.

## Seasonal Analysis for T-P mg/l



Figure (45) Outliers of Seasonal Analysis for T-P Variable

After adjustment the outliers, the new adjusted data are plotted in Figure (46), the figure shows that there are still outliers but these outliers cannot be omitted because they are real data so it can influence the statistics results. While comparing the old data (Figure 44) with the new adjusted data (Figure 46) it can be observed that two figures are quite the same and they have the same trend, so the effect of the outliers on the data was so little.


Figure (46): The New Adjusted Data of T-P mg/l

### 3.5.4.2 normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the T-P variable; the calculated values were as follows

| Month | Jan. | Feb. | Mar. | Apr. | May | Jun. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| T-P mg/l | 7.67 | 6.71 | 6.52 | 7.83 | 8.92 | 9.49 |
| Month | Jul. | Aug. | Sep. | Oct. | Nov. | Dec. |
| T-P mg/l | 10.34 | 10.04 | 10.4 | 10.8 | 10.4 | 9.78 |

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from figure (47) that the data of T-P is quite normal and there is a little skewness to the left and bulked to the right, but in general the graph gives an indication that the data is normal.


Figure (47) : Weibull Distribution Model Histogram for T-P Variable.

## B- Coefficient of Variation (COV), Preliminary Test:

The data were divided into four quarters; each quarter consists of 39 data. Table (45) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the T-P variable.

Table (45): The coefficient of variable for T-P

|  | MEAN |  | ST. DE. | C.O.V. |
| :--- | :--- | :---: | :---: | :---: | :---: |
| T-P $\left(1^{\text {st }}\right.$ Quarter $)$ | 7.0 | 5.9 | 2.4 | 0.3 |
| T-P $\left(2^{\text {nd }}\right.$ Quarter $)$ | 7.3 | 7.6 | 2.8 | 0.4 |
| T-P $\left(3^{\text {rd }}\right.$ Quarter $)$ | 10.2 | 7.9 | 2.8 | 0.3 |
| T-P $\left(4^{\text {th }}\right.$ Quarter $)$ | 12.1 | 10.8 | 3.3 | 0.3 |

It can be shown from the table that the value of the coefficient of variation for each quarter is less than 1 , which means that each quarter of the data has a little skewed (either to right or left), so the total data of the T-P variable has less skewness than each
of the four T-P quarters, it can be concluded that the T-P variable does not have skewness.

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K , which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (46) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of $K$, and the Kurtosis coefficient.

Table (46): The Kurtosis Coefficient for T-P

|  | MEAN | VARIANCE | ST. DE. <br> (S ) | K | Kurtosis <br> Coeff. $_{\mathrm{C}_{\mathrm{K}}}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| T-P (1st Quarter ) | 7.0 | 5.9 | 2.4 | 93.3 | -0.3 |
| T-P (2nd Quarter ) | 7.3 | 7.6 | 2.8 | 208.2 | 0.6 |
| T-P (3rd Quarter ) | 10.2 | 7.9 | 2.8 | 174.7 | -0.2 |
| T-P (4th Quarter ) | 12.1 | 10.8 | 3.3 | 555.7 | 1.8 |

From table (46) one can observe that the data in the first, second, and third quarters were normally distributed (mesokurtic), the fourth quarter was fairly leptokurtic. The total data of the T-P variable can be assumed to be normally distributed (mesokurtic).

## D- Shapiro-Wilk test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of $\left(\mathrm{a}_{\mathrm{n}-1+\mathrm{I}}\right)$, was taken for 20 data since the value of $n-1+i$ was equal to 20 , the value of $\left(a_{n-1+1}\right)$ was taken from Appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

From the Tables (47), (48), (49), and (50) it has been shown that the data in the first quarter is nonnormal, the second quarter has a normal data, the third has a quite not normal one, and the fourth one has a normal distribution. It can be assumed that the whole data has a tendency to be normal distribution, the T-P variable is assumed to have a normal distribution.

Table (47): Shapiro-Wilk Test for the Data of T-P's $1^{\text {st }}$ quarter

| No | $\begin{gathered} \text { T-P } \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-P <br> (1) | Inverse order T-P <br> (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 3.00 | 2.20 | 11.80 | 9.60 | 0.3989 | 3.83 |
| 2 | 2.20 | 2.70 | 11.50 | 8.80 | 0.2755 | 2.42 |
| 3 | 2.70 | 3.00 | 10.50 | 7.50 | 0.2380 | 1.78 |
| 4 | 4.50 | 3.10 | 10.40 | 7.30 | 0.2104 | 1.54 |
| 5 | 4.90 | 3.70 | 10.20 | 6.50 | 0.1880 | 1.22 |
| 6 | 5.20 | 4.50 | 10.10 | 5.60 | 0.1689 | 0.95 |
| 7 | 5.90 | 4.90 | 9.90 | 5.00 | 0.1520 | 0.76 |
| 8 | 6.10 | 5.05 | 8.97 | 3.92 | 0.1366 | 0.54 |
| 9 | 7.10 | 5.20 | 8.60 | 3.40 | 0.1225 | 0.42 |
| 10 | 6.50 | 5.67 | 8.50 | 2.83 | 0.1092 | 0.31 |
| 11 | 5.80 | 5.70 | 8.36 | 2.66 | 0.0967 | 0.26 |
| 12 | 6.70 | 5.70 | 8.13 | 2.43 | 0.0848 | 0.21 |
| 13 | 3.10 | 5.80 | 8.10 | 2.30 | 0.0733 | 0.17 |
| 14 | 5.05 | 5.89 | 8.10 | 2.21 | 0.0622 | 0.14 |
| 15 | 5.70 | 5.90 | 8.10 | 2.20 | 0.0515 | 0.11 |
| 16 | 5.67 | 6.10 | 8.10 | 2.00 | 0.0409 | 0.08 |
| 17 | 6.50 | 6.50 | 8.00 | 1.50 | 0.0305 | 0.05 |
| 18 | 5.89 | 6.50 | 7.10 | 0.60 | 0.0203 | 0.01 |
| 19 | 8.10 | 6.54 | 6.70 | 0.16 | 0.0101 | 0.00 |
| 20 | 6.54 | 6.54 | 6.54 | 0.00 |  | $\mathrm{b}=14.78$ |
| 21 | 11.80 | 6.70 | 6.54 | -0.16 |  | $\mathrm{S}=2.43$ |
| 22 | 8.36 | 7.10 | 6.50 | -0.60 |  |  |
| 23 | 8.97 | 8.00 | 6.50 | -1.50 |  |  |
| 24 | 8.60 | 8.10 | 6.10 | -2.00 |  | $\mathrm{W}=0.971>0.939$ |
| 25 | 8.10 | 8.10 | 5.90 | -2.20 |  |  |
| 26 | 8.10 | 8.10 | 5.89 | -2.21 |  | Satisfied |
| 27 | 5.70 | 8.10 | 5.80 | -2.30 |  |  |
| 28 | 6.54 | 8.13 | 5.70 | -2.43 |  |  |
| 29 | 8.10 | 8.36 | 5.70 | -2.66 |  |  |
| 30 | 10.40 | 8.50 | 5.67 | -2.83 |  |  |
| 31 | 8.50 | 8.60 | 5.20 | -3.40 |  |  |
| 32 | 8.00 | 8.97 | 5.05 | -3.92 |  |  |
| 33 | 9.90 | 9.90 | 4.90 | -5.00 |  |  |
| 34 | 10.20 | 10.10 | 4.50 | -5.60 |  |  |
| 35 | 10.10 | 10.20 | 3.70 | -6.50 |  |  |
| 36 | 11.50 | 10.40 | 3.10 | -7.30 |  |  |
| 37 | 10.50 | 10.50 | 3.00 | -7.50 |  |  |
| 38 | 8.13 | 11.50 | 2.70 | -8.80 |  |  |
| 39 | 3.70 | 11.80 | 2.20 | -9.60 |  |  |

Table (48): Shapiro-Wilk Test for the Data of T-P's $2^{\text {nd }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{P} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-P <br> (1) | Inverse order T-P <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 7.45 | 0.70 | 13.96 | 13.26 | 0.3989 | 5.29 |
| 41 | 8.70 | 2.05 | 12.61 | 10.56 | 0.2755 | 2.91 |
| 42 | 5.90 | 3.40 | 11.82 | 8.42 | 0.2380 | 2.00 |
| 43 | 7.78 | 3.70 | 10.91 | 7.21 | 0.2104 | 1.52 |
| 44 | 7.58 | 4.30 | 10.80 | 6.50 | 0.1880 | 1.22 |
| 45 | 8.51 | 4.60 | 10.00 | 5.40 | 0.1689 | 0.91 |
| 46 | 9.25 | 5.04 | 9.96 | 4.92 | 0.1520 | 0.75 |
| 47 | 10.80 | 5.50 | 9.62 | 4.12 | 0.1366 | 0.56 |
| 48 | 4.30 | 5.67 | 9.25 | 3.58 | 0.1225 | 0.44 |
| 49 | 3.40 | 5.90 | 9.24 | 3.34 | 0.1092 | 0.36 |
| 50 | 2.05 | 6.10 | 8.70 | 2.60 | 0.0967 | 0.25 |
| 51 | 0.70 | 6.20 | 8.51 | 2.31 | 0.0848 | 0.20 |
| 52 | 3.70 | 6.20 | 8.10 | 1.90 | 0.0733 | 0.14 |
| 53 | 4.60 | 6.21 | 7.78 | 1.57 | 0.0622 | 0.10 |
| 54 | 5.50 | 6.29 | 7.77 | 1.48 | 0.0515 | 0.08 |
| 55 | 6.40 | 6.30 | 7.58 | 1.28 | 0.0409 | 0.05 |
| 56 | 7.30 | 6.37 | 7.45 | 1.08 | 0.0305 | 0.03 |
| 57 | 6.84 | 6.40 | 7.30 | 0.90 | 0.0203 | 0.02 |
| 58 | 6.37 | 6.81 | 7.16 | 0.35 | 0.0101 | 0.00 |
| 59 | 6.29 | 6.84 | 6.84 | 0.00 |  | $\mathrm{b}=16.83$ |
| 60 | 6.20 | 7.16 | 6.81 | -0.35 |  | $\mathrm{S}=2.75$ |
| 61 | 6.20 | 7.30 | 6.40 | -0.90 |  |  |
| 62 | 5.04 | 7.45 | 6.37 | -1.08 |  |  |
| 63 | 5.67 | 7.58 | 6.30 | -1.28 |  | $\mathrm{W}=0.984>0.939$ |
| 64 | 6.30 | 7.77 | 6.29 | -1.48 |  |  |
| 65 | 7.77 | 7.78 | 6.21 | -1.57 |  | - Satisfied |
| 66 | 9.24 | 8.10 | 6.20 | -1.90 |  |  |
| 67 | 9.62 | 8.51 | 6.20 | -2.31 |  |  |
| 68 | 10.00 | 8.70 | 6.10 | -2.60 |  |  |
| 69 | 10.91 | 9.24 | 5.90 | -3.34 |  |  |
| 70 | 11.82 | 9.25 | 5.67 | -3.58 |  |  |
| 71 | 9.96 | 9.62 | 5.50 | -4.12 |  |  |
| 72 | 8.10 | 9.96 | 5.04 | -4.92 |  |  |
| 73 | 7.16 | 10.00 | 4.60 | -5.40 |  |  |
| 74 | 6.21 | 10.80 | 4.30 | -6.50 |  |  |
| 75 | 6.10 | 10.91 | 3.70 | -7.21 |  |  |
| 76 | 6.81 | 11.82 | 3.40 | -8.42 |  |  |
| 77 | 12.61 | 12.61 | 2.05 | -10.56 |  |  |
| 78 | 13.96 | 13.96 | 0.70 | -13.26 |  |  |

Table (49): Shapiro-Wilk Test for the Data of T-P's $3^{\text {rd }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{P} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-P <br> (1) | Inverse order T-P <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 79 | 13.37 | 3.34 | 15.00 | 11.65 | 0.3989 | 4.65 |
| 80 | 12.99 | 5.06 | 14.71 | 9.64 | 0.2755 | 2.66 |
| 81 | 14.71 | 5.89 | 14.27 | 8.38 | 0.2380 | 1.99 |
| 82 | 12.90 | 5.98 | 13.44 | 7.46 | 0.2104 | 1.57 |
| 83 | 5.89 | 6.26 | 13.41 | 7.16 | 0.1880 | 1.35 |
| 84 | 5.98 | 7.02 | 13.37 | 6.34 | 0.1689 | 1.07 |
| 85 | 8.01 | 7.89 | 12.99 | 5.10 | 0.1520 | 0.78 |
| 86 | 7.02 | 8.01 | 12.90 | 4.89 | 0.1366 | 0.67 |
| 87 | 8.71 | 8.10 | 12.84 | 4.73 | 0.1225 | 0.58 |
| 88 | 8.27 | 8.27 | 12.72 | 4.46 | 0.1092 | 0.49 |
| 89 | 10.66 | 8.57 | 12.38 | 3.81 | 0.0967 | 0.37 |
| 90 | 11.55 | 8.58 | 11.62 | 3.03 | 0.0848 | 0.26 |
| 91 | 13.41 | 8.67 | 11.57 | 2.90 | 0.0733 | 0.21 |
| 92 | 14.27 | 8.71 | 11.56 | 2.85 | 0.0622 | 0.18 |
| 93 | 3.34 | 9.05 | 11.55 | 2.50 | 0.0515 | 0.13 |
| 94 | 12.38 | 9.12 | 11.12 | 2.00 | 0.0409 | 0.08 |
| 95 | 11.62 | 9.42 | 11.07 | 1.65 | 0.0305 | 0.05 |
| 96 | 10.92 | 10.36 | 10.96 | 0.60 | 0.0203 | 0.01 |
| 97 | 5.06 | 10.66 | 10.93 | 0.28 | 0.0101 | 0.00 |
| 98 | 8.58 | 10.92 | 10.92 | 0.00 |  | $\mathrm{b}=17.09$ |
| 99 | 10.36 | 10.93 | 10.66 | -0.28 |  | $\mathrm{S}=2.82$ |
| 100 | 9.42 | 10.96 | 10.36 | -0.60 |  |  |
| 101 | 10.96 | 11.07 | 9.42 | -1.65 |  |  |
| 102 | 11.07 | 11.12 | 9.12 | -2.00 |  | $\mathrm{W}=0.968>0.939$ |
| 103 | 8.10 | 11.55 | 9.05 | -2.50 |  |  |
| 104 | 11.56 | 11.56 | 8.71 | -2.85 |  | F- Satisfied |
| 105 | 13.44 | 11.57 | 8.67 | -2.90 |  |  |
| 106 | 11.12 | 11.62 | 8.58 | -3.03 |  |  |
| 107 | 15.00 | 12.38 | 8.57 | -3.81 |  |  |
| 108 | 10.93 | 12.72 | 8.27 | -4.46 |  |  |
| 109 | 12.72 | 12.84 | 8.10 | -4.73 |  |  |
| 110 | 8.57 | 12.90 | 8.01 | -4.89 |  |  |
| 111 | 8.67 | 12.99 | 7.89 | -5.10 |  |  |
| 112 | 6.26 | 13.37 | 7.02 | -6.34 |  |  |
| 113 | 7.89 | 13.41 | 6.26 | -7.16 |  |  |
| 114 | 9.12 | 13.44 | 5.98 | -7.46 |  |  |
| 115 | 9.05 | 14.27 | 5.89 | -8.38 |  |  |
| 116 | 11.57 | 14.71 | 5.06 | -9.64 |  |  |
| 117 | 12.84 | 15.00 | 3.34 | -11.65 |  |  |

Table (50): Shapiro-Wilk Test for the Data of T-P's $4^{\text {th }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{P} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-P <br> (1) | Inverse order T-P <br> (2) | 2-1 | $a(n-1+i)$ | $(2-1) \times \mathrm{a}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 14.10 | 7.81 | 21.80 | 13.99 | 0.3989 | 5.58 |
| 119 | 10.50 | 8.32 | 20.02 | 11.70 | 0.2755 | 3.22 |
| 120 | 12.32 | 8.47 | 19.99 | 11.53 | 0.2380 | 2.74 |
| 121 | 7.81 | 8.53 | 16.03 | 7.49 | 0.2104 | 1.58 |
| 122 | 9.22 | 8.55 | 15.57 | 7.03 | 0.1880 | 1.32 |
| 123 | 8.32 | 8.84 | 14.64 | 5.80 | 0.1689 | 0.98 |
| 124 | 11.44 | 9.04 | 14.45 | 5.41 | 0.1520 | 0.82 |
| 125 | 8.47 | 9.22 | 14.14 | 4.92 | 0.1366 | 0.67 |
| 126 | 9.86 | 9.33 | 14.10 | 4.77 | 0.1225 | 0.58 |
| 127 | 12.71 | 9.49 | 13.73 | 4.24 | 0.1092 | 0.46 |
| 128 | 10.05 | 9.53 | 13.07 | 3.55 | 0.0967 | 0.34 |
| 129 | 9.33 | 9.86 | 12.76 | 2.91 | 0.0848 | 0.25 |
| 130 | 10.18 | 10.05 | 12.71 | 2.66 | 0.0733 | 0.20 |
| 131 | 9.49 | 10.18 | 12.71 | 2.53 | 0.0622 | 0.16 |
| 132 | 11.78 | 10.50 | 12.62 | 2.11 | 0.0515 | 0.11 |
| 133 | 12.16 | 11.08 | 12.43 | 1.35 | 0.0409 | 0.06 |
| 134 | 8.53 | 11.44 | 12.32 | 0.88 | 0.0305 | 0.03 |
| 135 | 8.84 | 11.46 | 12.22 | 0.76 | 0.0203 | 0.02 |
| 136 | 12.62 | 11.78 | 12.16 | 0.38 | 0.0101 | 0.00 |
| 137 | 12.71 | 12.07 | 12.07 | 0.00 |  | $\mathrm{b}=19.12$ |
| 138 | 14.64 | 12.16 | 11.78 | -0.38 |  | $\mathrm{S}=3.28$ |
| 139 | 11.46 | 12.22 | 11.46 | -0.76 |  |  |
| 140 | 9.04 | 12.32 | 11.44 | -0.88 |  |  |
| 141 | 14.45 | 12.43 | 11.08 | -1.35 |  | $\mathrm{W}=0.891<0.939$ |
| 142 | 13.73 | 12.62 | 10.50 | -2.11 |  |  |
| 143 | 16.03 | 12.71 | 10.18 | -2.53 |  |  |
| 144 | 21.80 | 12.71 | 10.05 | -2.66 |  | I-- Did not Satisfied |
| 145 | 12.43 | 12.76 | 9.86 | -2.91 |  |  |
| 146 | 8.55 | 13.07 | 9.53 | -3.55 |  |  |
| 147 | 9.53 | 13.73 | 9.49 | -4.24 |  |  |
| 148 | 12.76 | 14.10 | 9.33 | -4.77 |  |  |
| 149 | 12.07 | 14.14 | 9.22 | -4.92 |  |  |
| 150 | 11.08 | 14.45 | 9.04 | -5.41 |  |  |
| 151 | 19.99 | 14.64 | 8.84 | -5.80 |  |  |
| 152 | 15.57 | 15.57 | 8.55 | -7.03 |  |  |
| 153 | 12.22 | 16.03 | 8.53 | -7.49 |  |  |
| 154 | 13.07 | 19.99 | 8.47 | -11.53 |  |  |
| 155 | 14.14 | 20.02 | 8.32 | -11.70 |  |  |
| 156 | 20.02 | 21.80 | 7.81 | -13.99 |  |  |

### 3.5.4.3 order of (AR)

For water quality like King Talal Dam, the value of AR, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of T-P does not need more than 1 month till it analyze (Viessman and Lewis, 1996). From Figure (48) it can be seen that the value of AR is about 4 , but the value of $p$ that will be used is 1 for the T-P variable.


Figure (48) Autocorrelation Function for T-P Variable

### 3.5.4.4 order of moving average (MA)

After finding the value of AR , which was 1 , the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (49) shows the change between the real data of the variable T-P and it's moving average with different lengths of p .


Figure (49) Moving Average of T-P with Different Values of (p)

The moving average can be determined from Figure (49) when the difference between the previous length of p and the followed one have a small difference and that occurred when the value of p was 4 ( as shown in Figure (49) ), so the T-P variable has a value of MA(4).

### 3.5.4.5

 order of (I)The last coefficient of ARIMA's parameters is the integrated model ( I ), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (50) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that there are a difference between the original figure and the detrended one but in the seasonal case they are almost the same, which means that the detrended effect could take into consideration

Figure (50): Component Analysis for T-P mg/l


Two season; summer and winter can affect seasonality in Jordan, so if the data has no seasonality effect, then the value of $\mathrm{d}=0$ and if we have seasonality effect then the value of $d=2$. Figures (51), and (52) provide ARIMA model diagnostics for ARIMA $=(1,0,4)$ and $(1,2,4)$. It is seen from the two graphs that the residual in Figure (52) is less than Figure (51) so the coefficients of ARIMA that will be used are $(1,2,4)$


Figure (51): ARIMA (1,0,4) Model Diagnostic for T-P
$\qquad$
ARIMA Model Diagnostics: Data001\$T.P.....mg.I




Figure (52): ARIMA (1,2,4) Model Diagnostic for T-P

### 3.5.4.6

## forecasting future values

The following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last $10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (53).


Figure (53): Trend Analysis for T-P mg/l

It can be observed from the above figure and equation of the linear trend that the data is increasing. Table (51) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (51): The values of the predicted and actual data by linear regression for T-P variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 11.51 | 14.45 |
| 2 | 142 | 11.55 | 13.73 |
| 3 | 143 | 11.59 | 16.03 |
| 4 | 144 | 11.63 | 9.78 |
| 5 | 145 | 11.67 | 12.43 |
| 6 | 146 | 11.72 | 8.55 |
| 7 | 147 | 11.76 | 9.53 |
| 8 | 148 | 11.80 | 12.76 |
| 9 | 149 | 11.84 | 12.07 |
| 10 | 150 | 11.88 | 11.08 |
| 11 | 151 | 11.92 | 19.99 |
| 12 | 152 | 11.97 | 15.57 |
| 13 | 153 | 12.01 | 12.22 |
| 14 | 154 | 12.05 | 13.07 |
| 15 | 155 | 12.09 | 14.14 |
| 16 | 156 | 12.13 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $13.9 \%$, that the linear trend model did not satisfy the forecasting for the T-P variable.

## A2- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (54).


Figure (54):Trend Analysis for T-P mg/l
It can be observed from the above figure and the equation of the quadratic trend that the data is increasing upward. Table (52) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (52): The values of the predicted and actual data by quadratic regression for T-P variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 11.16 | 14.45 |
| 2 | 142 | 11.19 | 13.73 |
| 3 | 143 | 11.21 | 16.03 |
| 4 | 144 | 11.24 | 9.78 |
| 5 | 145 | 11.27 | 12.43 |
| 6 | 146 | 11.29 | 8.55 |
| 7 | 147 | 11.32 | 9.53 |
| 8 | 148 | 11.34 | 12.76 |
| 9 | 149 | 11.37 | 12.07 |
| 10 | 150 | 11.39 | 11.08 |
| 11 | 151 | 11.42 | 19.99 |
| 12 | 152 | 11.44 | 15.57 |
| 13 | 153 | 11.47 | 12.22 |
| 14 | 154 | 11.49 | 13.07 |
| 15 | 155 | 11.52 | 14.14 |
| 16 | 156 | 11.54 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $18.6 \%$, that the quadratic trend model did not satisfy the forecasting for the T-P variable.

## A3- Exponential Growth Regression Model

The regression of the additive exponential growth trend model is shown in Figure (55).


Figure (55): Trend Analysis for T-P mg/l
It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing trend. Table (53) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (53): The values of the predicted and actual data by exponential growth regression for T-P variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 11.72 | 14.45 |
| 2 | 142 | 11.79 | 13.73 |
| 3 | 143 | 11.85 | 16.03 |
| 4 | 144 | 11.92 | 9.78 |


| 5 | 145 | 11.99 | 12.43 |
| :---: | :---: | :---: | :---: |
| 6 | 146 | 12.05 | 8.55 |
| 7 | 147 | 12.12 | 9.53 |
| 8 | 148 | 12.19 | 12.76 |
| 9 | 149 | 12.26 | 12.07 |
| 10 | 150 | 12.33 | 11.08 |
| 11 | 151 | 12.39 | 19.99 |
| 12 | 152 | 12.46 | 15.57 |
| 13 | 153 | 12.53 | 12.22 |
| 14 | 154 | 12.60 | 13.07 |
| 15 | 155 | 12.67 | 14.14 |
| 16 | 156 | 12.75 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $10.1 \%$, that the exponential growth trend model did not satisfy the forecasting for the T-P variable.

## A4- Single Exponential Smoothing Model

The regression of the additive single exponential smoothing trend model is shown in Figure (56).


Figure (56): Single Exponential Smoothing for T-P mg/l
Table (54) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (54) it
can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (54): Forecasted, lower, upper and actual values by single exponential smoothing for T-P variable

| Row | Period (month) | $\frac{\text { Forecast }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | $\frac{\text { Lower }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | Upper mg/l | $\frac{\text { Actual }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 10.30 | 6.11 | 14.49 | 14.45 |
| 2 | 142 | 10.30 | 6.11 | 14.49 | 13.73 |
| 3 | 143 | 10.30 | 6.11 | 14.49 | 16.03 |
| 4 | 144 | 10.30 | 6.11 | 14.49 | 9.78 |
| 5 | 145 | 10.30 | 6.11 | 14.49 | 12.43 |
| 6 | 146 | 10.30 | 6.11 | 14.49 | 8.55 |
| 7 | 147 | 10.30 | 6.11 | 14.49 | 9.53 |
| 8 | 148 | 10.30 | 6.11 | 14.49 | 12.76 |
| 9 | 149 | 10.30 | 6.11 | 14.49 | 12.07 |
| 10 | 150 | 10.30 | 6.11 | 14.49 | 11.08 |
| 11 | 151 | 10.30 | 6.11 | 14.49 | 19.99 |
| 12 | 152 | 10.30 | 6.11 | 14.49 | 15.57 |
| 13 | 153 | 10.30 | 6.11 | 14.49 | 12.22 |
| 14 | 154 | 10.30 | 6.11 | 14.49 | 13.07 |
| 15 | 155 | 10.30 | 6.11 | 14.49 | 14.14 |
| 16 | 156 | 10.30 | 6.11 | 14.49 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $30.7 \%$, that the simple exponential smoothing trend model has not satisfied the forecasting for the T-P variable.

## B- stochastic forecasting

## B1- auto regression model

Table (55) shows the $\mathrm{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (55) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (55): Forecasted, lower, upper and actual values by AR(1) for T-P variable

| Row | $\underline{\text { Period }}$ | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | $\underline{\mathrm{mg} / \mathrm{l}}$ | $\underline{\mathrm{mg} / \mathrm{l}}$ | mg/l | mg/l |
| 1 | 141 | 8.86 | 4.47 | 12.25 | 14.45 |
| 2 | 142 | 8.74 | 3.44 | 14.05 | 13.73 |
| 3 | 143 | 8.66 | 2.99 | 14.34 | 16.03 |
| 4 | 144 | 8.61 | 2.77 | 14.45 | 9.78 |
| 5 | 145 | 8.57 | 2.66 | 14.49 | 12.43 |
| 6 | 146 | 8.55 | 2.60 | 14.49 | 8.55 |
| 7 | 147 | 8.53 | 2.56 | 14.49 | 9.53 |
| 8 | 148 | 8.52 | 2.55 | 14.49 | 12.76 |
| 9 | 149 | 8.51 | 2.53 | 14.48 | 12.07 |
| 10 | 150 | 8.50 | 2.53 | 14.48 | 11.08 |
| 11 | 151 | 8.50 | 2.52 | 14.48 | 19.99 |
| 12 | 152 | 8.50 | 2.52 | 14.47 | 15.57 |
| 13 | 153 | 8.50 | 2.52 | 14.47 | 12.22 |
| 14 | 154 | 8.49 | 2.52 | 14.47 | 13.07 |
| 15 | 155 | 8.49 | 2.52 | 14.47 | 14.14 |
| 16 | 156 | 8.49 | 2.52 | 14.47 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $55.6 \%$, that the $\mathrm{AR}(1)$ trend model has not satisfied the forecasting for the T-P variable.

## B2- Moving Average Regression Model

The regression of the additive MA (4) trend model is shown in Figure (57).


Figure (57): Moving Average for T-P mg/l

Table (56) shows the $\mathrm{MA}(4)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (56) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | mg/l | mg/l | mg/l | $\mathrm{mg} / \mathrm{l}$ |
| 1 | 141 | 12.41 | 7.06 | 17.76 | 14.45 |
| 2 | 142 | 12.41 | 7.06 | 17.76 | 13.73 |
| 3 | 143 | 12.41 | 7.06 | 17.76 | 16.03 |
| 4 | 144 | 12.41 | 7.06 | 17.76 | 9.78 |
| 5 | 145 | 12.41 | 7.06 | 17.76 | 12.43 |
| 6 | 146 | 12.41 | 7.06 | 17.76 | 8.55 |
| 7 | 147 | 12.41 | 7.06 | 17.76 | 9.53 |
| 8 | 148 | 12.41 | 7.06 | 17.76 | 12.76 |
| 9 | 149 | 12.41 | 7.06 | 17.76 | 12.07 |
| 10 | 150 | 12.41 | 7.06 | 17.76 | 11.08 |
| 11 | 151 | 12.41 | 7.06 | 17.76 | 19.99 |
| 12 | 152 | 12.41 | 7.06 | 17.76 | 15.57 |
| 13 | 153 | 12.41 | 7.06 | 17.76 | 12.22 |
| 14 | 154 | 12.41 | 7.06 | 17.76 | 13.07 |
| 15 | 155 | 12.41 | 7.06 | 17.76 | 14.14 |
| 16 | 156 | 12.41 | 7.06 | 17.76 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $8.5 \%$, that the MA(4) trend model has satisfied the forecasting for the T-P variable.

## B3- ARIMA modeling

Table (57) shows the ARIMA(1,2,4) prediction values for the next $10 \%$ of the predicted and the real data, which equals to 16 observations. In Table (57) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (57): Forecasted, lower, upper and actual values by ARIMA(1,2,4) for T-P variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (month) | mg/l | mg/l | mg/l | mg/l |
| 1 | 141 | 9.45 | 4.78 | 14.13 | 14.45 |
| 2 | 142 | 9.85 | 3.98 | 15.71 | 13.73 |
| 3 | 143 | 10.05 | 3.05 | 17.05 | 16.03 |
| 4 | 144 | 9.98 | 2.26 | 17.69 | 9.78 |
| 5 | 145 | 10.06 | 1.51 | 18.60 | 12.43 |
| 6 | 146 | 10.05 | 0.81 | 19.30 | 8.55 |
| 7 | 147 | 10.10 | 0.13 | 20.06 | 9.53 |
| 8 | 148 | 10.38 | 0.00 | 20.75 | 12.76 |
| 9 | 149 | 10.72 | 0.00 | 21.44 | 12.07 |
| 10 | 150 | 11.06 | 0.00 | 22.11 | 11.08 |
| 11 | 151 | 11.39 | 0.00 | 22.77 | 19.99 |
| 12 | 152 | 11.71 | 0.00 | 23.41 | 15.57 |
| 13 | 153 | 12.03 | 0.00 | 24.05 | 12.22 |
| 14 | 154 | 12.34 | 0.00 | 24.68 | 13.07 |
| 15 | 155 | 12.66 | 0.00 | 25.31 | 14.14 |
| 16 | 156 | 12.96 | 0.00 | 25.92 | 20.02 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $23.2 \%$, that the $\operatorname{ARIMA}(1,2,4)$ trend model did not satisfy the forecasting for the T-P variable.

### 3.5.4.7 results of prediction

The results of error are summarized in the following Table (58), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (58) : Percentage of error of each model for T-P variable

| Model | Percentage of <br> Mean Error |
| :--- | :---: |
| Linear Method | $13.9 \%$ |
| Quadratic Method | $18.6 \%$ |
| Exponential Growth Method | $10.1 \%$ |
| Simple Exponential Smoothing | $30.7 \%$ |
| Auto Regression, AR(1) | $55.6 \%$ |
| Moving Average, MA(4) | $8.50 \%$ |
| ARIMA (1,2,4) | $23.2 \%$ |

The previous Table (58) shows that the method, which has satisfied the $10 \%$ acceptable prediction limits, is the Moving Average MA(4) model. The best model that gave the least error is the Moving Average (4) method.

### 3.5.5 Total nitrogen (T-N) variable:

The consequences that were used to analyze the T-N variable were as follows:

### 3.5.5.1 detection of missing data and outliers:

From the table (1) it is observed that the data do not contain any missing data, so the second step is to find the outliers, data should be drawn in a scatter diagram (see Figure (58)) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have a clearly trend after the month July 1993. So the data will be separated into two parts. The first one is before July 1993, and the second one after it. In the first part there is just one low data that could be an outlier and this data is in August 1988. In the second part it can be seen that the data approximately do not have any outlier. In August 1988 the amount of rainfall was somehow high so the data is assumed to be a real data and not an outlier (Appendix (1)). So the real data is on August 1988, and there is not any outlier according to the data.


Figure (58): Original Data of T-N mg/l

Figure (59) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are three outliers in the residual data in the seasonal condition. Also figure (59) shows the variation in the data for the same month, it can be observe that the variation was the highest on January, and was the lowest on October. Another three outliers were found in the seasonal graph, they are July 1995, August 1996, and January 1999, these three outliers were not observed in the original data. In July 1995, and August 1996 the amount of rainfall was low, these two months were treated as real data. In January 1999 the amount of rainfall was high (Appendix (1)), the data was treated as an outlier, the adjustment was made according to the average monthly way.


Figure (59) Outliers of Seasonal Analysis for T-N Variable

After adjustment the outliers, the new adjusted data are plotted in Figure (60), the figure shows that their still outliers but these outliers cannot be omitted because they are real data so it can influence the statistics results. While comparing the old data
(Figure 58) with the new adjusted data (Figure 60) it can be observed that two figures are quite the same and they have the same trend, so the effect of the outliers on the data was so little.


Figure (60): The New Adjusted Data of T-N mg/l

### 3.5.5.2 normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the T-N variable; the calculated values were as follows

Month Jan. Feb. Mar Apr. May Jun. Jul. Aug Sep. Oct. Nov Dec.<br>$\begin{array}{lllllllllllll}\mathrm{T}-\mathrm{N} ~ \mathrm{mg} / \mathrm{l} & 43.2 & 38.8 & 42.6 & 43.0 & 46.0 & 42.3 & 41.1 & 41.2 & 43.1 & 49.0 & 51.8 & 51.7\end{array}$

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from Figure (61) that the data of T-N is quite normal and there is a little skewness to the left and bulked to the right, but in general the graph gives an indication that the data is normal.


Figure (61) : Weibull Distribution Model Histogram

## B- coefficient of variation (COV), preliminary test:

The data were divided into four quarters; each quarter consists of 39 data. Table (59) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the T-N variable.

Table (59): The coefficient of variable for T-N

|  | MEAN | VARIANCE | ST. DE. <br> $(S)$ | C.O.V. |
| :--- | :---: | :---: | :---: | :---: |
| T-N ( $1^{\text {st }}$ Quarter $)$ | 30.9 | 98.2 | 9.9 | 0.3 |
| T-N (2 $2^{\text {nd }}$ Quarter $)$ | 30.9 | 79.3 | 8.9 | 0.3 |
| T-N ( $3^{\text {rd }}$ Quarter $)$ | 54.1 | 121.2 | 11.0 | 0.2 |
| T-N ( $4^{\text {th }}$ Quarter $)$ | 62.7 | 92.2 | 9.6 | 0.2 |

It can be shown from the table that the value of the coefficient of variation for each quarter is less than 1 , which means that each quarter of the data has a little skewed (either to right or left), so the total data of the T-N variable has less skewness than each of the four T-N quarters, it can be concluded that the T-N variable does not have skewness.

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K , which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (60) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of K , and the Kurtosis coefficient.

Table (60): The Kurtosis Coefficient for T-N

|  | MEAN | VARIANCE | ST. DE. <br> (S ) | K | Kurtosis <br> Coeff. <br> $\mathrm{C}_{\mathrm{K}}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| T-N (1st Quarter ) | 30.9 | 98.2 | 9.9 | 28379.0 | -0.1 |
| T-N (2 ${ }^{\text {nd }}$ Quarter ) | 30.9 | 79.3 | 8.9 | 14628.0 | -0.7 |
| T-N (3 $3^{\text {rd }}$ Quarter ) | 54.1 | 121.2 | 11.0 | 41056.8 | -0.2 |
| T-N (4 $4^{\text {th }}$ Quarter $)$ | 62.7 | 92.2 | 9.6 | 20660.6 | -0.6 |

From table (60) one can observe that the data in each quarter normally distributed (mesokurtic). The total data of the T-N variable can be assumed to be normally distributed (mesokurtic).

## D- Shapiro-Wilk test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of $\left(a_{n-1+1}\right)$, was taken for 20 data since the value of $n-1+i$ was equal to 20 , the value of $\left(a_{n-1+1}\right)$ was taken from appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

From the Tables (61), (62), (63), and (64) it has been shown that the data in each quarter was normal. It can be say that the whole data has a tendency to be normal distribution. It can be safely say that TSS variable is normally distributed.

Table (61): Shapiro-Wilk Test for the Data of T-N's $1^{\text {st }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-N <br> (1) | Inverse order <br> T-N <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | (2-1) $\mathrm{x} \mathrm{a}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 52.70 | 11.00 | 52.70 | 41.70 | 0.3989 | 16.63 |
| 2 | 21.50 | 13.90 | 51.94 | 38.04 | 0.2755 | 10.48 |
| 3 | 27.30 | 19.10 | 47.40 | 28.30 | 0.2380 | 6.74 |
| 4 | 20.20 | 19.90 | 45.46 | 25.56 | 0.2104 | 5.38 |
| 5 | 23.70 | 20.20 | 43.30 | 23.10 | 0.1880 | 4.34 |
| 6 | 20.20 | 20.20 | 42.90 | 22.70 | 0.1689 | 3.83 |
| 7 | 13.90 | 20.21 | 40.80 | 20.59 | 0.1520 | 3.13 |
| 8 | 11.00 | 21.50 | 39.40 | 17.90 | 0.1366 | 2.45 |
| 9 | 19.90 | 22.13 | 36.93 | 14.79 | 0.1225 | 1.81 |
| 10 | 27.10 | 22.80 | 36.10 | 13.30 | 0.1092 | 1.45 |
| 11 | 42.90 | 23.21 | 35.75 | 12.55 | 0.0967 | 1.21 |
| 12 | 34.60 | 23.70 | 35.20 | 11.50 | 0.0848 | 0.98 |
| 13 | 23.21 | 25.90 | 35.00 | 9.10 | 0.0733 | 0.67 |
| 14 | 30.02 | 27.10 | 34.60 | 7.50 | 0.0622 | 0.47 |
| 15 | 35.75 | 27.10 | 33.57 | 6.47 | 0.0515 | 0.33 |
| 16 | 36.93 | 27.30 | 32.96 | 5.66 | 0.0409 | 0.23 |
| 17 | 33.57 | 28.24 | 32.50 | 4.26 | 0.0305 | 0.13 |
| 18 | 20.21 | 29.00 | 32.50 | 3.50 | 0.0203 | 0.07 |
| 19 | 27.10 | 30.02 | 31.40 | 1.37 | 0.0101 | 0.01 |
| 20 | 22.13 | 30.70 | 30.70 | 0.00 |  | $\mathrm{b}=60.35$ |
| 21 | 28.24 | 31.40 | 30.02 | -1.37 |  | $\mathrm{S}=9.91$ |
| 22 | 32.96 | 32.50 | 29.00 | -3.50 |  |  |
| 23 | 45.46 | 32.50 | 28.24 | -4.26 |  |  |
| 24 | 51.94 | 32.96 | 27.30 | -5.66 |  | $\mathrm{W}=0.976>0.939$ |
| 25 | 47.40 | 33.57 | 27.10 | -6.47 |  |  |
| 26 | 29.00 | 34.60 | 27.10 | -7.50 |  | = Satisfied |
| 27 | 30.70 | 35.00 | 25.90 | -9.10 |  |  |
| 28 | 35.20 | 35.20 | 23.70 | -11.50 |  |  |
| 29 | 32.50 | 35.75 | 23.21 | -12.55 |  |  |
| 30 | 31.40 | 36.10 | 22.80 | -13.30 |  |  |
| 31 | 25.90 | 36.93 | 22.13 | -14.79 |  |  |
| 32 | 22.80 | 39.40 | 21.50 | -17.90 |  |  |
| 33 | 19.10 | 40.80 | 20.21 | -20.59 |  |  |
| 34 | 32.50 | 42.90 | 20.20 | -22.70 |  |  |
| 35 | 40.80 | 43.30 | 20.20 | -23.10 |  |  |
| 36 | 43.30 | 45.46 | 19.90 | -25.56 |  |  |
| 37 | 39.40 | 47.40 | 19.10 | -28.30 |  |  |
| 38 | 36.10 | 51.94 | 13.90 | -38.04 |  |  |
| 39 | 35.00 | 52.70 | 11.00 | -41.70 |  |  |

Table (62): Shapiro-Wilk Test for the Data of T-N's $2^{\text {nd }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-N <br> (1) | Inverse order T-N (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 29.00 | 17.00 | 49.55 | 32.55 | 0.3989 | 12.98 |
| 41 | 27.70 | 17.30 | 48.20 | 30.91 | 0.2755 | 8.51 |
| 42 | 21.30 | 18.50 | 43.12 | 24.62 | 0.2380 | 5.86 |
| 43 | 19.50 | 18.50 | 41.90 | 23.40 | 0.2104 | 4.92 |
| 44 | 28.00 | 19.00 | 41.00 | 22.00 | 0.1880 | 4.14 |
| 45 | 27.19 | 19.50 | 40.71 | 21.21 | 0.1689 | 3.58 |
| 46 | 31.00 | 21.00 | 40.68 | 19.68 | 0.1520 | 2.99 |
| 47 | 41.00 | 21.25 | 40.62 | 19.37 | 0.1366 | 2.65 |
| 48 | 39.50 | 21.30 | 39.50 | 18.20 | 0.1225 | 2.23 |
| 49 | 17.00 | 25.20 | 38.60 | 13.40 | 0.1092 | 1.46 |
| 50 | 19.00 | 26.20 | 38.29 | 12.09 | 0.0967 | 1.17 |
| 51 | 21.00 | 26.80 | 37.51 | 10.71 | 0.0848 | 0.91 |
| 52 | 28.80 | 27.00 | 36.52 | 9.52 | 0.0733 | 0.70 |
| 53 | 27.00 | 27.19 | 35.76 | 8.57 | 0.0622 | 0.53 |
| 54 | 25.20 | 27.51 | 35.00 | 7.49 | 0.0515 | 0.39 |
| 55 | 21.25 | 27.70 | 33.22 | 5.52 | 0.0409 | 0.23 |
| 56 | 17.30 | 27.70 | 31.00 | 3.30 | 0.0305 | 0.10 |
| 57 | 26.20 | 27.88 | 30.55 | 2.67 | 0.0203 | 0.05 |
| 58 | 35.00 | 28.00 | 29.00 | 1.00 | 0.0101 | 0.01 |
| 59 | 26.80 | 28.80 | 28.80 | 0.00 |  | $\mathrm{b}=53.41$ |
| 60 | 18.50 | 29.00 | 28.00 | -1.00 |  | $\mathrm{S}=8.90$ |
| 61 | 18.50 | 30.55 | 27.88 | -2.67 |  |  |
| 62 | 27.51 | 31.00 | 27.70 | -3.30 |  |  |
| 63 | 27.70 | 33.22 | 27.70 | -5.52 |  | $\mathrm{W}=0.947>0.939$ |
| 64 | 27.88 | 35.00 | 27.51 | -7.49 |  |  |
| 65 | 30.55 | 35.76 | 27.19 | -8.57 |  | Satisfied |
| 66 | 33.22 | 36.52 | 27.00 | -9.52 |  |  |
| 67 | 35.76 | 37.51 | 26.80 | -10.71 |  |  |
| 68 | 38.29 | 38.29 | 26.20 | -12.09 |  |  |
| 69 | 40.71 | 38.60 | 25.20 | -13.40 |  |  |
| 70 | 43.12 | 39.50 | 21.30 | -18.20 |  |  |
| 71 | 41.90 | 40.62 | 21.25 | -19.37 |  |  |
| 72 | 40.68 | 40.68 | 21.00 | -19.68 |  |  |
| 73 | 38.60 | 40.71 | 19.50 | -21.21 |  |  |
| 74 | 36.52 | 41.00 | 19.00 | -22.00 |  |  |
| 75 | 40.62 | 41.90 | 18.50 | -23.40 |  |  |
| 76 | 37.51 | 43.12 | 18.50 | -24.62 |  |  |
| 77 | 49.55 | 48.20 | 17.30 | -30.91 |  |  |
| 78 | 48.20 | 49.55 | 17.00 | -32.55 |  |  |

Table (63): Shapiro-Wilk Test for the Data of T-N's $3{ }^{\text {rd }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | Ordering T-N <br> (1) | Inverse order <br> T-N <br> (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 79 | 48.81 | 36.11 | 79.33 | 43.22 | 0.3989 | 17.24 |
| 80 | 47.46 | 37.09 | 76.76 | 39.67 | 0.2755 | 10.93 |
| 81 | 50.20 | 37.54 | 71.69 | 34.15 | 0.2380 | 8.13 |
| 82 | 53.74 | 39.48 | 70.23 | 30.75 | 0.2104 | 6.47 |
| 83 | 40.27 | 40.27 | 69.23 | 28.96 | 0.1880 | 5.44 |
| 84 | 36.11 | 43.98 | 68.84 | 24.86 | 0.1689 | 4.20 |
| 85 | 52.10 | 44.24 | 66.05 | 21.81 | 0.1520 | 3.32 |
| 86 | 37.09 | 46.40 | 65.54 | 19.15 | 0.1366 | 2.62 |
| 87 | 47.06 | 47.06 | 63.41 | 16.35 | 0.1225 | 2.00 |
| 88 | 47.35 | 47.35 | 63.27 | 15.93 | 0.1092 | 1.74 |
| 89 | 50.24 | 47.46 | 61.38 | 13.91 | 0.0967 | 1.35 |
| 90 | 47.91 | 47.91 | 58.54 | 10.63 | 0.0848 | 0.90 |
| 91 | 69.23 | 48.26 | 57.55 | 9.30 | 0.0733 | 0.68 |
| 92 | 39.48 | 48.81 | 55.98 | 7.17 | 0.0622 | 0.45 |
| 93 | 50.37 | 49.15 | 55.88 | 6.73 | 0.0515 | 0.35 |
| 94 | 58.54 | 49.74 | 55.30 | 5.56 | 0.0409 | 0.23 |
| 95 | 49.76 | 49.76 | 54.52 | 4.76 | 0.0305 | 0.15 |
| 96 | 55.98 | 50.20 | 53.74 | 3.54 | 0.0203 | 0.07 |
| 97 | 46.40 | 50.24 | 52.10 | 1.85 | 0.0101 | 0.02 |
| 98 | 44.24 | 50.37 | 50.37 | 0.00 |  | $\mathrm{b}=66.27$ |
| 99 | 49.74 | 52.10 | 50.24 | -1.85 |  | $\mathrm{S}=11.01$ |
| 100 | 55.30 | 53.74 | 50.20 | -3.54 |  |  |
| 101 | 71.69 | 54.52 | 49.76 | -4.76 |  |  |
| 102 | 65.54 | 55.30 | 49.74 | -5.56 |  | $\mathrm{W}=0.954>0.939$ |
| 103 | 54.52 | 55.88 | 49.15 | -6.73 |  |  |
| 104 | 79.33 | 55.98 | 48.81 | -7.17 |  | I- Satisfied |
| 105 | 68.84 | 57.55 | 48.26 | -9.30 |  |  |
| 106 | 70.23 | 58.54 | 47.91 | -10.63 |  |  |
| 107 | 76.76 | 61.38 | 47.46 | -13.91 |  |  |
| 108 | 63.27 | 63.27 | 47.35 | -15.93 |  |  |
| 109 | 66.05 | 63.41 | 47.06 | -16.35 |  |  |
| 110 | 43.98 | 65.54 | 46.40 | -19.15 |  |  |
| 111 | 55.88 | 66.05 | 44.24 | -21.81 |  |  |
| 112 | 37.54 | 68.84 | 43.98 | -24.86 |  |  |
| 113 | 48.26 | 69.23 | 40.27 | -28.96 |  |  |
| 114 | 63.41 | 70.23 | 39.48 | -30.75 |  |  |
| 115 | 49.15 | 71.69 | 37.54 | -34.15 |  |  |
| 116 | 61.38 | 76.76 | 37.09 | -39.67 |  |  |
| 117 | 57.55 | 79.33 | 36.11 | -43.22 |  |  |

Table (64): Shapiro-Wilk Test for the Data of T-N's $4^{\text {th }}$ quarter

| No | $\begin{gathered} \mathrm{T}-\mathrm{N} \\ \mathrm{mg} / \mathrm{l} \end{gathered}$ | $\begin{aligned} & \text { Ordering } \\ & \text { T-N } \\ & \text { (1) } \end{aligned}$ | Inverse order T-N <br> (2) | 2-1 | $a(n-1+i)$ | $(2-1) \mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 62.05 | 45.94 | 79.85 | 33.90 | 0.3989 | 13.52 |
| 119 | 60.84 | 46.58 | 79.51 | 32.93 | 0.2755 | 9.07 |
| 120 | 52.56 | 47.96 | 79.45 | 31.49 | 0.2380 | 7.49 |
| 121 | 53.29 | 50.32 | 77.60 | 27.28 | 0.2104 | 5.74 |
| 122 | 45.94 | 51.78 | 77.49 | 25.72 | 0.1880 | 4.83 |
| 123 | 51.78 | 52.56 | 76.89 | 24.33 | 0.1689 | 4.11 |
| 124 | 47.96 | 53.03 | 73.67 | 20.64 | 0.1520 | 3.14 |
| 125 | 57.36 | 53.29 | 72.22 | 18.93 | 0.1366 | 2.59 |
| 126 | 53.62 | 53.62 | 69.34 | 15.72 | 0.1225 | 1.93 |
| 127 | 54.28 | 54.28 | 68.63 | 14.35 | 0.1092 | 1.57 |
| 128 | 46.58 | 54.41 | 67.00 | 12.60 | 0.0967 | 1.22 |
| 129 | 53.03 | 57.36 | 66.51 | 9.15 | 0.0848 | 0.78 |
| 130 | 57.55 | 57.55 | 66.41 | 8.86 | 0.0733 | 0.65 |
| 131 | 65.72 | 57.92 | 66.07 | 8.15 | 0.0622 | 0.51 |
| 132 | 76.89 | 59.50 | 65.72 | 6.22 | 0.0515 | 0.32 |
| 133 | 66.41 | 59.85 | 64.68 | 4.83 | 0.0409 | 0.20 |
| 134 | 66.51 | 60.84 | 64.64 | 3.80 | 0.0305 | 0.12 |
| 135 | 68.63 | 61.80 | 63.13 | 1.33 | 0.0203 | 0.03 |
| 136 | 77.60 | 62.05 | 62.58 | 0.53 | 0.0101 | 0.01 |
| 137 | 66.07 | 62.11 | 62.11 | 0.00 |  | $\mathrm{b}=57.81$ |
| 138 | 62.11 | 62.58 | 62.05 | -0.53 |  | $\mathrm{S}=9.60$ |
| 139 | 50.32 | 63.13 | 61.80 | -1.33 |  |  |
| 140 | 61.80 | 64.64 | 60.84 | -3.80 |  |  |
| 141 | 54.41 | 64.68 | 59.85 | -4.83 |  | $\mathrm{W}=0.954>0.939$ |
| 142 | 59.50 | 65.72 | 59.50 | -6.22 |  |  |
| 143 | 72.22 | 66.07 | 57.92 | -8.15 |  |  |
| 144 | 79.45 | 66.41 | 57.55 | -8.86 |  | - $=$ Satisfied |
| 145 | 63.13 | 66.51 | 57.36 | -9.15 |  |  |
| 146 | 67.00 | 67.00 | 54.41 | -12.60 |  |  |
| 147 | 62.58 | 68.63 | 54.28 | -14.35 |  |  |
| 148 | 77.49 | 69.34 | 53.62 | -15.72 |  |  |
| 149 | 79.51 | 72.22 | 53.29 | -18.93 |  |  |
| 150 | 57.92 | 73.67 | 53.03 | -20.64 |  |  |
| 151 | 64.64 | 76.89 | 52.56 | -24.33 |  |  |
| 152 | 59.85 | 77.49 | 51.78 | -25.72 |  |  |
| 153 | 64.68 | 77.60 | 50.32 | -27.28 |  |  |
| 154 | 73.67 | 79.45 | 47.96 | -31.49 |  |  |
| 155 | 69.34 | 79.51 | 46.58 | -32.93 |  |  |
| 156 | 79.85 | 79.85 | 45.94 | -33.90 |  |  |

### 3.5.5.3 order of (AR)

For water quality like King Talal Dam, the value of AR, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of T-N does not need more than 1 month till it analyze (Viessman and Lewis, 1996). From Figure (62) it can be seen that the autocorrelation between data is high and that means that the data has random distributions, but the value of p that will be used is 1 for the $\mathrm{T}-\mathrm{N}$ variable.


Figure (62) Autocorrelation Function for T-N Variable

### 3.5.5.4 <br> order of moving average (MA)

After finding the value of AR, which was 1, the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (63) shows the change between the real data of the variable T-N and it's moving average with different lengths of p .


Figure (63) Moving Average of T-N with Different Values of (p)

The moving average can be determined from Figure (63) when the difference between the previous length of p and the followed one have a small difference and that occurred when the value of p was 5 (as shown in Figure (63)), so the T-N variable has a value of MA (5).

### 3.5.5.5

 order of ( I )The last coefficient of ARIMA's parameters is the integrated model ( I ), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (64) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that there are a difference between the original figure and the detrended one but in the seasonal case they are almost the same, which means that the detrended effect should be taken into consideration


Figure (64): Component Analysis for T-N mg/l

Two season; summer and winter can affect seasonality in Jordan, so if the data has no seasonality effect, then the value of $\mathrm{d}=0$ and if we have seasonality effect then the value of $d=2$. Figures (65), and (66) provide ARIMA model diagnostics for ARIMA $=(1,0,5)$ and $(1,2,5)$. It is seen from the two graphs that the residual in Figure $(65)$ is quite the same as in Figure (66) so the coefficients of ARIMA that will be used are $(1,2,5)$


Figure (65): ARIMA (1,0,5) Model Diagnostics for T-N


Figure (66): ARIMA $(1,2,5)$ Model Diagnostics for T-N


### 3.5.5.6

 forecasting future valuesThe following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last $10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (67).


Figure (67): Trend Analysis for T-N mg/l
It can be observed from the above figure and equation of the linear trend that the data is increasing. Table (65) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (65): The values of the predicted and actual data by linear regression for T-N variable

| Row Period (months) Forecasted $(\mathrm{mg} / \mathrm{l})$  <br> 1 141  Actual $(\mathrm{mg} / \mathrm{l})$ <br> 2 142 63.36 54.41 <br> 3 143 63.67 59.50 <br> 4 144 63.97 72.22 <br> 5 145 64.27 79.45 <br> 6 146 64.58 63.13 <br> 7 147 64.88 67.00 <br> 8 148 65.18 62.58 <br> 9 149 65.49 77.49 <br> 10 150 65.79 79.51 <br> 11 151 66.09 57.92 <br> 12 152 66.40 64.64 <br> 13 153 66.70 59.85 <br> 14 154 67.00 64.68 <br> 15 155 67.31 73.67 <br> 16 156 67.61 69.34 |  | 67.91 | 79.85 |
| :---: | :---: | :---: | :---: |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $3.3 \%$, that the linear trend model has satisfied the forecasting for the T-N variable.

## A2- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (68).


Figure (68): Trend Analysis for T-N mg/l

It can be observed from Figure (68) and the equation of the quadratic trend that the data is increasing upward. Table (66) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (66): The values of the predicted and actual data by quadratic regression for T-N variable

| Row | Period (months) | Forecasted (mg/l) | Actual (mg/l) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 68.95 | 54.41 |
| 2 | 142 | 69.49 | 59.50 |
| 3 | 143 | 70.03 | 72.22 |
| 4 | 144 | 70.58 | 79.45 |
| 5 | 145 | 71.13 | 63.13 |
| 6 | 146 | 71.68 | 67.00 |
| 7 | 147 | 72.24 | 62.58 |
| 8 | 148 | 72.80 | 77.49 |
| 9 | 149 | 73.37 | 79.51 |
| 10 | 150 | 73.93 | 57.92 |
| 11 | 151 | 74.51 | 64.64 |
| 12 | 152 | 75.08 | 59.85 |
| 13 | 153 | 75.66 | 64.68 |
| 14 | 154 | 76.24 | 73.67 |
| 15 | 155 | 76.82 | 69.34 |
| 16 | 156 | 77.41 | 79.85 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $7.2 \%$, that the quadratic trend model has satisfied the forecasting for the T-N variable.

## A3- exponential growth regression model

The regression of the additive exponential growth trend model is shown in Figure (69).


Figure (69): Trend Analysis for T-N mg/l

It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing trend. Table (67) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (67): The values of the predicted and actual data by exponential growth regression for $\mathrm{T}-\mathrm{N}$ variable

| Row | Period (months) | Forecasted $(\mathrm{mg} / \mathrm{l})$ |  | Actual $(\mathrm{mg} / \mathrm{l})$ |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 141 |  | 66.50 | 54.41 |
| 2 | 142 | 67.01 | 59.50 |  |
| 3 | 143 |  | 67.52 | 72.22 |
| 4 | 144 | 68.04 | 79.45 |  |
| 5 | 145 | 68.56 | 63.13 |  |
| 6 | 146 | 69.09 | 67.00 |  |
| 7 | 147 | 69.62 | 62.58 |  |
| 8 | 148 | 70.15 | 77.49 |  |
| 9 | 149 | 70.69 | 79.51 |  |
| 10 | 150 | 71.23 | 57.92 |  |
| 11 | 151 | 71.78 | 64.64 |  |
| 12 | 152 | 72.33 | 59.85 |  |
| 13 | 153 | 72.88 | 64.68 |  |
| 14 | 154 | 73.44 | 73.67 |  |
| 15 | 155 | 74.01 | 69.34 |  |
| 16 | 156 | 74.57 | 79.85 |  |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $3.7 \%$, that the exponential growth trend model has satisfied the forecasting for the T-N variable.

## A4- single exponential smoothing model

The regression of the additive single exponential smoothing trend model is shown in Figure (70).


Figure (70): Single Exponential Smoothing for T-N mg/l
Table (68) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (68) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (68): Forecasted, lower, upper and actual values by single exponential smoothing for T-N variable

| Row | $\frac{\text { Period }}{(\text { month })}$ | Forecast mg/l | $\frac{\text { Lower }}{\underline{\mathrm{mg} / \mathrm{l}}}$ | Upper mg/l | $\frac{\text { Actual }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 59.60 | 43.59 | 75.61 | 54.41 |
| 2 | 142 | 59.60 | 43.59 | 75.61 | 59.50 |
| 3 | 143 | 59.60 | 43.59 | 75.61 | 72.22 |


| 4 | 144 | 59.60 | 43.59 | 75.61 | 79.45 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 5 | 145 | 59.60 | 43.59 | 75.61 | 63.13 |
| 6 | 146 | 59.60 | 43.59 | 75.61 | 67.00 |
| 7 | 147 | 59.60 | 43.59 | 75.61 | 62.58 |
| 8 | 148 | 59.60 | 43.59 | 75.61 | 77.49 |
| 9 | 149 | 59.60 | 43.59 | 75.61 | 79.51 |
| 10 | 150 | 59.60 | 43.59 | 75.61 | 57.92 |
| 11 | 151 | 59.60 | 43.59 | 75.61 | 64.64 |
| 12 | 152 | 59.60 | 43.59 | 75.61 | 59.85 |
| 13 | 153 | 59.60 | 43.59 | 75.61 | 64.68 |
| 14 | 154 | 59.60 | 43.59 | 75.61 | 73.67 |
| 15 | 155 | 59.60 | 43.59 | 75.61 | 69.34 |
| 16 | 156 | 59.60 | 43.59 | 75.61 | 79.85 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $13.8 \%$, that the simple exponential smoothing trend model did not satisfy the forecasting for the T-N variable.

## B- stochastic forecasting

## B1- auto regression model

Table (69) shows the $\mathrm{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (69) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (69): Forecasted, lower, upper and actual values by AR(1) for T-N variable

| Row | Period (month) | Forecast $\mathrm{mg} / \mathrm{l}$ | Lower mg/l | Upper $\mathrm{mg} / \mathrm{l}$ | Actua $\mathrm{mg} / \mathrm{l}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 59.77 | 42.30 | 77.23 | 54.41 |
| 2 | 142 | 58.00 | 34.87 | 81.14 | 59.50 |
| 3 | 143 | 56.47 | 29.84 | 83.10 | 72.22 |
| 4 | 144 | 55.14 | 26.15 | 84.12 | 79.45 |
| 5 | 145 | 53.98 | 23.34 | 84.63 | 63.13 |
| 6 | 146 | 52.98 | 21.14 | 84.82 | 67.00 |
| 7 | 147 | 52.10 | 19.39 | 84.82 | 62.58 |
| 8 | 148 | 51.34 | 17.99 | 84.70 | 77.49 |
| 9 | 149 | 50.68 | 16.85 | 84.52 | 79.51 |


| 10 | 150 | 50.11 | 15.92 | 84.30 | 57.92 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 11 | 151 | 49.62 | 15.16 | 84.07 | 64.64 |
| 12 | 152 | 49.18 | 14.53 | 83.84 | 59.85 |
| 13 | 153 | 48.81 | 14.00 | 83.61 | 64.68 |
| 14 | 154 | 48.48 | 13.56 | 83.40 | 73.67 |
| 15 | 155 | 48.20 | 13.19 | 83.20 | 69.34 |
| 16 | 156 | 47.95 | 12.88 | 83.02 | 79.85 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $30.3 \%$, that the $\operatorname{AR}(1)$ trend model did not satisfy the forecasting for the T-N variable.

## B2- moving average regression model

The regression of the additive MA (5) trend model is shown in Figure (71).


Figure (71): Moving Average for TSS mg/l

Table (70) shows the MA(5) prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (70) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (70): Forecasted, lower, upper and actual values by MA(5) for T-N variable

| Row | Period | Forecast | Lower | Upper | Actual |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Row | (month) | mg/l | $\underline{\mathrm{mg} / 1}$ | mg/l | $\underline{\mathrm{mg} / 1}$ |
| 1 | 141 | 63.58 | 45.61 | 81.55 | 54.41 |
| 2 | 142 | 63.58 | 45.61 | 81.55 | 59.50 |
| 3 | 143 | 63.58 | 45.61 | 81.55 | 72.22 |
| 4 | 144 | 63.58 | 45.61 | 81.55 | 79.45 |
| 5 | 145 | 63.58 | 45.61 | 81.55 | 63.13 |
| 6 | 146 | 63.58 | 45.61 | 81.55 | 67.00 |
| 7 | 147 | 63.58 | 45.61 | 81.55 | 62.58 |
| 8 | 148 | 63.58 | 45.61 | 81.55 | 77.49 |
| 9 | 149 | 63.58 | 45.61 | 81.55 | 79.51 |
| 10 | 150 | 63.58 | 45.61 | 81.55 | 57.92 |
| 11 | 151 | 63.58 | 45.61 | 81.55 | 64.64 |
| 12 | 152 | 63.58 | 45.61 | 81.55 | 59.85 |
| 13 | 153 | 63.58 | 45.61 | 81.55 | 64.68 |
| 14 | 154 | 63.58 | 45.61 | 81.55 | 73.67 |
| 15 | 155 | 63.58 | 45.61 | 81.55 | 69.34 |
| 16 | 156 | 63.58 | 45.61 | 81.55 | 79.85 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $6.7 \%$, that the MA(5) trend model has satisfied the forecasting for the T-N variable.

## B3- ARIMA modeling

Table (71) shows the ARIMA $(1,2,5)$ prediction values for the next $10 \%$ of the predicted and the real data, which equals to 16 observations. In Table (71) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (71): Forecasted, lower, upper and actual values by ARIMA $(1,2,5)$ for T-N variable

| Row | $\left.\frac{\text { Period }}{(m o n t h}\right)$ | $\frac{\text { Forecast }}{}$ |  | $\frac{\text { Lower }}{m g / l}$ |  | $\frac{\text { Upper }}{\underline{\mathrm{mg} / \mathrm{l}}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |


| 7 | 147 | 70.13 | 40.67 | 99.60 | 62.58 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8 | 148 | 70.27 | 39.37 | 101.17 | 77.49 |
| 9 | 149 | 72.10 | 39.70 | 104.50 | 79.51 |
| 10 | 150 | 72.26 | 38.47 | 106.05 | 57.92 |
| 11 | 151 | 74.11 | 38.88 | 109.35 | 64.64 |
| 12 | 152 | 74.29 | 37.71 | 110.88 | 59.85 |
| 13 | 153 | 76.17 | 38.18 | 114.17 | 64.68 |
| 14 | 154 | 76.37 | 37.06 | 115.69 | 73.67 |
| 15 | 155 | 78.28 | 37.58 | 118.98 | 69.34 |
| 16 | 156 | 78.50 | 36.50 | 120.50 | 79.85 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $4.8 \%$, that the $\operatorname{ARIMA}(1,2,5)$ trend model has satisfied the forecasting for the T-N variable.

### 3.5.5.7 results of prediction

The results of error are summarized in the following Table (72), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (72) : Percentage of error of each model for T-N variable

| Model | Percentage of <br> Mean Error |
| :--- | :---: |
| Linear Method | $3.3 \%$ |
| Quadratic Method | $7.2 \%$ |
| Exponential Growth Method | $3.7 \%$ |
| Simple Exponential Smoothing | $13.8 \%$ |
| Auto Regression, AR(1) | $30.3 \%$ |
| Moving Average, MA(5) | $6.7 \%$ |
| ARIMA (1,2,5) | $4.8 \%$ |

The previous Table (72) shows that the methods, which have satisfied the $10 \%$ acceptable prediction limits, are linear, quadratic, exponential growth, moving average MA(5), ARIMA $(1,2,5)$. The best model is ARIMA( $1,2,5$ ) model.

### 3.5.6 Flow rate ( Q ) variable:

The consequences that were used to analyze the Q variable were as follows:

### 3.5.6.1 detection of missing data and outliers:

From table (1) it is observed that the data do not contain any missing data, so the second step is to find the outliers, data should be drawn in a scatter diagram (see Figure (72)) so that outliers will be clearly observed. These data, which contains 156 observations from January 1988 till December 2000, have approximately one outlier and it is in February 1992. It was observed that the rainfall in February 1992 was so high, and it is known that when the rainfall is high (Appendix (1)), the amount of flow rate will be high. So this data is assumed to be a real data and no adjustment will be made on it.


Figure (72): Origin Data of Zarqa River Flow (MCM/month)

Figure (73) shows the outliers for the seasonal trends for the original and the residual data, one can conclude from the charts that there are three outliers in the
residual data in the seasonal condition. Also Figure (73) shows the variation in the data for the same month, it can be observe that the variation was the highest on February, and was the lowest on September. Nine outliers were found in the seasonal graph, they are in February 1988, December 1991, January 1992, February 1992, March 1992, April 1992, May 1992, June 1992, and November 1994. The amount of rainfall on each of the nine months was high (Appendix (1)), the data was treated as a real data, no adjustment was made on the data. The new adjusted data will be the same as the original data.


Figure (73) Outliers of Seasonal Analysis for Q Variable

### 3.5.6.2 normality of data

In this section, normality of data will be checked through four procedures; first one is by drawing a histogram for Weibull's distribution model, second one is through
calculating the coefficient of variance, the third one is through calculating the Kurtosis coefficient, and the fourth one is through calculating the Shapiro-Wilk test. From these four procedures, if the data was not normal then a lognormal transformation to the data will be made.

## A- Weibull's distribution model histogram:

Data will be transformed to the average monthly value for the Q variable; the calculated values were as follows

Month Jan. Feb. Mar. Apr. May Jun. Jul. Aug. Sep. Oct. Nov. Dec.

| Flow | 11.4 | 14.4 | 9.9 | 6.1 | 5.4 | 4.7 | 4.1 | 4.1 | 4.0 | 4.5 | 7.0 | 10.9 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

The Weibull's distribution histogram is drawn for these twelve data. It can be observed from Figure (74) that the data of Q has skewness to the right and bulked to the left, so the data did not satisfy Weibull's Distribution Model for normality. A lognormal test should be made to the data.

Month Jan Feb Mar Apr May Jun Jul. Aug Sep Oct Nov Dec Log (flow) 1.06


Figure (74) : Weibull Distribution Model Histogram


Figure (75) : Weibull Distribution Model Histogram

The Weibull's distribution histogram for the lognormal of the flow is drawn for these twelve data. It can be observed from Figure (75) that the data of $\log (\mathrm{Q})$ has not skewness, so the data is normal.

## B- coefficient of variation (COV), preliminary test:

The data were divided into four quarters; each quarter consists of 39 data.
Table (73) provides the value of the mean, variance, standard deviation, and the coefficient of variation for the $\log (\mathrm{Q})$ variable.

Table (73): The coefficient of variable for $\log (\mathrm{Q})$

|  | MEAN | VARIANCE | ST. DE. | C.O.V $)$ |
| :--- | :--- | :---: | :---: | :---: | :---: |
| $\log (Q) \quad\left(1^{\text {st }}\right.$ Quarter $)$ | 0.702 | 0.051 | 0.226 | 0.32 |
| $\log (\mathrm{Q}) \quad\left(2^{\text {nd }}\right.$ Quarter $)$ | 0.877 | 0.095 | 0.309 | 0.35 |
| $\log (\mathrm{Q}) \quad\left(3^{\text {rd }}\right.$ Quarter $)$ | 0.788 | 0.041 | 0.202 | 0.26 |
| $\log (\mathrm{Q}) \quad\left(4^{\text {th }}\right.$ Quarter $)$ | 0.711 | 0.026 | 0.162 | 0.23 |

It can be shown from the table that the value of the coefficient of variation for all quarters were less than 1 ,it can be concluded that the variable $\log (\mathrm{Q})$ does not have skewness.

## C- Kurtosis coefficient (peakedness), vertical test:

To find the Kurtosis coefficient, one should find the value of K , which depends on the fourth moment about the mean and the number of samples, so that the Kurtosis can be calculated. The Kurtosis will give a good indication if the distribution is leptokurtic or platykutric. The data was divided into four quarters, Table (74) provides the values of the Kurtosis coefficient for each quarter and it provides also the calculations needed to calculate the Kurtosis coefficient, which they are: the mean, the variance or standard error, the value of K , and the Kurtosis coefficient.

Table (74): The Kurtosis Coefficient for Log (Q)

|  | MEAN | VARIANCE | ST. DE. <br> (S ) | K | Kurtosis <br> Coeff. <br> $C_{K}^{\prime}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $\log (\mathrm{Q})(1$ st Quarter) | 0.702 | 0.051 | 0.226 | 0.016 | 3.1 |
| $\log (\mathrm{Q})\left(2^{\text {nd }}\right.$ Quarter) | 0.877 | 0.095 | 0.309 | 0.042 | 1.6 |
| $\log (\mathrm{Q})\left(3^{\text {rd }}\right.$ Quarter $)$ | 0.788 | 0.041 | 0.202 | 0.009 | 2.4 |
| $\log (\mathrm{Q})\left(4^{\text {th }}\right.$ Quarter $)$ | 0.711 | 0.026 | 0.162 | 0.002 | -0.1 |

From table (74) one can observe that the data in quarter one, two, and three are quietly normally distributed, the fourth quarter is normally distributed. The total data of the $\log (\mathrm{Q})$ variable can be assumed to be normally distributed (mesokurtic).

## D- Shapiro-Wilk test

This is another test to show that the data we have is normal or not. Data that have been collected were divided into equal quarters, the value of $\left(a_{n-l+1}\right)$, was taken for 20 data since the value of $n-1+i$ was equal to 20 , the value of $\left(a_{n-1+1}\right)$ was taken from appendix (2). The Shapiro-Wilk value was compared with the five percent critical value for sample size 20 in Appendix (3), if the value of the Shapiro-Wilk test was greater than it then the data will not show evidence of nonnormality.

Table (75): Shapiro-Wilk Test for the Data of $\log$ (Q's) $1^{\text {st }}$ quarter

| No | $\log (\mathrm{Q})$ | Ordering | Inverse order |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MCM/ | Log (Q) | Log (Q) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i}}$ |
|  | MONTH | (1) | (2) |  |  |  |
| 1 | 1.01 | 0.46 | 1.50 | 1.04 | 0.40 | 0.42 |
| 2 | 1.50 | 0.47 | 1.21 | 0.74 | 0.28 | 0.21 |
| 3 | 1.02 | 0.47 | 1.02 | 0.55 | 0.24 | 0.13 |
| 4 | 0.70 | 0.50 | 1.01 | 0.51 | 0.21 | 0.11 |
| 5 | 0.72 | 0.50 | 0.94 | 0.44 | 0.19 | 0.08 |
| 6 | 0.65 | 0.51 | 0.92 | 0.42 | 0.17 | 0.07 |
| 7 | 0.60 | 0.51 | 0.90 | 0.39 | 0.15 | 0.06 |
| 8 | 0.52 | 0.52 | 0.88 | 0.36 | 0.14 | 0.05 |
| 9 | 0.51 | 0.52 | 0.88 | 0.36 | 0.12 | 0.04 |
| 10 | 0.52 | 0.52 | 0.86 | 0.34 | 0.11 | 0.04 |
| 11 | 0.56 | 0.55 | 0.84 | 0.29 | 0.10 | 0.03 |
| 12 | 1.21 | 0.55 | 0.83 | 0.28 | 0.08 | 0.02 |
| 13 | 0.86 | 0.55 | 0.72 | 0.17 | 0.07 | 0.01 |
| 14 | 0.71 | 0.56 | 0.71 | 0.15 | 0.06 | 0.01 |
| 15 | 0.83 | 0.56 | 0.71 | 0.14 | 0.05 | 0.01 |
| 16 | 0.68 | 0.58 | 0.70 | 0.12 | 0.04 | 0.01 |
| 17 | 0.59 | 0.59 | 0.69 | 0.11 | 0.03 | 0.00 |
| 18 | 0.46 | 0.59 | 0.68 | 0.10 | 0.02 | 0.00 |
| 19 | 0.52 | 0.60 | 0.65 | 0.05 | 0.01 | 0.00 |
| 20 | 0.50 | 0.63 | 0.63 | 0.00 |  | $\mathrm{b}=1.29$ |
| 21 | 0.47 | 0.65 | 0.60 | -0.05 |  | $\mathrm{S}=0.23$ |
| 22 | 0.47 | 0.68 | 0.59 | -0.10 |  |  |
| 23 | 0.55 | 0.69 | 0.59 | -0.11 |  |  |
| 24 | 0.69 | 0.70 | 0.58 | -0.12 |  | $\mathrm{W}=0.861<0.939$ |
| 25 | 0.88 | 0.71 | 0.56 | -0.14 |  |  |
| 26 | 0.88 | 0.71 | 0.56 | -0.15 | - | - Did not satisfy |
| 27 | 0.90 | 0.72 | 0.55 | -0.17 |  |  |
| 28 | 0.71 | 0.83 | 0.55 | -0.28 |  |  |
| 29 | 0.59 | 0.84 | 0.55 | -0.29 |  |  |
| 30 | 0.50 | 0.86 | 0.52 | -0.34 |  |  |
| 31 | 0.56 | 0.88 | 0.52 | -0.36 |  |  |
| 32 | 0.55 | 0.88 | 0.52 | -0.36 |  |  |
| 33 | 0.51 | 0.90 | 0.51 | -0.39 |  |  |
| 34 | 0.55 | 0.92 | 0.51 | -0.42 |  |  |
| 35 | 0.58 | 0.94 | 0.50 | -0.44 |  |  |
| 36 | 0.63 | 1.01 | 0.50 | -0.51 |  |  |
| 37 | 0.92 | 1.02 | 0.47 | -0.55 |  |  |
| 38 | 0.84 | 1.21 | 0.47 | -0.74 |  |  |
| 39 | 0.94 | 1.50 | 0.46 | -1.04 |  |  |

Table (76): Shapiro-Wilk Test for the Data of $\log$ (Q's) $2^{\text {nd }}$ quarter

| No | $\begin{gathered} \log (\mathrm{Q}) \\ \text { MCM/ } \\ \text { MONTH } \end{gathered}$ | Ordering $\log (\mathrm{Q})$ <br> (1) | Inverse order Log (Q) (2) | 2-1 | $a(n-1+i)$ | (2-1) $\mathrm{xa}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | 0.61 | 0.45 | 1.84 | 1.38 | 0.40 | 0.55 |
| 41 | 0.59 | 0.46 | 1.53 | 1.07 | 0.28 | 0.30 |
| 42 | 0.51 | 0.46 | 1.47 | 1.00 | 0.24 | 0.24 |
| 43 | 0.46 | 0.51 | 1.39 | 0.88 | 0.21 | 0.19 |
| 44 | 0.46 | 0.51 | 1.18 | 0.67 | 0.19 | 0.13 |
| 45 | 0.45 | 0.59 | 1.17 | 0.58 | 0.17 | 0.10 |
| 46 | 0.51 | 0.59 | 1.17 | 0.57 | 0.15 | 0.09 |
| 47 | 0.59 | 0.60 | 1.05 | 0.44 | 0.14 | 0.06 |
| 48 | 1.53 | 0.61 | 1.01 | 0.40 | 0.12 | 0.05 |
| 49 | 1.39 | 0.66 | 1.00 | 0.34 | 0.11 | 0.04 |
| 50 | 1.84 | 0.69 | 0.99 | 0.30 | 0.10 | 0.03 |
| 51 | 1.47 | 0.70 | 0.94 | 0.24 | 0.08 | 0.02 |
| 52 | 1.18 | 0.71 | 0.92 | 0.21 | 0.07 | 0.02 |
| 53 | 1.01 | 0.72 | 0.92 | 0.21 | 0.06 | 0.01 |
| 54 | 0.94 | 0.73 | 0.91 | 0.19 | 0.05 | 0.01 |
| 55 | 0.84 | 0.79 | 0.91 | 0.12 | 0.04 | 0.00 |
| 56 | 0.82 | 0.82 | 0.91 | 0.10 | 0.03 | 0.00 |
| 57 | 0.79 | 0.82 | 0.91 | 0.09 | 0.02 | 0.00 |
| 58 | 0.82 | 0.84 | 0.88 | 0.04 | 0.01 | 0.00 |
| 59 | 0.91 | 0.84 | 0.84 | 0.00 |  | $\mathrm{b}=1.83$ |
| 60 | 1.17 | 0.88 | 0.84 | -0.04 |  | $\mathrm{S}=0.31$ |
| 61 | 1.17 | 0.91 | 0.82 | -0.09 |  |  |
| 62 | 1.05 | 0.91 | 0.82 | -0.10 |  |  |
| 63 | 1.00 | 0.91 | 0.79 | -0.12 |  | $\mathrm{W}=0.921<0.939$ |
| 64 | 0.91 | 0.91 | 0.73 | -0.19 |  |  |
| 65 | 0.91 | 0.92 | 0.72 | -0.21 |  | =- Did not satisfy |
| 66 | 0.92 | 0.92 | 0.71 | -0.21 |  |  |
| 67 | 0.70 | 0.94 | 0.70 | -0.24 |  |  |
| 68 | 0.71 | 0.99 | 0.69 | -0.30 |  |  |
| 69 | 0.69 | 1.00 | 0.66 | -0.34 |  |  |
| 70 | 0.73 | 1.01 | 0.61 | -0.40 |  |  |
| 71 | 0.88 | 1.05 | 0.60 | -0.44 |  |  |
| 72 | 0.84 | 1.17 | 0.59 | -0.57 |  |  |
| 73 | 0.99 | 1.17 | 0.59 | -0.58 |  |  |
| 74 | 0.92 | 1.18 | 0.51 | -0.67 |  |  |
| 75 | 0.91 | 1.39 | 0.51 | -0.88 |  |  |
| 76 | 0.72 | 1.47 | 0.46 | -1.00 |  |  |
| 77 | 0.66 | 1.53 | 0.46 | -1.07 |  |  |
| 78 | 0.60 | 1.84 | 0.45 | -1.38 |  |  |

Table (77): Shapiro-Wilk Test for the Data of $\log (Q ’ s) 3^{\text {rd }}$ quarter

| No | $\log (\mathrm{Q})$ <br> MCM/ <br> MONTH | Ordering $\log (\mathrm{Q})$ <br> (1) | Inverse order $\log (\mathrm{Q})$ <br> (2) | 2-1 | $a(n-1+i)$ | (2-1) $\times \mathrm{a}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 79 | 0.52 | 0.52 | 1.42 | 0.89 | 0.40 | 0.36 |
| 80 | 0.58 | 0.53 | 1.28 | 0.74 | 0.28 | 0.21 |
| 81 | 0.53 | 0.57 | 1.26 | 0.70 | 0.24 | 0.17 |
| 82 | 0.69 | 0.58 | 1.06 | 0.48 | 0.21 | 0.10 |
| 83 | 1.42 | 0.58 | 1.04 | 0.46 | 0.19 | 0.09 |
| 84 | 1.26 | 0.61 | 0.97 | 0.36 | 0.17 | 0.06 |
| 85 | 0.78 | 0.63 | 0.91 | 0.28 | 0.15 | 0.04 |
| 86 | 0.91 | 0.65 | 0.91 | 0.25 | 0.14 | 0.03 |
| 87 | 0.89 | 0.67 | 0.90 | 0.23 | 0.12 | 0.03 |
| 88 | 0.82 | 0.68 | 0.89 | 0.21 | 0.11 | 0.02 |
| 89 | 0.82 | 0.68 | 0.84 | 0.16 | 0.10 | 0.02 |
| 90 | 0.72 | 0.68 | 0.82 | 0.14 | 0.08 | 0.01 |
| 91 | 0.71 | 0.68 | 0.82 | 0.13 | 0.07 | 0.01 |
| 92 | 0.71 | 0.69 | 0.81 | 0.12 | 0.06 | 0.01 |
| 93 | 0.68 | 0.69 | 0.78 | 0.10 | 0.05 | 0.00 |
| 94 | 0.68 | 0.71 | 0.78 | 0.07 | 0.04 | 0.00 |
| 95 | 0.71 | 0.71 | 0.75 | 0.04 | 0.03 | 0.00 |
| 96 | 0.81 | 0.71 | 0.75 | 0.03 | 0.02 | 0.00 |
| 97 | 1.04 | 0.72 | 0.74 | 0.02 | 0.01 | 0.00 |
| 98 | 0.78 | 0.72 | 0.72 | 0.00 |  | $\mathrm{b}=1.16$ |
| 99 | 0.97 | 0.74 | 0.72 | -0.02 |  | $\mathrm{S}=0.2$ |
| 100 | 0.75 | 0.75 | 0.71 | -0.03 |  |  |
| 101 | 0.75 | 0.75 | 0.71 | -0.04 |  |  |
| 102 | 0.65 | 0.78 | 0.71 | -0.07 |  | $\mathrm{W}=0.866<0.939$ |
| 103 | 0.61 | 0.78 | 0.69 | -0.10 |  |  |
| 104 | 0.58 | 0.81 | 0.69 | -0.12 |  | $\underline{=}$ - Did not satisfy |
| 105 | 0.63 | 0.82 | 0.68 | -0.13 |  |  |
| 106 | 0.57 | 0.82 | 0.68 | -0.14 |  |  |
| 107 | 0.90 | 0.84 | 0.68 | -0.16 |  |  |
| 108 | 0.84 | 0.89 | 0.68 | -0.21 |  |  |
| 109 | 1.28 | 0.90 | 0.67 | -0.23 |  |  |
| 110 | 1.06 | 0.91 | 0.65 | -0.25 |  |  |
| 111 | 0.91 | 0.91 | 0.63 | -0.28 |  |  |
| 112 | 0.72 | 0.97 | 0.61 | -0.36 |  |  |
| 113 | 0.74 | 1.04 | 0.58 | -0.46 |  |  |
| 114 | 0.68 | 1.06 | 0.58 | -0.48 |  |  |
| 115 | 0.69 | 1.26 | 0.57 | -0.70 |  |  |
| 116 | 0.68 | 1.28 | 0.53 | -0.74 |  |  |
| 117 | 0.67 | 1.42 | 0.52 | -0.89 |  |  |

Table (78): Shapiro-Wilk Test for the Data of $\log$ (Q's) $4^{\text {th }}$ quarter

| No | $\log (\mathrm{Q})$ <br> MCM/ <br> MONTH | Ordering Log (Q) <br> (1) | Inverse order Log (Q) <br> (2) | 2-1 | $\mathrm{a}(\mathrm{n}-1+\mathrm{i})$ | $(2-1) \times \mathrm{a}_{(\mathrm{n}-1+\mathrm{i})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 118 | 0.66 | 0.50 | 1.09 | 0.59 | 0.40 | 0.24 |
| 119 | 0.84 | 0.52 | 1.03 | 0.51 | 0.28 | 0.14 |
| 120 | 1.03 | 0.52 | 1.01 | 0.49 | 0.24 | 0.12 |
| 121 | 1.01 | 0.53 | 0.99 | 0.46 | 0.21 | 0.10 |
| 122 | 0.78 | 0.54 | 0.99 | 0.45 | 0.19 | 0.08 |
| 123 | 0.99 | 0.55 | 0.94 | 0.40 | 0.17 | 0.07 |
| 124 | 0.77 | 0.56 | 0.86 | 0.30 | 0.15 | 0.05 |
| 125 | 0.56 | 0.56 | 0.85 | 0.29 | 0.14 | 0.04 |
| 126 | 0.54 | 0.58 | 0.85 | 0.27 | 0.12 | 0.03 |
| 127 | 0.58 | 0.60 | 0.84 | 0.24 | 0.11 | 0.03 |
| 128 | 0.61 | 0.60 | 0.82 | 0.22 | 0.10 | 0.02 |
| 129 | 0.64 | 0.61 | 0.78 | 0.17 | 0.08 | 0.01 |
| 130 | 0.65 | 0.62 | 0.77 | 0.15 | 0.07 | 0.01 |
| 131 | 0.65 | 0.63 | 0.74 | 0.12 | 0.06 | 0.01 |
| 132 | 0.65 | 0.63 | 0.72 | 0.08 | 0.05 | 0.00 |
| 133 | 0.86 | 0.64 | 0.68 | 0.04 | 0.04 | 0.00 |
| 134 | 0.99 | 0.64 | 0.67 | 0.03 | 0.03 | 0.00 |
| 135 | 0.72 | 0.64 | 0.66 | 0.02 | 0.02 | 0.00 |
| 136 | 0.62 | 0.65 | 0.65 | 0.01 | 0.01 | 0.00 |
| 137 | 0.68 | 0.65 | 0.65 | 0.00 |  | $\mathrm{b}=0.95$ |
| 138 | 0.60 | 0.65 | 0.65 | -0.01 |  | $\mathrm{S}=0.16$ |
| 139 | 0.50 | 0.66 | 0.64 | -0.02 |  |  |
| 140 | 0.52 | 0.67 | 0.64 | -0.03 |  |  |
| 141 | 0.55 | 0.68 | 0.64 | -0.04 |  | $\mathrm{W}=0.902<0.939$ |
| 142 | 0.64 | 0.72 | 0.63 | -0.08 |  |  |
| 143 | 0.67 | 0.74 | 0.63 | -0.12 |  |  |
| 144 | 0.74 | 0.77 | 0.62 | -0.15 |  | Did not satisfy |
| 145 | 1.09 | 0.78 | 0.61 | -0.17 |  |  |
| 146 | 0.82 | 0.82 | 0.60 | -0.22 |  |  |
| 147 | 0.85 | 0.84 | 0.60 | -0.24 |  |  |
| 148 | 0.63 | 0.85 | 0.58 | -0.27 |  |  |
| 149 | 0.63 | 0.85 | 0.56 | -0.29 |  |  |
| 150 | 0.60 | 0.86 | 0.56 | -0.30 |  |  |
| 151 | 0.56 | 0.94 | 0.55 | -0.40 |  |  |
| 152 | 0.52 | 0.99 | 0.54 | -0.45 |  |  |
| 153 | 0.53 | 0.99 | 0.53 | -0.46 |  |  |
| 154 | 0.85 | 1.01 | 0.52 | -0.49 |  |  |
| 155 | 0.64 | 1.03 | 0.52 | -0.51 |  |  |
| 156 | 0.94 | 1.09 | 0.50 | -0.59 |  |  |

From the Tables (75), (76), (77), and (78) it has been shown that the data in each quarter did not satisfy Shapiro-Wilk test for normality. So this test did not give any indication about the whole data if it is normal or not.

### 3.5.6.3 order of (AR)

For water quality like King Talal Dam, the value of $A R$, which is expressed by the item (p) shall not be more than 1 since the autocorrelation for a particle of $\log (\mathrm{Q})$ does not need more than 1 month till it analyze (Viessman and Lewis, 1996). From Figure (76) it can be seen that there is a high correlation for the first autocorrelation, the value of p that will be used is 1 for the Q variable.


Figure (76) Autocorrelation Function for Q Variable

### 3.5.6.4 order of moving average (MA)

After finding the value of AR, which was 1 , the following procedure is to determine the value of MA, which is expressed by the item (q). Figure (77) shows the change between the real data of the variable $\log (\mathrm{Q})$ and it's moving average with different lengths of p .


Figure (77) Moving Average of Log (Q) with Different Values of (p)


Cont. Figure (77) Moving Average of $\log (\mathrm{Q})$ with Different Values of (p)

The moving average can be determined from Figure (77) when the difference between the previous length of p and the followed one have a small difference and that occurred when the value of p was 6 ( as shown in Figure (77) ), so the $\log (\mathrm{Q})$ variable has a value of MA(6).

### 3.5.6.5 order of (I)

The last coefficient of ARIMA's parameters is the integrated model ( I ), which expressed by the item (d). The data should be differenced when there is trend or shift or seasonality in the data, otherwise there is no need to make differentiation for the data. Figure (78) consists of four graphs, which provides a good idea if there is a difference between the original, detrended, seasonally adjusted, and seasonally adjusted and detrended data. It is shown from these four graphs that there are a difference between the original figure and the seasonally one but in the detrended case they are almost the same, which means that the seasonally effect should be taken into consideration.


Figure (78): Component Analysis for $\log$ (Q) MCM/month

Two season; summer and winter can affect seasonality in Jordan, so if the data has no seasonality effect, then the value of $\mathrm{d}=0$ and if we have seasonality effect then the value of $\mathrm{d}=2$. Figures (79), and (80) provide ARIMA model diagnostics for ARIMA $=(1,0,6)$ and $(1,2,6)$. It is seen from the two graphs that the residual in Figure (79) is quite the same as in Figure (80) so the coefficients of ARIMA that will be used are (1,2,6).


Figure (79): ARIMA (1,0,6) Model Diagnostic for $\log (\mathrm{Q})$
$\qquad$
ARIMA Model Diagnostics: Data001\$Log..Q..MCM.month


ACF Plot of Residuals




Figure (80): ARIMA (1,2,6) Model Diagnostic for $\log (\mathrm{Q})$

### 3.5.6.6 forecasting future values

The following procedure will be used in the forecasting: The values of the data collected will be divided into two parts, the first part consists of $90 \%$ of the real data, and this data will be analyzed and predicted. And the second part consists of the last
$10 \%$ of the real data, and this part will be compared with the predicted values in the mean. The best model is the one that gives the least error in mean.

## A- deterministic forecasting

## A1- linear regression model

The regression of the additive linear trend is shown in Figure (81).


Figure (81):Trend Analysis for $\log (Q)$ MCM/month

It can be observed from the above figure and equation of the linear trend that the data is decreasing very slowly. Table (79) shows the linear prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (79): The values of the predicted and actual data by linear regression for $\log (\mathrm{Q})$ variable

| Row | Period (months) | Forecasted (MCM) | Actual (MCM) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 0.7753 | 0.55 |
| 2 | 142 | 0.7752 | 0.64 |
| 3 | 143 | 0.7752 | 0.67 |
| 4 | 144 | 0.7752 | 0.74 |
| 5 | 145 | 0.7752 | 1.09 |
| 6 | 146 | 0.7752 | 0.82 |
| 7 | 147 | 0.7751 | 0.85 |


| 8 | 148 | 0.7751 | 0.63 |
| :---: | :---: | :---: | :---: |
| 9 | 149 | 0.7751 | 0.63 |
| 10 | 150 | 0.7751 | 0.60 |
| 11 | 151 | 0.7750 | 0.56 |
| 12 | 152 | 0.7750 | 0.52 |
| 13 | 153 | 0.7750 | 0.53 |
| 14 | 154 | 0.7750 | 0.85 |
| 15 | 155 | 0.7749 | 0.64 |
| 16 | 156 | 0.7749 | 0.94 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $9.2 \%$, that the linear trend model has satisfied the forecasting for $\log (\mathrm{Q})$ variable.

## A2- quadratic regression model

The regression of the additive quadratic trend is shown in Figure (82).


Figure (82): Trend Analysis for log (Q) MCM/month
It can be observed from Figure (82) and the equation of the quadratic trend that the data is increasing upward and then it is decreasing. Table (80) shows the quadratic prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (80): The values of the predicted and actual data by quadratic regression for $\log$ (Q) variable

| Row | Period (months) | Forecasted (MCM) | Actual (MCM) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 0.6523 | 0.55 |
| 2 | 142 | 0.6471 | 0.64 |
| 3 | 143 | 0.6418 | 0.67 |
| 4 | 144 | 0.6364 | 0.74 |
| 5 | 145 | 0.6309 | 1.09 |
| 6 | 146 | 0.6253 | 0.82 |
| 7 | 147 | 0.6197 | 0.85 |
| 8 | 148 | 0.6140 | 0.63 |
| 9 | 149 | 0.6082 | 0.63 |
| 10 | 150 | 0.6024 | 0.60 |
| 11 | 151 | 0.5965 | 0.56 |
| 12 | 152 | 0.5905 | 0.52 |
| 13 | 153 | 0.5844 | 0.53 |
| 14 | 154 | 0.5783 | 0.85 |
| 15 | 155 | 0.5721 | 0.64 |
| 16 | 156 | 0.5658 | 0.94 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $15.3 \%$, that the quadratic trend model did not satisfy the forecasting for $\log (\mathrm{Q})$ variable.

## A3- exponential growth regression model

The regression of the additive exponential growth trend model is shown in Figure (83).


Figure (83): Trend Analysis for $\log (Q) M C M /$ month

It can be observed from the above figure and equation of the exponential growth trend that the data has an increasing (very slowly) trend. Table (81) shows the exponential growth prediction of the next $10 \%$ of the predicted and real data, which equals to 16 observations.

Table (81): The values of the predicted and actual data by exponential growth regression for $\log (\mathrm{Q})$ variable

| Row | Period (months) | Forecasted (MCM) | Actual (MCM) |
| :---: | :---: | :---: | :---: |
| 1 | 141 | 0.7620 | 0.55 |
| 2 | 142 | 0.7623 | 0.64 |
| 3 | 143 | 0.7625 | 0.67 |
| 4 | 144 | 0.7628 | 0.74 |
| 5 | 145 | 0.7630 | 1.09 |
| 6 | 146 | 0.7633 | 0.82 |
| 7 | 147 | 0.7636 | 0.85 |
| 8 | 148 | 0.7638 | 0.63 |
| 9 | 149 | 0.7641 | 0.63 |
| 10 | 150 | 0.7643 | 0.60 |
| 11 | 151 | 0.7646 | 0.56 |
| 12 | 152 | 0.7649 | 0.52 |
| 13 | 153 | 0.7651 | 0.53 |
| 14 | 154 | 0.7654 | 0.85 |
| 15 | 155 | 0.7656 | 0.64 |
| 16 | 156 | 0.7659 | 0.94 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $7.9 \%$, that the exponential growth trend model has satisfied the forecasting for the $\log (\mathrm{Q})$ variable.

## A4- single exponential smoothing model

The regression of the additive single exponential smoothing trend model is shown in Figure (84).


Figure (84): Single Exponential Smoothing for $\log (Q)$ MCM/month
Table (82) shows the single exponential smoothing prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (82) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (82): Forecasted, lower, upper and actual values by single exponential smoothing for $\log (\mathrm{Q})$ variable

| $\underline{\text { Row }}$ | Period <br> $\underline{(m o n t h})$ | $\frac{\text { Forecast }}{\underline{(M C M)}}$ | $\underline{\text { Lower }}$ <br> $1(\mathrm{MCM})$ | $\underline{(\mathrm{UPp})}$ | $\underline{\underline{\text { Actual }}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 0.5191 | 0.2226 | 0.8156 | 0.55 |
| 2 | 142 | 0.5191 | 0.2226 | 0.8156 | 0.64 |
| 3 | 143 | 0.5191 | 0.2226 | 0.8156 | 0.67 |


| 4 | 144 | 0.5191 | 0.2226 | 0.8156 | 0.74 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 5 | 145 | 0.5191 | 0.2226 | 0.8156 | 1.09 |
| 6 | 146 | 0.5191 | 0.2226 | 0.8156 | 0.82 |
| 7 | 147 | 0.5191 | 0.2226 | 0.8156 | 0.85 |
| 8 | 148 | 0.5191 | 0.2226 | 0.8156 | 0.63 |
| 9 | 149 | 0.5191 | 0.2226 | 0.8156 | 0.63 |
| 10 | 150 | 0.5191 | 0.2226 | 0.8156 | 0.60 |
| 11 | 151 | 0.5191 | 0.2226 | 0.8156 | 0.56 |
| 12 | 152 | 0.5191 | 0.2226 | 0.8156 | 0.52 |
| 13 | 153 | 0.5191 | 0.2226 | 0.8156 | 0.53 |
| 14 | 154 | 0.5191 | 0.2226 | 0.8156 | 0.85 |
| 15 | 155 | 0.5191 | 0.2226 | 0.8156 | 0.64 |
| 16 | 156 | 0.5191 | 0.2226 | 0.8156 | 0.94 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $35.6 \%$, that the simple exponential smoothing trend model did not satisfy the forecasting for the Q variable.

## B- stochastic forecasting

## B1- auto regression model

Table (83) shows the $\mathrm{AR}(1)$ prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (83) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (83): Forecasted, lower, upper and actual values by $\mathrm{AR}(1)$ for Q variable

| Row | Period (month) | Forecast (MCM) | Lower $(\mathrm{MCM})$ | Upper <br> (MCM) | Actual <br> (MCM) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 141 | 0.5975 | 0.2487 | 0.9463 | 0.55 |
| 2 | 142 | 0.6516 | 0.2262 | 1.0770 | 0.64 |
| 3 | 143 | 0.6894 | 0.2313 | 1.1474 | 0.67 |
| 4 | 144 | 0.7157 | 0.2425 | 1.1889 | 0.74 |
| 5 | 145 | 0.7341 | 0.2537 | 1.2145 | 1.09 |
| 6 | 146 | 0.7470 | 0.2631 | 1.2308 | 0.82 |
| 7 | 147 | 0.7559 | 0.2704 | 1.2414 | 0.85 |


| 8 | 148 | 0.7622 | 0.2758 | 1.2485 | 0.63 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 9 | 149 | 0.7665 | 0.2798 | 1.2533 | 0.63 |
| 10 | 150 | 0.7696 | 0.2827 | 1.2565 | 0.60 |
| 11 | 151 | 0.7717 | 0.2847 | 1.2587 | 0.56 |
| 12 | 152 | 0.7732 | 0.2861 | 1.2602 | 0.52 |
| 13 | 153 | 0.7742 | 0.2872 | 1.2613 | 0.53 |
| 14 | 154 | 0.7750 | 0.2879 | 1.2620 | 0.85 |
| 15 | 155 | 0.7755 | 0.2884 | 1.2626 | 0.64 |
| 16 | 156 | 0.7758 | 0.2887 | 1.2629 | 0.94 |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $4.9 \%$, that the $\operatorname{AR}(1)$ trend model has satisfied the forecasting for the Q variable.

## B2- moving average regression model

The regression of the additive MA (6) trend model is shown in Figure (85).


Figure(85): Moving Average for $\log$ (Q) $M C M /$ month
Table (84) shows the MA(6) prediction for the next $10 \%$ of the predicted and real data, which equals to 16 observations. In Table (84) it can be seen the upper and lower values, the forecasted value was the average between the upper and lower values.

Table (84): Forecasted, lower, upper and actual values by MA(6) for $\log (Q)$ variable

| $\frac{\text { Row }}{}$ | $\frac{\text { Period }}{(\text { month })}$ | $\frac{\frac{\text { Forecast }}{(M C M)}}{}$ |  | $\frac{\text { Lower }}{(\mathrm{MCM})}$ | $\frac{\text { Upper }}{(\mathrm{MCM})}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |

Comparing the actual values with the predicted ones, one can conclude, after calculating the predication error, which equals to $9.0 \%$, that the MA(6) trend model has satisfied the forecasting for the Q variable.

## B3- ARIMA modeling

The software (Minitab 13) cannot estimate the value of ARIMA (1,2,6), because each of ARIMA parameter should not exceed 5 .

### 3.5.6.7 results of prediction

The results of error are summarized in the following Table (85), which provides a summary of the models name used in the prediction and also it provides the percentage error.

Table (85) : Percentage of error of each model for $\log (\mathrm{Q})$ variable

| Model | Percentage of <br> Mean Error |
| :--- | :---: |
| Linear Method | $9.2 \%$ |
| Quadratic Method | $15.3 \%$ |
| Exponential Growth Method | $7.9 \%$ |
| Simple Exponential Smoothing | $35.6 \%$ |
| Auto Regression, AR(1) | $4.9 \%$ |
| Moving Average, MA(6) | $9.0 \%$ |
| ARIMA $(1,2,6)$ | -- |

The previous Table (85) shows that the methods, which have satisfied the $10 \%$ acceptable prediction limits, are linear, exponential growth, Auto regression $\operatorname{AR}(1)$, and moving average $\mathrm{MA}(5)$. The best model is $\mathrm{AR}(1)$ model with a least error of $4.9 \%$.

### 3.6 Cross and Distance Correlation

In this section, the relations between Zarka's River variables, which is expressed in term of cross variable and then the relation between each variable in Zarka River and Samra's effluent, the figures of each relation is shown in Appendix (4).

### 3.6.1 Cross correlation in Zarka River variables:

The cross correlation procedure will be implemented between all the six variables; Flow, TSS, $\mathrm{BOD}_{5}, \mathrm{COD}, \mathrm{T}-\mathrm{P}$, and T-N, in Zarka River.
3.6.1.1 cross correlation between Zarka River flow in MCM/month and TSS in mg/l

The relation between flow and TSS has a cyclic shape. From the figure in Appendix4 (A4.1) it can be seen that the relation between the flow and TSS in Zarka River varies every about six months.
3.6.1.2 cross correlation between Zarka River flow in MCM/month and $\mathrm{BOD}_{5}$ in $\mathrm{mg} / \mathrm{l}$

The relation between flow and $\mathrm{BOD}_{5}$ has a cyclic shape. From the figure in Appendix4 (A4.2) it can be seen that the relation between the flow and $\mathrm{BOD}_{5}$ in Zarka River is negative in most of the times. But there are some months where the relation between them is positive.

### 3.6.1.3 cross correlation between Zarka River flow in MCM/month and COD in mg/l

The relation between flow and COD has a cyclic shape. From the figure in Appendix4 (A4.3) it can be seen that the relation between the flow and COD in Zarka River is negative in most of the times. But there are some months where the relation between them is positive.

### 3.6.1.4 cross correlation between Zarka River flow in MCM/month and T-P in $\mathbf{m g} / \mathrm{l}$

The relation between flow and T-P has a cyclic shape. From the figure in Appendix4 (A4.4) it can be seen that the relation between the flow and T-P in Zarka River varies every about six months. Also it can be seen that the relation is negative in most of the times.
3.6.1.5 cross correlation between Zarka River flow in MCM/month and T-N in mg/l

The relation between flow and T-N has a cyclic shape. From the figure in Appendix4 (A4.5) it can be seen that the relation between the flow and T-N in Zarka River varies every about six months. Also it can be seen that the relation is negative in most of the times.

### 3.6.1.6 cross correlation between Zarka River TSS in mg/l and $\mathrm{BOD}_{5}$ in mg/l

The relation between TSS and $\mathrm{BOD}_{5}$ has a cyclic shape. From the figure in Appendix4 (A4.6) it can be seen that the relation between the TSS and $\mathrm{BOD}_{5}$ in Zarka

River varies every about six months. Also it can be seen that the relation is positive in most of the times.

### 3.6.1.7 cross correlation between Zarka River TSS in mg/l and COD in mg/l

The relation between TSS and COD has a cyclic shape. From the figure in Appendix4 (A4.7) it can be seen that the relation between the TSS and COD in Zarka River varies every about six months. Also it can be seen that the relation is positive in most of the times.

### 3.6.1.8 cross correlation between Zarka River TSS in mg/l and T-P in mg/l

The relation between TSS and T-P has a cyclic shape. From the figure in Appendix4 (A4.9) it can be seen that the relation between the TSS and T-P in Zarka River varies every about six months. Also it can be seen that the correlation in the positive direction is more than in the negative direction.

### 3.6.1.9 cross correlation between Zarka River TSS in mg/l and T-N in mg/l

It can be seen from the figure in Appendix4 (A4.9) that when TSS increases the amount of T-N increases. So the relation between the TSS and T-N is increasing in most of the times.
3.6.1.10 cross correlation between Zarka River $\mathrm{BOD}_{5}$ in $\mathrm{mg} / \mathrm{l}$ and COD in mg/l

It can be seen from the figure in Appendix4 (A4.10) that when $\mathrm{BOD}_{5}$ increases the amount of COD increases, and that the correlation between them is high. So the relation between the $\mathrm{BOD}_{5}$ and COD is an increasing regression.
3.6.1.11 cross correlation between Zarka's River $\mathrm{BOD}_{5}$ in $\mathrm{mg} / \mathrm{l}$ and T-P in mg/l

It can be seen from the figure in Appendix4 (A4.11) that when $\mathrm{BOD}_{5}$ increases the amount of T-P increases, and that the correlation between them is high. So the relation between the $\mathrm{BOD}_{5}$ and $\mathrm{T}-\mathrm{P}$ is an increasing regression.
> 3.6.1.12 cross correlation between Zarka River $\mathrm{BOD}_{5}$ in mg/l and T-N in mg/l

It can be seen from the figure in Appendix4 (A4.12) that when $\mathrm{BOD}_{5}$ increases the amount of T-N increases, and that the correlation between them is high. So the relation between the $\mathrm{BOD}_{5}$ and $\mathrm{T}-\mathrm{N}$ is an increasing regression.

### 3.6.1.13 cross correlation between Zarka River COD in mg/l and T-P in mg/l

It can be seen from the figure in Appendix4 (A4.13) that when COD increases the amount of T-P increases, and that the correlation between them is high. So the relation between the COD and T-P is an increasing regression.
3.6.1.14 cross correlation between Zarka River COD in mg/l and T-N in mg/l

It can be seen from the figure in Appendix4 (A4.14) that when COD increases the amount of T-N increases, and that the correlation between them is high. So the relation between the COD and T-N is an increasing regression.

Figure (86): Example on the Cross Correlation Function: BOD5 mg/l; COD mg/l
CCF - correlates BOD5 $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and COD $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$

|  |  |  |
| :---: | :---: | :---: |
| -22 | 0.086 | XXX |
| -21 | 0.046 | XX |
| -20 | 0.044 | XX |
| -19 | 0.060 | XXX |
| -18 | 0.099 | XXX |
| -17 | 0.053 | XX |
| -16 | 0.047 | XX |
| -15 | 0.041 | XX |
| -14 | 0.028 | XX |
| -13 | 0.117 | XXXX |
| -12 | 0.177 | XXXXX |
| -11 | 0.233 | XXXXXXX |
| -10 | 0.226 | XXXXXXX |
| -9 | 0.194 | XXXXXX |
| -8 | 0.141 | XXXXX |
| -7 | 0.153 | XXXXX |
| -6 | 0.209 | XXXXXX |
| -5 | 0.172 | XXXXX |
| -4 | 0.085 | XXX |
| -3 | 0.171 | XXXXX |
| -2 | 0.166 | XXXXX |
| -1 | 0.251 | XXXXXXX |
| 0 | 0.453 | XXXXXXXXXXXX |
| 1 | 0.403 | XXXXXXXXXXX |
| 2 | 0.351 | XXXXXXXXXX |
| 3 | 0.271 | XXXXXXXX |
| 4 | 0.285 | XXXXXXXX |
| 5 | 0.279 | XXXXXXXX |
| 6 | 0.264 | XXXXXXXX |
| 7 | 0.296 | XXXXXXXX |
| 8 | 0.262 | XXXXXXXX |
| 9 | 0.146 | XXXXX |
| 10 | 0.226 | XXXXXXX |
| 11 | 0.203 | XXXXXX |
| 12 | 0.319 | XXXXXXXXX |
| 13 | 0.320 | XXXXXXXXX |
| 14 | 0.210 | XXXXXX |
| 15 | 0.262 | XXXXXXXX |
| 16 | 0.236 | XXXXXXX |
| 17 | 0.162 | XXXXX |
| 18 | 0.265 | XXXXXXXX |
| 19 | 0.203 | XXXXXX |
| 20 | 0.180 | XXXXXX |
| 21 | 0.194 | XXXXXX |
| 22 | 0.178 | XXXXX |

3.6.1.15 cross correlation between Zarka River T-P in mg/l and T-N in mg/l

It can be seen from the figure in Appendix4 (A4.15) that when T-P increases the amount of T-N increases, and that the correlation between them is so high. So the relation between the T-P and T-N is an increasing regression.

### 3.6.2 Distance cross correlation between Zarka River and Samra's effluent

The distance cross correlation procedure will be implemented between each variable; Flow, TSS, $\mathrm{BOD}_{5}$, COD, T-P, and T-N, in Zarka River and its arbitrary in Samra's effluent.

### 3.6.2.1 distance cross correlation function: Zarka Flow MCM/month; Samra flow MCM/month

It can be seen from the figure in Appendix4 (A4.16) that the relation between the Zarka River flow and Samra's effluent is decreasing regression.

### 3.6.2.2 distance cross correlation function: Zarka TSS mg/l; TSS Samra mg/l

The relation between TSS in Zarka River and TSS in Samra's effluent has a cyclic shape. From the figure in Appendix4 (A4.17) it can be seen that the relation between the TSS in Zarka and Samra varies every about six months. Also it can be seen that the relation is negative in most of the times.

### 3.6.2.3 distance cross correlation function: Zarka $\mathrm{BOD}_{5} \mathrm{mg} /$; $\mathrm{BOD}_{5}$ Samra mg/l

The relation between $\mathrm{BOD}_{5}$ in Zarka River and Samra's effluent is positively in most of the times, see Appendix4 (A4.18).

### 3.6.2.4 distance cross correlation function: Zarka COD mg/l; COD Samra mg/l

The relation between COD in Zarka River and Samra's effluent is positively in most of the times see Appendix4 (A4.19).

### 3.6.2.5 distance cross correlation function: Zarka T-P mg/l; T-P Samra mg/l

The relation between T-P in Zarka River and Samra's effluent is positively in most of the times, but the correlation is not very high between these two variables see Appendix4 (A4.20).

### 3.6.2.6 distance cross correlation function: Zarka T-N mg/l; T-N Samra mg/l

The relation between T-N in Zarka River and Samra's effluent is positively in most of the times, but the correlation is not very high between these two variables see Appendix 4 (A4.21).

Figure (87): Example on the Distance Correlation Function: BOD5 in Zarka River mg/l; BOD5 Samra mg/l

| CCF - correlates BOD5 mg/l(t) and BOD5 Samra mg/l $1(\mathrm{t}+\mathrm{k}$ ) |  |  |
| :---: | :---: | :---: |
|  |  |  |
| -22 | -0.000 | X |
| -21 | 0.039 | XX |
| -20 | 0.101 | XXXX |
| -19 | 0.096 | XXX |
| -18 | 0.179 | XXXXX |
| -17 | 0.152 | XXXXX |
| -16 | 0.193 | XXXXXX |
| -15 | 0.173 | XXXXX |
| -14 | 0.093 | XXX |
| -13 | 0.028 | XX |
| -12 | -0.016 | X |
| -11 | -0.043 | XX |
| -10 | 0.008 | X |
| -9 | 0.037 | XX |
| -8 | 0.047 | XX |
| -7 | 0.215 | XXXXXX |
| -6 | 0.214 | XXXXXX |
| -5 | 0.273 | XXXXXXXX |
| -4 | 0.282 | XXXXXXXX |
| -3 | 0.211 | XXXXXX |
| -2 | 0.192 | XXXXXX |
| -1 | 0.110 | XXXX |
| 0 | 0.188 | Xxxxxx |
| 1 | 0.019 | X |
| 2 | 0.053 | XX |
| 3 | 0.100 | XXX |
| 4 | 0.138 | XXXX |
| 5 | 0.145 | XXXXX |
| 6 | 0.216 | XXXXXX |
| 7 | 0.267 | XXXXXXXX |
| 8 | 0.245 | XXXXXXX |
| 9 | 0.211 | XXXXXX |
| 10 | 0.084 | XXX |
| 11 | 0.095 | XXX |
| 12 | 0.030 | XX |
| 13 | 0.070 | XXX |
| 14 | 0.062 | XXX |
| 15 | 0.067 | XXX |
| 16 | 0.205 | XXXXXX |
| 17 | 0.185 | XXXXXX |
| 18 | 0.249 | XXXXXXX |
| 19 | 0.297 | XXXXXXXX |
| 20 | 0.279 | XXXXXXXX |
| 21 | 0.245 | XXXXXXX |
| 22 | 0.186 | XXXXXX |

## 4. DISCUSSION, CONCLUSION, AND RECOMENDATION

### 4.1 Discussion \& conclusion:

At the beginning of the analysis, the data had some missing values; these missing values in Zarka River were less than the values in Samra's effluent. The only missing data in Zarka River were in December 1999 and it was in the TSS variable, but in Samra's effluent there were nine missing data, and all of them were only in the year 2000 and they were in different types of Samra's effluent variables.

Missing data were estimated by taking the average of the same months. After substituting the missing data, the data of Zarka River were plotted against time in a scatter diagram. The data showed some abnormal observations, which had to be discussed either as real or not. Some observations were real data and kept as is. Other data could not be explained and were considered as outliers.

The abnormality in the real abnormal observed data were explained, this explanation was due to an external condition that influenced the value of the data, such condition is the huge amount of rainfall that the kingdom received in some months during 1992 and other years (see appendix (1)). In these months, the rainfall was mixed with water in Zarka River \& decreased its variables, such as; BOD5, COD, TSS, T-P, and T-N, in the other way it increased the amount of Zarka River flow. The other observed data, which could not have any explanation why this value is abnormal, was treated as an outlier and assumed that the abnormal value was due to human error, such as: human error in reading or calculating the value, equipment error, or any other errors related to human, and it was treated as missing data (average of same months).

Next, normality test was used to plot the normality test of the data; Weibull distribution was used in plotting the histogram of the normality test for each variable. There are other normality histogram plots that could have been used, such as California, Hazen, Beard and other distributions. But Weibull distinguished in its simplicity and it does not have a $100 \%$ probability.

The other three normality tests were; the coefficient of variation, which is a preliminary test to determine the horizontal skewness, the Kurtosis test, which is a good test in determining the data normality in the vertical direction, and finally the ShapiroWilk test; which showed good results in the normality analysis.

From the Kurtosis and Shapiro -Wilk tests, it has been evident that the two tests had almost identical results in most data quarters, such as; the first quarter of TSS variable, the fourth quarter of BOD5, and other quarters. This proves that the ShapiroWilk test is a good method for testing the normality in the vertical direction

The five variables (TSS, BOD5, COD, T-P, and T-N) had a normal distribution, but the flow variable showed abnormality due to the variation in data and the large amount of abnormal observations ( 9 in total). These abnormal observations were kept and used as it is since they were real data. A lognormal test was made to the flow variable; the lognormal test showed a big difference in data normality between the flow and the $\log$ flow. The Weibull distribution model was skewed to right in the flow variable distribution, the coefficient of variation was more than one, and the Kurtosis
coefficient reached a value of 20 . The lognormal test decreased the skewness and made the data of the flow variable better in normality.

After analyzing the data and finding 16-forecasted values for each variable, the percentage of mean error was calculated. For the TSS variable it was shown that the least percentage of mean error was in the exponential growth method, with an error of $1.7 \%$, but this model's method is a deterministic one, the least percentage of mean error in the stochastic model for TSS variable is an $\operatorname{AR}(1)$, which equals to $5.4 \%$, and this was the best method to be used. ARIMA $(1,0,4)$ was also a good model in forecasting the TSS with a mean error of $8.2 \%$.

In forecasting the BOD5 variable, none of the models have satisfied the $10 \%$ error. The least percentage of mean error for $\mathrm{BOD}_{5}$ variable was in $\operatorname{ARMA}(1,3)$, which equals to $16.1 \%$.

The stochastic model for COD variable did not satisfy the $10 \%$ error, thus, none of the stochastic methods was used in forecasting the COD variable, but in the deterministic modeling, the quadratic method gave the least percentage of mean error which equaled to $3.8 \%$, and it is the best model to be used in forecasting the COD.

For the T-P variable, the only method that satisfied the $10 \%$ error was the moving average, $\mathrm{MA}(4)$, which gave a percentage of mean error equals to $8.5 \%$, so the moving average $\mathrm{MA}(4)$ is the best method to be used in forecasting the T-P variable. In forecasting the T-N variable, many methods have satisfied the $10 \%$ of the mean error, the method that gave the least error was the linear method with an error of $3.3 \%$, but the
best model to be used in forecasting the T-N variable is $\operatorname{ARIMA}(1,2,5)$, which gave $4.8 \%$ of mean error.

Finally, for the flow variable, the forecasting was made for $\log (\mathrm{Q})$ variable. Many models have satisfied the $10 \%$ mean error. However, the computer software (Minitab13) could not estimate the values of ARIMA $(1,2,6)$, because the coefficients of ARIMA model ( $\mathrm{p}, \mathrm{d}, \mathrm{q}$ ) should be less than 5 in order to be estimated. The best method that gives the least error and to be used in forecasting is $\operatorname{AR}(1)$, which gives mean error of $4.9 \%$.

From the above results it can be concluded that the ARIMA model has satisfied the forecasting model for most of the variables. The least amount of mean error resulting from ARIMA model was when calculating the percentage of mean error for ARIMA(1,2,5), which was $4.8 \%$.

The cross and distance correlation gave a good indication about the relations between variables in Zarka River it self and between the Zarka River with Samra's effluent.

- From the relation between the flow and the TSS in Zarka River, it can be concluded that the relation depends on the season, whether it's summer or winter. At winter, the amount of TSS exerted from rainfall in the run off of Zarka River catchment area was high. But in summer, the amount of TSS was low since the suspended solids had been precipitated.
- From the relation between the flow and BOD5 in Zarka River, it can be concluded that the flow coming from the northern south direction has the ability to dilute the Samra's affluent and decrease the amount of BOD5.
- From the relation between the flow and COD in Zarka River, it can be seen that it has the same relation between the flow and BOD5, but with less correlation, that is because the BOD5 measures only the organic matters. Where as, the COD measures organic and inorganic matters, so the variation in COD with flow would be less.
- The relation between the flow and T-P has the same correlation with the flow and TSS. This indicates that most of the phosphorus in Zarka River is in the form of suspended solids.
- From the relation between the flow and T-N, it can be seen that it has the same correlation between the flow rate with $\mathrm{BOD}_{5}$ and COD . This is because the source of T $\mathrm{N}, \mathrm{BOD} 5$, and COD are from the domestic waste. In some cases the relation between the flow and T-N is proportional because some T-N is exerted in the flow of Samra's effluent.
- From the relation between TSS with BOD5 and COD, it can be concluded that most of the organic and inorganic matters are in the form of solids. The correlation between TSS and COD is less than correlation between TSS and BOD5, because the COD variable has more components than the BOD5 variable.
- From the relation between TSS and T-P, it can be concluded that in winter most of the phosphorus is in a solid form. Where as in summer it will be dissolved in Samra's affluent.
- From the relations between the BOD5, COD, T-P, and T-N in Zarka River, it can be seen that all of them have a proportional correlation, so this ensures that the source of these variables is mostly the wastewater.
- From the relation between the flow in Zarka River and the flow from Samra's effluent, it can be concluded that the time of the huge flow in Zarka River, which is mostly between December and March, differs from the huge releases in Samra's effluent, which is mostly between June and July.
- From the relations between the BOD5, COD, T-P, and T-N in Zarka River and Samra's effluent, it can be seen that in each variable the quality of water in Samra's effluent affects positively the quality of water in Zarka River.


### 4.2 Recommendations:

1. Improving the quality of wastewater in Samra's effluent, which affects mainly Zarka River water quality. This improvement could be achieved by increasing the capacity of Al Samra WWTP with decreasing the concentration of the variables (TSS, BOD5, COD, T-P, and T-N) of the Samra's effluent.
2. Finding alternatives and solutions for the usage of Al Samra's effluent and trying to decrease the amount of the effluent that goes to the Zarka River
3. The necessity of studying the environmental impact assessment (EIA) for future projects in Zarka River catchment area, which could influence the water quality in this River; such as: Industrial factories, agricultural projects, and wastewater treatment plants.
4. Decreasing the amount of TSS in Zarka River exerted from soil erosion by rainfall. Also cleaning the precipitated solids in King Talal Reservoir, taking into consideration that the amount of precipitated solids reach up to 13.0 million cubic meters in the year 2000. (RSS reports, 2001).
5. A study should be made to find the correlation between the cross and distance correlation in the vertical (depth) and horizontal directions, so that the variables
achieved from one site could be used to know the similar variables in any other site along the whole track of the flow.
6. Decreasing the amount of phosphorus and Nitrogen, because there presence in water in a concentration of $300 \mathrm{mg} / \mathrm{m} 3$ for phosphorus and $5 \mathrm{mg} / \mathrm{m} 3$ for nitrogen will create the eutrophication process and encourage the algae to build up. The algae is already exist in Zarka River, decreasing the concentration of T-P and TN using Macrophytes, which lives on algae will help in minimizing the amount of algae.
7. For the time series analysis, as increasing the amount of months, as the time series forecasting will be more accurate. A thirteen years data collected is not enough to show the trend and seasonal affect in a proper way.
8. The values of the variable should be very accurate, since the samples are taken to Amman and then tested there, this can decrease the accuracy of the data. A site laboratory will increase the accuracy of the data.
9. The time of sample should be in the peak hours; usually the samples were taken in the working day (between 8 and 3 ), which is not necessary to be the peak hours. Taking three or more daily readings for each point will increase the accuracy of the data.
10. More information about any strange reading should be recorded, so that any reading to be considered as an outlier should be justified by comprehensive reasons.
11. Decreasing the amount of missing data will help in decreasing the randomness in data.
12. There should be Jordanian specifications for Irrigation and Agriculture instead of using the FAO specifications or any other specifications in analyzing the water quality.

## APPENDICES

Appendix (1): The amount of rainfall in Al Zarka River catchment area

## Amount of Rainfall in MCM/Month

| Year | Aan. | Feb. | March | April | May | June |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Jannt of Rainfall in MCM/Month |  |  |  |  |  |
| 1988 | 0.051 | 0.207 | 0.135 | 0.020 | 0.000 | 0.000 |
| 1989 | 0.073 | 0.050 | 0.093 | 0.000 | 0.000 | 0.000 |
| 1990 | 0.102 | 0.045 | 0.084 | 0.035 | 0.000 | 0.000 |
| 1991 | 0.070 | 0.048 | 0.058 | 0.015 | 0.000 | 0.000 |
| 1992 | 0.267 | 0.466 | 0.054 | 0.000 | 0.003 | 0.008 |
| 1993 | 0.115 | 0.493 | 0.045 | 0.002 | 0.014 | 0.000 |
| 1994 | 0.164 | 0.054 | 0.076 | 0.008 | 0.000 | 0.007 |
| 1995 | 0.021 | 0.129 | 0.033 | 0.018 | 0.000 | 0.000 |
| 1996 | 0.157 | 0.015 | 0.218 | 0.016 | 0.000 | 0.000 |
| 1997 | 0.152 | 0.190 | 0.178 | 0.012 | 0.014 | 0.000 |
| 1998 | 0.184 | 0.065 | 0.258 | 0.000 | 0.000 | 0.000 |
| 1999 | 0.128 | 0.095 | 0.049 | 0.017 | 0.000 | 0.000 |
| 2000 | 0.126 | 0.046 | 0.054 | 0.000 | 0.000 | 0.000 |


|  | Amount of Rainfall in MCM/Month |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | July | Aug. | Sep. | Oct. | Nov. | Dec. |  |
| 1988 | 0.020 | 0.000 | 0.000 | 0.003 | 0.026 | 0.149 |  |
| 1989 | 0.000 | 0.000 | 0.000 | 0.003 | 0.031 | 0.059 |  |
| 1990 | 0.000 | 0.000 | 0.000 | 0.005 | 0.016 | 0.001 |  |
| 1991 | 0.000 | 0.000 | 0.000 | 0.000 | 0.032 | 0.162 |  |
| 1992 | 0.000 | 0.000 | 0.000 | 0.000 | 0.088 | 0.225 |  |
| 1993 | 0.000 | 0.000 | 0.000 | 0.004 | 0.012 | 0.011 |  |
| 1994 | 0.000 | 0.000 | 0.000 | 0.005 | 0.200 | 0.142 |  |
| 1995 | 0.000 | 0.000 | 0.000 | 0.000 | 0.032 | 0.034 |  |
| 1996 | 0.000 | 0.000 | 0.000 | 0.013 | 0.023 | 0.071 |  |
| 1997 | 0.000 | 0.000 | 0.000 | 0.030 | 0.068 | 0.154 |  |
| 1998 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.009 |  |
| 1999 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.007 |  |
| 2000 | 0.000 | 0.000 | 0.000 | 0.027 | 0.006 | 0.063 |  |

## Appendix (2):Coefficients ( $\mathbf{a}_{\mathrm{N}-\mathrm{It}}$ ) for Shapiro-Wilk $W$-test of normality

| $\mathbf{i} / \mathbf{n}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ | $\mathbf{6}$ | $\mathbf{7}$ | $\mathbf{8}$ | $\mathbf{9}$ | $\mathbf{1 0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.7071 | 0.7071 | 0.6872 | 0.6646 | 0.6431 | 0.6233 | 0.6052 | 0.5888 | 0.5739 |
| 2 | - | 0.0000 | 0.1677 | 0.2413 | 0.2806 | 0.0310 | 0.3164 | 0.3244 | 0.3291 |
| 3 | - | - | - | 0.0000 | 0.0875 | 0.1401 | 0.1743 | 0.1976 | 0.2141 |
| 4 | - | - | - | - | - | 0.0000 | 0.0561 | 0.0947 | 0.1224 |
| 5 | - | - | - | - | - | - | - | 0.0000 | 0.0399 |


| $\mathbf{i} / \mathbf{n}$ | $\mathbf{1 1}$ | $\mathbf{1 2}$ | $\mathbf{1 3}$ | $\mathbf{1 4}$ | $\mathbf{1 5}$ | $\mathbf{1 6}$ | $\mathbf{1 7}$ | $\mathbf{1 8}$ | $\mathbf{1 9}$ | $\mathbf{2 0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.5601 | 0.5475 | 0.5359 | 0.5251 | 0.5150 | 0.5056 | 0.4968 | 0.4886 | 0.4808 | 0.4734 |
| 2 | 0.3315 | 0.3325 | 0.3325 | 0.3318 | 0.3306 | 0.3290 | 0.3273 | 0.3253 | 0.3232 | 0.3211 |
| 3 | 0.2260 | 0.2347 | 0.2412 | 0.2460 | 0.2495 | 0.2521 | 0.2540 | 0.2553 | 0.2561 | 0.2565 |
| 4 | 0.1429 | 0.1586 | 0.1707 | 0.1802 | 0.1878 | 0.1939 | 0.1988 | 0.2027 | 0.2059 | 0.2085 |
| 5 | 0.0695 | 0.0922 | 0.1099 | 0.1240 | 0.1353 | 0.1447 | 0.1524 | 0.1587 | 0.1641 | 0.1686 |
| 6 | 0.0000 | 0.0303 | 0.0539 | 0.0727 | 0.0880 | 0.1005 | 0.1109 | 0.1197 | 0.1271 | 0.1334 |
| 7 | - | - | 0.0000 | 0.0240 | 0.0433 | 0.0593 | 0.0725 | 0.0837 | 0.0932 | 0.1013 |
| 8 | - | - | - | - | 0.0000 | 0.0196 | 0.0359 | 0.0496 | 0.0621 | 0.0711 |
| 9 | - | - | - | - | - | - | 0.0000 | 0.0163 | 0.0303 | 0.0422 |
| 10 | - | - | - | - | - | - | - | - | 0.0000 | 0.0140 |

Cont. Appendix (2):Coefficients ( $\mathbf{a}_{\mathrm{N}-\mathrm{I}+1}$ ) for Shapiro-Wilk $W$-test of normality

| $\mathbf{i} / \mathbf{n}$ | $\mathbf{2 1}$ | $\mathbf{2 2}$ | $\mathbf{2 3}$ | $\mathbf{2 4}$ | $\mathbf{2 5}$ | $\mathbf{2 6}$ | $\mathbf{2 7}$ | $\mathbf{2 8}$ | $\mathbf{2 9}$ | $\mathbf{3 0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.4643 | 0.4590 | 0.4542 | 0.4493 | 0.4450 | 0.4407 | 0.4366 | 0.4328 | 0.4291 | 0.4254 |
| 2 | 0.3185 | 0.3156 | 0.3126 | 0.3098 | 0.3069 | 0.3043 | 0.3018 | 0.2992 | 0.2968 | 0.2944 |
| 3 | 0.2578 | 0.2571 | 0.2563 | 0.2554 | 0.2543 | 0.2533 | 0.2522 | 0.2510 | 0.2499 | 0.2487 |
| 4 | 0.2119 | 0.2131 | 0.2139 | 0.2145 | 0.2148 | 0.2151 | 0.2152 | 0.2151 | 0.2150 | 0.2148 |
| 5 | 0.1736 | 0.1764 | 0.1787 | 0.1807 | 0.1822 | 0.1836 | 0.1848 | 0.1857 | 0.1864 | 0.1870 |
| 6 | 0.1399 | 0.1443 | 0.1480 | 0.1512 | 0.1539 | 0.1563 | 0.1584 | 0.1601 | 0.1616 | 0.1630 |
| 7 | 0.1092 | 0.1150 | 0.1201 | 0.1245 | 0.1283 | 0.1316 | 0.1346 | 0.1372 | 0.1395 | 0.1415 |
| 8 | 0.0804 | 0.0878 | 0.0941 | 0.0997 | 0.1046 | 0.1089 | 0.1128 | 0.1162 | 0.1192 | 0.1219 |
| 9 | 0.0530 | 0.0618 | 0.0696 | 0.0764 | 0.0823 | 0.0876 | 0.0923 | 0.0965 | 0.1002 | 0.1036 |
| 10 | 0.0263 | 0.0368 | 0.0459 | 0.0539 | 0.0610 | 0.0672 | 0.0728 | 0.0778 | 0.0822 | 0.0862 |
| 11 | 0.0000 | 0.0122 | 0.0228 | 0.0321 | 0.0403 | 0.0476 | 0.0540 | 0.0598 | 0.0650 | 0.0697 |
| 12 | - | - | 0.0000 | 0.0107 | 0.0200 | 0.0284 | 0.0358 | 0.0424 | 0.0483 | 0.0537 |
| 13 | - | - | - | - | 0.0000 | 0.0094 | 0.0178 | 0.0253 | 0.0320 | 0.0381 |
| 14 | - | - | - | - | - | - | 0.0000 | 0.0084 | 0.0159 | 0.0227 |
| 15 | - | - | - | - | - | - | - | - | 0.0000 | 0.0076 |

Cont. Appendix (2):Coefficients ( $\mathbf{a}_{\mathrm{N}-\mathrm{I}+1}$ ) for Shapiro-Wilk $\boldsymbol{W}$-test of normality

| $\mathbf{i} / \mathbf{n}$ | $\mathbf{3 1}$ | $\mathbf{3 2}$ | $\mathbf{3 3}$ | $\mathbf{3 4}$ | $\mathbf{3 5}$ | $\mathbf{3 6}$ | $\mathbf{3 7}$ | $\mathbf{3 8}$ | $\mathbf{3 9}$ | $\mathbf{4 0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.4220 | 0.4188 | 0.4156 | 0.4127 | 0.4096 | 0.4068 | 0.4040 | 0.4015 | 0.3989 | 0.3964 |
| 2 | 0.2921 | 0.2898 | 0.2876 | 0.2854 | 0.2834 | 0.2813 | 0.2794 | 0.2774 | 0.2755 | 0.2737 |
| 3 | 0.2475 | 0.2463 | 0.2451 | 0.2439 | 0.2427 | 0.2415 | 0.2403 | 0.2391 | 0.2380 | 0.2368 |
| 4 | 0.2145 | 0.2141 | 0.2137 | 0.2132 | 0.2127 | 0.2121 | 0.2116 | 0.2110 | 0.2104 | 0.2098 |
| 5 | 0.1874 | 0.1878 | 0.1880 | 0.1882 | 0.1883 | 0.1883 | 0.1883 | 0.1881 | 0.1880 | 0.1878 |
| 6 | 0.1641 | 0.1651 | 0.1660 | 0.1667 | 0.1673 | 0.1678 | 0.1683 | 0.1686 | 0.1689 | 0.1691 |
| 7 | 0.1433 | 0.1449 | 0.1463 | 0.1475 | 0.1487 | 0.1496 | 0.1503 | 0.1513 | 0.1520 | 0.1526 |
| 8 | 0.1243 | 0.1265 | 0.1284 | 0.1301 | 0.1317 | 0.1331 | 0.1344 | 0.1356 | 0.1366 | 0.1376 |
| 9 | 0.1066 | 0.1093 | 0.1118 | 0.1140 | 0.1160 | 0.1179 | 0.1196 | 0.1211 | 0.1250 | 0.1237 |
| 10 | 0.0899 | 0.0931 | 0.0961 | 0.0988 | 0.1013 | 0.1036 | 0.1056 | 0.1075 | 0.1092 | 0.1108 |
| 11 | 0.0739 | 0.0777 | 0.0812 | 0.0844 | 0.0873 | 0.0900 | 0.0924 | 0.0947 | 0.0967 | 0.0896 |
| 12 | 0.0585 | 0.0629 | 0.0669 | 0.0706 | 0.0739 | 0.0770 | 0.0798 | 0.0824 | 0.0848 | 0.0870 |
| 13 | 0.0435 | 0.0485 | 0.0530 | 0.0572 | 0.0610 | 0.0645 | 0.0677 | 0.0706 | 0.0733 | 0.0759 |
| 14 | 0.0289 | 0.0344 | 0.0395 | 0.0441 | 0.0484 | 0.0523 | 0.0559 | 0.0592 | 0.0622 | 0.0651 |
| 15 | 0.0144 | 0.0206 | 0.0262 | 0.0314 | 0.0361 | 0.0404 | 0.0444 | 0.0481 | 0.0515 | 0.0546 |
| 16 | 0.0000 | 0.0068 | 0.0131 | 0.0187 | 0.0239 | 0.0287 | 0.0331 | 0.0372 | 0.0409 | 0.0444 |
| 17 | - | - | 0.0000 | 0.0062 | 0.0119 | 0.0172 | 0.0220 | 0.0264 | 0.0305 | 0.0343 |
| 18 | - | - | - | - | 0.0000 | 0.0057 | 0.0110 | 0.0158 | 0.0203 | 0.0244 |
| 19 | - | - | - | - | - | - | 0.0000 | 0.0053 | 0.0101 | 0.0146 |
| 20 | - | - | - | - | - | - | - | - | 0.0000 | 0.0049 |
| 10 |  |  |  |  |  |  |  |  |  |  |

## Cont. Appendix (2):Coefficients ( $\mathbf{a}_{\mathrm{N}-\mathrm{I}+1}$ ) for Shapiro-Wilk W-test of normality

| $\mathbf{i} / \mathbf{n}$ | $\mathbf{4 1}$ | $\mathbf{4 2}$ | $\mathbf{4 3}$ | $\mathbf{4 4}$ | $\mathbf{4 5}$ | $\mathbf{4 6}$ | $\mathbf{4 7}$ | $\mathbf{4 8}$ | $\mathbf{4 9}$ | $\mathbf{5 0}$ |
| ---: | :---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.3940 | 0.3917 | 0.3894 | 0.3872 | 0.3850 | 0.3830 | 0.3808 | 0.3789 | 0.3000 | 0.3751 |
| 2 | 0.2719 | 0.2701 | 0.2684 | 0.2667 | 0.2651 | 0.2635 | 0.2620 | 0.2604 | 0.2589 | 0.2574 |
| 3 | 0.2357 | 0.2345 | 0.2334 | 0.2323 | 0.2313 | 0.2302 | 0.2291 | 0.2281 | 0.2271 | 0.2260 |
| 4 | 0.2091 | 0.2085 | 0.2078 | 0.2072 | 0.2065 | 0.2058 | 0.2052 | 0.2045 | 0.2038 | 0.2032 |
| 5 | 0.1876 | 0.1874 | 0.1871 | 0.1868 | 0.1865 | 0.1862 | 0.1859 | 0.1855 | 0.1851 | 0.1847 |
| 6 | 0.1693 | 0.1694 | 0.1695 | 0.1695 | 0.1695 | 0.1695 | 0.1695 | 0.1693 | 0.1692 | 0.1691 |
| 7 | 0.1531 | 0.1535 | 0.1539 | 0.1542 | 0.1545 | 0.1548 | 0.1550 | 0.1551 | 0.1553 | 0.1554 |
| 8 | 0.1384 | 0.1392 | 0.1398 | 0.1405 | 0.1410 | 0.1415 | 0.1420 | 0.1423 | 0.1427 | 0.1430 |
| 9 | 0.1249 | 0.1259 | 0.1269 | 0.1278 | 0.1286 | 0.1293 | 0.1300 | 0.1306 | 0.1312 | 0.1317 |
| 10 | 0.1123 | 0.1136 | 0.1149 | 0.1160 | 0.1170 | 0.1180 | 0.1189 | 0.1197 | 0.1205 | 0.1212 |
| 11 | 0.1004 | 0.1020 | 0.1035 | 0.1049 | 0.1062 | 0.1073 | 0.1085 | 0.1095 | 0.1105 | 0.1113 |
| 12 | 0.0891 | 0.0909 | 0.0927 | 0.0943 | 0.0959 | 0.0972 | 0.0986 | 0.0998 | 0.1010 | 0.1020 |
| 13 | 0.0782 | 0.0804 | 0.0824 | 0.0842 | 0.0860 | 0.0876 | 0.0892 | 0.0906 | 0.0919 | 0.0932 |
| 14 | 0.0677 | 0.0701 | 0.0724 | 0.0745 | 0.0775 | 0.0785 | 0.0801 | 0.0817 | 0.0832 | 0.0846 |
| 15 | 0.0575 | 0.0602 | 0.0628 | 0.0651 | 0.0673 | 0.0694 | 0.0713 | 0.0731 | 0.0748 | 0.0764 |
| 16 | 0.0476 | 0.0506 | 0.0534 | 0.0560 | 0.0584 | 0.0607 | 0.0628 | 0.0648 | 0.0662 | 0.0685 |
| 17 | 0.0379 | 0.0411 | 0.0442 | 0.0471 | 0.0497 | 0.0522 | 0.0546 | 0.0568 | 0.0588 | 0.0608 |
| 18 | 0.0283 | 0.0318 | 0.0352 | 0.0383 | 0.0412 | 0.0439 | 0.0465 | 0.0489 | 0.0511 | 0.0532 |
| 19 | 0.0188 | 0.0227 | 0.0263 | 0.0296 | 0.0328 | 0.0357 | 0.0385 | 0.0411 | 0.0436 | 0.0459 |
| 20 | 0.0094 | 0.0316 | 0.0175 | 0.0211 | 0.0245 | 0.0277 | 0.0307 | 0.0335 | 0.0361 | 0.0386 |
| 21 | 0.0000 | 0.0045 | 0.0087 | 0.0126 | 0.0163 | 0.0197 | 0.0229 | 0.0259 | 0.0288 | 0.0314 |
| 22 | - | - | 0.0000 | 0.0042 | 0.0081 | 0.0118 | 0.0153 | 0.0185 | 0.0215 | 0.0244 |
| 23 | - | - | - | - | 0.0000 | 0.0039 | 0.0076 | 0.0111 | 0.0143 | 0.0174 |
| 24 | - | - | - | - | - | - | 0.0000 | 0.0037 | 0.0071 | 0.0104 |
| 25 | - | - | - | - | - | - | - | 0.0000 | 0.0035 |  |
| 10 |  |  |  |  |  |  |  |  |  |  |

## Appendix (3): Percentage points of the Shapiro-Wilk W-test

| n | 1 - Confidence Interval |  | n | 1 - Confidence Interval |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.01 | 0.05 |  | 0.01 | 0.05 |
| 3 | 0.753 | 0.767 | 27 | 0.894 | 0.923 |
| 4 | 0.687 | 0.748 | 28 | 0.896 | 0.924 |
| 5 | 0.686 | 0.762 | 29 | 0.898 | 0.926 |
| 6 | 0.713 | 0.788 | 30 | 0.900 | 0.927 |
| 7 | 0.730 | 0.803 | 31 | 0.902 | 0.929 |
| 8 | 0.749 | 0.818 | 32 | 0.904 | 0.930 |
| 9 | 0.764 | 0.829 | 33 | 0.906 | 0.931 |
| 10 | 0.781 | 0.842 | 34 | 0.908 | 0.933 |
| 11 | 0.792 | 0.850 | 35 | 0.910 | 0.934 |
| 12 | 0.805 | 0.859 | 36 | 0.912 | 0.935 |
| 13 | 0.814 | 0.866 | 37 | 0.914 | 0.936 |
| 14 | 0.825 | 0.874 | 38 | 0.914 | 0.938 |
| 15 | 0.835 | 0.881 | 39 | 0.917 | 0.939 |
| 16 | 0.844 | 0.887 | 40 | 0.919 | 0.940 |
| 17 | 0.851 | 0.892 | 41 | 0.920 | 0.941 |
| 18 | 0.858 | 0.897 | 42 | 0.922 | 0.942 |
| 19 | 0.863 | $0.901$ | 43 | 0.923 | 0.943 |
| 20 | 0.868 | 0.905 | 44 | 0.924 | 0.944 |
| 21 | 0.873 | 0.908 | 45 | 0.926 | 0.945 |
| 22 | 0.878 | 0.911 | 46 | 0.927 | 0.945 |
| 23 | 0.881 | 0.914 | 47 | 0.928 | 0.946 |
| 24 | 0.884 | 0.916 | 48 | 0.929 | 0.947 |
| 25 | 0.888 | 0.918 | 49 | 0.929 | 0.947 |
| 26 | 0.891 | 0.920 | 50 | 0.930 | 0.947 |

Appendix (4):

## A4.1: Cross Correlation Function: Zarka Flow MCM/month; TSS Zarka mg/l



## Cont. Appendix (4):

## A4.2: Cross Correlation Function: Zarka Flow MCM/month; BOD5 Zarka mg/l

CCF - correlates Zarka Flow MCM/month(t) and BOD5 Zarka mg/l(t+k)

| -1.0-0.8-0.6-0.4-0.2 0.0 0.2 $0.4 \begin{array}{lllllll}0.6 & 0.8 & 1.0\end{array}$ |  |  |
| :---: | :---: | :---: |
|  | +--- | +----+---+-- |
| -22 | -0.151 | XXXXX |
| -21 | -0.088 | XXX |
| -20 | 0.005 | X |
| -19 | 0.031 | XX |
| -18 | -0.015 | X |
| -17 | -0.025 | XX |
| -16 | 0.036 | XX |
| -15 | 0.005 | X |
| -14 | -0.000 | X |
| -13 | -0.041 | XX |
| -12 | -0.086 | XXX |
| -11 | -0.117 | XXXX |
| -10 | -0.110 | XXXX |
| -9 | -0.083 | XXX |
| -8 | -0.024 | XX |
| -7 | -0.018 | X |
| -6 | -0.047 | XX |
| -5 | -0.015 | X |
| -4 | 0.013 | X |
| -3 | -0.025 | XX |
| -2 | -0.095 | XXX |
| -1 | -0.126 | XXXX |
| 0 | -0.250 | XXXXXXX |
| 1 | -0.224 | XXXXXXX |
| 2 | -0.098 | XXX |
| 3 | -0.053 | XX |
| 4 | -0.020 | XX |
| 5 | 0.062 | XXX |
| 6 | 0.003 | X |
| 7 | -0.029 | XX |
| 8 | -0.096 | XXX |
| 9 | -0.156 | XXXXX |
| 10 | -0.145 | XXXXX |
| 11 | -0.193 | XXXXXX |
| 12 | -0.171 | XXXXX |
| 13 | -0.093 | XXX |
| 14 | -0.009 | X |
| 15 | 0.077 | XXX |
| 16 | 0.096 | XXX |
| 17 | 0.167 | XXXXX |
| 18 | 0.160 | XXXXX |
| 19 | 0.119 | XXXX |
| 20 | 0.073 | XXX |
| 21 | 0.096 | XXX |
| 22 | 0.060 | XXX |

## Cont. Appendix (4):

A4.3: Cross Correlation Function: Zarka Flow MCM/month; COD Zarka mg/l

CCF - correlates Zarka Flow MCM/month( t ) and COD Zarka mg/l(t+k)


## Cont. Appendix (4):

A4.4: Cross Correlation Function: Zarka Flow MCM/month; T-P Zarka mg/l

CCF - correlates Zarka Flow MCM/month(t) and T-P Zarka mg/l $(\mathrm{t}+\mathrm{k})$

| $\begin{array}{r} -1.0-1 \\ +-- \end{array}$ |  |
| :---: | :---: |
| -22 -0.127 | XXXX |
| -21-0.079 | XXX |
| -20 0.015 | X |
| -19 0.052 | XX |
| -18 0.101 | XXXX |
| -17 0.110 | XXXX |
| -16 0.099 | XXX |
| -15 0.121 | XXXX |
| -14 0.090 | XXX |
| -13 0.024 | XX |
| -12-0.103 | XXXX |
| -11-0.208 | XXXXXX |
| -10-0.185 | XXXXXX |
| -9 -0.159 | XXXXX |
| -8-0.116 | XXXX |
| -7-0.028 | XX |
| -6 0.009 | X |
| -5 0.075 | XXX |
| -4 0.063 | XXX |
| -3 0.054 | XX |
| -2 -0.079 | XXX |
| -1-0.250 | XXXXXXX |
| 0 -0.409 | XXXXXXXXXXX |
| $1-0.462$ | XXXXXXXXXXXXX |
| $2-0.368$ | XXXXXXXXXX |
| $3-0.309$ | XXXXXXXXX |
| $4-0.195$ | XXXXXX |
| $5-0.126$ | XXXX |
| $6-0.048$ | XX |
| $7-0.008$ | X |
| $8-0.000$ | X |
| 9-0.047 | XX |
| $10-0.141$ | XXXXX |
| $11-0.159$ | XXXXX |
| $12-0.222$ | XXXXXXX |
| $13-0.233$ | XXXXXXX |
| $14-0.221$ | XXXXXXX |
| $15-0.116$ | XXXX |
| $16-0.032$ | XX |
| 170.068 | XXX |
| 180.128 | XXXX |
| 190.149 | XXXXX |
| 200.115 | XXXX |
| 210.110 | XXXX |
| 220.032 | XX |

## Cont. Appendix (4):

## A4.5: Cross Correlation Function: Zarka Flow MCM/month; T-N Zarka mg/l

CCF - correlates Zarka Flow MCM/month( t ) and T-N Zarka mg/l(t+k)

|  |  |
| :---: | :---: |
| -22 -0.130 | XXXX |
| -21-0.142 | XXXXX |
| -20-0.124 | XXXX |
| -19-0.148 | XXXXX |
| -18-0.148 | XXXXX |
| -17-0.156 | XXXXX |
| -16-0.069 | XXX |
| -15-0.022 | XX |
| -14 0.013 | X |
| -13-0.008 | X |
| -12-0.058 | XX |
| -11-0.101 | XXXX |
| -10-0.147 | XXXXX |
| -9 -0.155 | XXXXX |
| -8 -0.179 | XXXXX |
| -7 -0.184 | XXXXXX |
| -6-0.154 | XXXXX |
| -5 -0.130 | XXXX |
| -4 -0.094 | XXX |
| -3 -0.036 | XX |
| -2 -0.035 | XX |
| -1 -0.114 | XXXX |
| $0-0.214$ | XXXXXX |
| $1-0.259$ | XXXXXXX |
| $2-0.211$ | XXXXXX |
| $3-0.230$ | XXXXXXX |
| 4-0.206 | XXXXXX |
| $5-0.232$ | XXXXXXX |
| $6-0.246$ | XXXXXXX |
| $7-0.164$ | XXXXX |
| 8-0.116 | XXXX |
| $9-0.108$ | XXXX |
| $10-0.126$ | XXXX |
| $11-0.151$ | XXXXX |
| $12-0.151$ | XXXXX |
| $13-0.153$ | XXXXX |
| $14-0.137$ | XXXX |
| $15-0.122$ | XXXX |
| $16-0.103$ | XXXX |
| $17-0.060$ | XX |
| $18-0.034$ | XX |
| $19-0.017$ | X |
| $20 \quad 0.024$ | XX |
| 210.069 | XXX |
| 220.069 | XXX |

## Cont. Appendix (4):

## A4.6: Cross Correlation Function: TSS mg/l ; BOD5 mg/l for Zarka River

| CCF - correlates TSS | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and BOD5 $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: |
|  |  |
| -22 0.153 | XXXXX |
| -21 0.008 | X |
| -20-0.028 | XX |
| -19-0.021 | XX |
| -18 0.016 | X |
| -17 0.037 | XX |
| -16 0.210 | XXXXXX |
| -15 0.105 | XXXX |
| -14 0.175 | XXXXX |
| -13 0.138 | XXXX |
| $\begin{array}{lll}-12 & 0.181\end{array}$ | XXXXXX |
| -11 0.088 | XXX |
| -10-0.047 | XX |
| -9 -0.026 | XX |
| -8 0.036 | XX |
| -7 0.049 | XX |
| -6 0.035 | XX |
| -5 0.088 | XXX |
| -4 0.121 | XXXX |
| -3 0.159 | XXXXX |
| -2 0.251 | XXXXXXX |
| -1 0.291 | XXXXXXXX |
| $0 \quad 0.394$ | XXXXXXXXXXX |
| 10.130 | XXXX |
| $2-0.001$ | X |
| 30.082 | XXX |
| 40.008 | X |
| $5-0.031$ | XX |
| 60.039 | XX |
| 70.119 | XXXX |
| 80.125 | XXXX |
| 90.096 | XXX |
| 100.246 | XXXXXXX |
| 110.184 | XXXXXX |
| 120.155 | XXXXX |
| 130.118 | XXXX |
| 140.049 | XX |
| $15-0.082$ | XXX |
| $16-0.043$ | XX |
| $17 \quad 0.024$ | XX |
| 180.055 | XX |
| 190.073 | XXX |
| $20 \quad 0.042$ | XX |
| 210.196 | XXXXXX |
| 220.214 | XXXXXX |

## Cont. Appendix (4):

A4.7: Cross Correlation Function: TSS mg/l; COD mg/l in Zarka River

| CCF - correlates TSS | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and COD $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: |
| $\begin{array}{r} -1.0-0.8 \text {-0.6-0. } \\ +---+-------+~ \end{array}$ |  |
| -22 0.122 | XXXX |
| $\begin{array}{lll}-21 & 0.011\end{array}$ | X |
| -20-0.018 | X |
| -19-0.026 | XX |
| -18 0.008 | X |
| -17-0.049 | XX |
| -16 0.044 | XX |
| -15 0.051 | XX |
| -14 0.134 | XXXX |
| $\begin{array}{lll}-13 & 0.121\end{array}$ | XXXX |
| -12 0.147 | XXXXX |
| -11 0.135 | XXXX |
| -10 0.071 | XXX |
| -9 0.006 | X |
| -8 -0.018 | X |
| -7 -0.030 | XX |
| -6 -0.039 | XX |
| -5 -0.045 | XX |
| -4 -0.009 | X |
| -3 0.064 | XXX |
| -2 0.114 | XXXX |
| -1 0.254 | XXXXXXX |
| 00.338 | XXXXXXXXX |
| 10.272 | XXXXXXXX |
| 20.093 | XXX |
| 30.021 | XX |
| 40.070 | XXX |
| $5-0.057$ | XX |
| 6 -0.140 | XXXXX |
| 7 -0.024 | XX |
| 80.009 | X |
| $9-0.058$ | XX |
| 100.109 | XXXX |
| 110.228 | XXXXXXX |
| 120.205 | XXXXXX |
| 130.204 | XXXXXX |
| 140.064 | XXX |
| 150.029 | XX |
| $16-0.082$ | XXX |
| $17-0.053$ | XX |
| $18-0.040$ | XX |
| $19-0.023$ | XX |
| $20-0.174$ | XXXXX |
| $21-0.037$ | XX |
| 220.024 | XX |

## Cont. Appendix (4):

A4.8: Cross Correlation Function: TSS mg/l; T-P mg/l in Zarka River

| CCF - correlates TSS | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and T-P $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: |
|  |  |
| -22 0.153 | XXXXX |
| -21 0.037 | XX |
| -20 0.085 | XXX |
| -19-0.038 | XX |
| -18-0.058 | XX |
| -17-0.073 | XXX |
| -16-0.002 | X |
| -15 0.099 | XXX |
| -14 0.171 | XXXXX |
| -13 0.279 | XXXXXXXX |
| -12 0.203 | XXXXXX |
| $\begin{array}{lll}-11 & 0.235\end{array}$ | XXXXXXX |
| -10 0.226 | XXXXXXX |
| -9 0.141 | XXXXX |
| -8 -0.068 | XXX |
| -7-0.071 | XXX |
| -6 -0.081 | XXX |
| -5 -0.085 | XXX |
| -4 -0.135 | XXXX |
| -3 0.0 .054 | XX |
| -2 0.160 | XXXXX |
| -1 0.291 | XXXXXXXX |
| $0 \quad 0.352$ | XXXXXXXXXX |
| 10.272 | XXXXXXXX |
| 20.154 | XXXXX |
| 30.078 | XXX |
| 40.007 | X |
| $5-0.102$ | XXXX |
| 6 -0.155 | XXXXX |
| $7-0.110$ | XXXX |
| $8-0.036$ | XX |
| 90.056 | XX |
| 100.200 | XXXXXX |
| 110.216 | XXXXXX |
| 120.247 | XXXXXXX |
| 130.233 | XXXXXXX |
| 140.169 | XXXXX |
| 150.080 | XXX |
| $16-0.054$ | XX |
| 17 -0.088 | XXX |
| $18-0.151$ | XXXXX |
| $19-0.098$ | XXX |
| $20-0.074$ | XXX |
| $21-0.048$ | XX |
| 220.071 | XXX |

## Cont. Appendix (4):

## A4.9: Cross Correlation Function: TSS mg/l; T-N mg/l in Zarka River

| CCF | F - correlates TSS | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and T-N $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: | :---: |
|  | $\begin{array}{r} -1.0-0.8 \text {-0.6-0.4 } \\ +---+----+-+. \end{array}$ |  |
| -22 | 0.141 | XXXXX |
| -21 | 0.072 | XXX |
| -20 | 0.095 | XXX |
| -19 | 0.090 | XXX |
| -18 | 0.076 | XXX |
| -17 | 0.021 | XX |
| -16 | 0.032 | XX |
| -15 | 0.012 | X |
| -14 | 0.060 | XXX |
| -13 | 0.098 | XXX |
| -12 | 0.150 | XXXXX |
| -11 | 0.158 | XXXXX |
| -10 | 0.219 | XXXXXX |
| -9 | 0.152 | XXXXX |
| -8 | 0.210 | XXXXXX |
| -7 | 0.147 | XXXXX |
| -6 | 0.166 | XXXXX |
| -5 | 0.117 | XXXX |
| -4 | 0.056 | XX |
| -3 | 0.076 | XXX |
| -2 | 0.096 | XXX |
| -1 | 0.159 | XXXXX |
| 0 | 0.252 | XXXXXXX |
| 1 | 0.151 | XXXXX |
| 2 | 0.183 | XXXXXX |
| 3 | 0.164 | XXXXX |
| 4 | 0.213 | XXXXXX |
| 5 | 0.140 | XXXX |
| 6 | 0.065 | XXX |
| 7 | 0.047 | XX |
| 8 | 0.023 | XX |
| 9 | 0.054 | XX |
| 10 | 0.101 | XXXX |
| 11 | 0.103 | XXXX |
| 12 | 0.169 | XXXXX |
| 13 | 0.155 | XXXXX |
| 14 | 0.119 | XXXX |
| 15 | 0.099 | XXX |
| 16 | 0.125 | XXXX |
| 17 | 0.109 | XXXX |
| 18 | 0.082 | XXX |
| 19 | -0.001 | X |
| 20 | -0.007 | X |
| 21 | -0.000 | X |
| 22 | 0.083 | XXX |

## Cont. Appendix (4):

## A4.10: Cross Correlation Function: BOD5 mg/l; COD mg/l in Zarka River

| CCF - correlates BOD5 |  | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and COD $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: | :---: |
|  |  | $\begin{array}{lllll} 0.2 & 0.0 & 0.2 & 0.4 & 0.6 \\ --+--------------------------------~ \end{array}$ |
| -22 | 0.086 | XXX |
| -21 | 0.046 | XX |
| -20 | 0.044 | XX |
| -19 | 0.060 | XXX |
| -18 | 0.099 | XXX |
| -17 | 0.053 | XX |
| -16 | 0.047 | XX |
| -15 | 0.041 | XX |
| -14 | 0.028 | XX |
| -13 | 0.117 | XXXX |
| -12 | 0.177 | XXXXX |
| -11 | 0.233 | XXXXXXX |
| -10 | 0.226 | XXXXXXX |
| -9 | 0.194 | XXXXXX |
| -8 | 0.141 | XXXXX |
| -7 | 0.153 | XXXXX |
| -6 | 0.209 | XXXXXX |
| -5 | 0.172 | XXXXX |
| -4 | 0.085 | XXX |
| -3 | 0.171 | XXXXX |
| -2 | 0.166 | XXXXX |
| -1 | 0.251 | XXXXXXX |
| 0 | 0.453 | XXXXXXXXXXXX |
| 1 | 0.403 | XXXXXXXXXXX |
| 2 | 0.351 | XXXXXXXXXX |
| 3 | 0.271 | XXXXXXXX |
| 4 | 0.285 | XXXXXXXX |
| 5 | 0.279 | XXXXXXXX |
| 6 | 0.264 | XXXXXXXX |
| 7 | 0.296 | XXXXXXXX |
| 8 | 0.262 | XXXXXXXX |
| 9 | 0.146 | XXXXX |
| 10 | 0.226 | XXXXXXX |
| 11 | 0.203 | XXXXXX |
| 12 | 0.319 | XXXXXXXXX |
| 13 | 0.320 | XXXXXXXXX |
| 14 | 0.210 | XXXXXX |
| 15 | 0.262 | XXXXXXXX |
| 16 | 0.236 | XXXXXXX |
| 17 | 0.162 | XXXXX |
| 18 | 0.265 | XXXXXXXX |
| 19 | 0.203 | XXXXXX |
| 20 | 0.180 | XXXXXX |
| 21 | 0.194 | XXXXXX |
| 22 | 0.178 | XXXXX |

Cont. Appendix (4):
A4.11: Cross Correlation Function: BOD5 mg/l; T-P mg/l in Zarka River

| CCF - correlates BOD5 |  | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and T-P $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: | :---: |
|  |  |  |
| -22 | 0.159 | XXXXX |
| -21 | 0.157 | XXXXX |
| -20 | 0.191 | XXXXXX |
| -19 | 0.141 | XXXXX |
| -18 | 0.123 | XXXX |
| -17 | 0.063 | XXX |
| -16 | 0.119 | XXXX |
| -15 | 0.075 | XXX |
| -14 | 0.083 | XXX |
| -13 | 0.276 | XXXXXXXX |
| -12 | 0.312 | XXXXXXXXX |
| -11 | 0.334 | XXXXXXXXX |
| -10 | 0.352 | XXXXXXXXXX |
| -9 | 0.349 | XXXXXXXXXX |
| -8 | 0.335 | XXXXXXXXX |
| -7 | 0.285 | XXXXXXXX |
| -6 | 0.282 | XXXXXXXX |
| -5 | 0.216 | XXXXXX |
| -4 | 0.150 | XXXXX |
| -3 | 0.273 | XXXXXXXX |
| -2 | 0.289 | XXXXXXXX |
| -1 | 0.368 | XXXXXXXXXX |
| 0 | 0.451 | XXXXXXXXXXXX |
| 1 | 0.437 | XXXXXXXXXXXX |
| 2 | 0.453 | XXXXXXXXXXXX |
| 3 | 0.379 | XXXXXXXXXX |
| 4 | 0.290 | XXXXXXXX |
| 5 | 0.288 | XXXXXXXX |
| 6 | 0.304 | XXXXXXXXX |
| 7 | 0.263 | XXXXXXXX |
| 8 | 0.152 | XXXXX |
| 9 | 0.215 | XXXXXX |
| 10 | 0.275 | XXXXXXXX |
| 11 | 0.317 | XXXXXXXXX |
| 12 | 0.276 | XXXXXXXX |
| 13 | 0.339 | XXXXXXXXX |
| 14 | 0.263 | XXXXXXXX |
| 15 | 0.278 | XXXXXXXX |
| 16 | 0.293 | XXXXXXXX |
| 17 | 0.203 | XXXXXX |
| 18 | 0.202 | XXXXXX |
| 19 | 0.191 | XXXXXX |
| 20 | 0.180 | XXXXXX |
| 21 | 0.118 | XXXX |
| 22 | 0.183 | XXXXXX |

## Cont. Appendix (4):

## A4.12: Cross Correlation Function: BOD5 mg/l; T-N mg/l in Zarka River

CCF - correlates BOD5 $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and T-N $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$


## Cont. Appendix (4):

A4.13: Cross Correlation Function: COD mg/l; T-P mg/l in Zarka River

| CCF - correlates COD |  | $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and T-P $\mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$ |
| :---: | :---: | :---: |
|  | $\begin{gathered} -1.0-0.8-0.6-0.4 \\ +---+----+--+-1 \end{gathered}$ | $\begin{array}{ccccc} -0.2 & 0.0 & 0.2 & 0.4 & 0.6 \end{array} 0.8 \quad 1.0$ |
| -22 | 0.146 | XXXXX |
| -21 | 0.185 | XXXXXX |
| -20 | 0.193 | XXXXXX |
| -19 | 0.159 | XXXXX |
| -18 | 0.150 | XXXXX |
| -17 | 0.178 | XXXXX |
| -16 | 0.156 | XXXXX |
| -15 | 0.278 | XXXXXXXX |
| -14 | 0.302 | XXXXXXXXX |
| -13 | 0.389 | XXXXXXXXXXX |
| -12 | 0.339 | XXXXXXXXX |
| -11 | 0.379 | XXXXXXXXXX |
| -10 | 0.333 | XXXXXXXXX |
| -9 | 0.283 | XXXXXXXX |
| -8 | 0.203 | XXXXXX |
| -7 | 0.218 | XXXXXX |
| -6 | 0.206 | XXXXXX |
| -5 | 0.264 | XXXXXXXX |
| -4 | 0.280 | XXXXXXXX |
| -3 | 0.351 | XXXXXXXXXX |
| -2 | 0.505 | XXXXXXXXXXXXXX |
| -1 | 0.557 | XXXXXXXXXXXXXXX |
| 0 | 0.644 | XXXXXXXXXXXXXXXXX |
| 1 | 0.474 | XXXXXXXXXXXXX |
| 2 | 0.419 | XXXXXXXXXXX |
| 3 | 0.392 | XXXXXXXXXXX |
| 4 | 0.375 | XXXXXXXXXX |
| 5 | 0.243 | XXXXXXX |
| 6 | 0.198 | XXXXXX |
| 7 | 0.195 | XXXXXX |
| 8 | 0.222 | XXXXXXX |
| 9 | 0.228 | XXXXXXX |
| 10 | 0.287 | XXXXXXXX |
| 11 | 0.327 | XXXXXXXXX |
| 12 | 0.370 | XXXXXXXXXX |
| 13 | 0.313 | XXXXXXXXX |
| 14 | 0.211 | XXXXXX |
| 15 | 0.225 | XXXXXXX |
| 16 | 0.155 | XXXXX |
| 17 | 0.094 | XXX |
| 18 | 0.125 | XXXX |
| 19 | 0.119 | XXXX |
| 20 | -0.017 | X |
| 21 | 0.009 | X |
| 22 | 0.075 | XXX |

## Cont. Appendix (4):

## A4.14: Cross Correlation Function: COD mg/l; T-N mg/l in Zarka River

CCF - correlates COD $\mathrm{mg} / \mathrm{l}(\mathrm{t})$ and $\mathrm{T}-\mathrm{N} \quad \mathrm{mg} / \mathrm{l}(\mathrm{t}+\mathrm{k})$

|  |  |  |
| :---: | :---: | :---: |
| -22 | 0.299 | XXXXXXXX |
| -21 | 0.319 | XXXXXXXXX |
| -20 | 0.324 | XXXXXXXXX |
| -19 | 0.371 | XXXXXXXXXX |
| -18 | 0.392 | XXXXXXXXXXX |
| -17 | 0.329 | XXXXXXXXX |
| -16 | 0.351 | XXXXXXXXXX |
| -15 | 0.315 | XXXXXXXXX |
| -14 | 0.336 | XXXXXXXXX |
| -13 | 0.408 | XXXXXXXXXXX |
| -12 | 0.403 | XXXXXXXXXXX |
| -11 | 0.380 | XXXXXXXXXXX |
| -10 | 0.374 | XXXXXXXXXX |
| -9 | 0.410 | XXXXXXXXXXX |
| -8 | 0.455 | XXXXXXXXXXXX |
| -7 | 0.480 | XXXXXXXXXXXXX |
| -6 | 0.460 | XXXXXXXXXXXXX |
| -5 | 0.469 | XXXXXXXXXXXXX |
| -4 | 0.472 | XXXXXXXXXXXXX |
| -3 | 0.463 | XXXXXXXXXXXXX |
| -2 | 0.482 | XXXXXXXXXXXXX |
| -1 | 0.567 | XXXXXXXXXXXXXXX |
| 0 | 0.561 | XXXXXXXXXXXXXXX |
| 1 | 0.489 | XXXXXXXXXXXXX |
| 2 | 0.461 | XXXXXXXXXXXXX |
| 3 | 0.473 | XXXXXXXXXXXXX |
| 4 | 0.447 | XXXXXXXXXXXX |
| 5 | 0.458 | XXXXXXXXXXXX |
| 6 | 0.422 | XXXXXXXXXXXX |
| 7 | 0.352 | XXXXXXXXXX |
| 8 | 0.351 | XXXXXXXXXX |
| 9 | 0.332 | XXXXXXXXX |
| 10 | 0.369 | XXXXXXXXXX |
| 11 | 0.420 | XXXXXXXXXXXX |
| 12 | 0.408 | XXXXXXXXXXX |
| 13 | 0.294 | XXXXXXXX |
| 14 | 0.282 | XXXXXXXX |
| 15 | 0.288 | XXXXXXXX |
| 16 | 0.307 | XXXXXXXXX |
| 17 | 0.269 | XXXXXXXX |
| 18 | 0.245 | XXXXXXX |
| 19 | 0.129 | XXXX |
| 20 | 0.106 | XXXX |
| 21 | 0.082 | XXX |
| 22 | 0.128 | XXXX |

## Cont. Appendix (4):

## A4.15: Cross Correlation Function: T-P mg/l; T-N mg/l in Zarka River

| CCF - correlates T-P mg/l(t) and T-N mg/l(t+k) |  |  |
| :---: | :---: | :---: |
|  | $\begin{array}{r} -1.0-0.8-0.6-0 . \\ +------+---1 . \end{array}$ | $\begin{array}{lllll}  & -0.2 & 0.0 & 0.2 & 0.4 \\ 0 \end{array}$ |
| -22 | 0.276 | XXXXXXXX |
| -21 | 0.251 | XXXXXXX |
| -20 | 0.319 | XXXXXXXXX |
| -19 | 0.300 | XXXXXXXXX |
| -18 | 0.274 | XXXXXXXX |
| -17 | 0.256 | XXXXXXX |
| -16 | 0.273 | XXXXXXXX |
| -15 | 0.252 | XXXXXXX |
| -14 | 0.302 | XXXXXXXXX |
| -13 | 0.382 | XXXXXXXXXXX |
| -12 | 0.391 | XXXXXXXXXXX |
| -11 | 0.458 | XXXXXXXXXXXX |
| -10 | 0.443 | XXXXXXXXXXXX |
| -9 | 0.467 | XXXXXXXXXXXXX |
| -8 | 0.468 | XXXXXXXXXXXXX |
| -7 | 0.514 | XXXXXXXXXXXXXX |
| -6 | 0.459 | XXXXXXXXXXXX |
| -5 | 0.468 | XXXXXXXXXXXXX |
| -4 | 0.438 | XXXXXXXXXXXX |
| -3 | 0.486 | XXXXXXXXXXXXX |
| -2 | 0.536 | XXXXXXXXXXXXXX |
| -1 | 0.620 | XXXXXXXXXXXXXXXXX |
| 0 | 0.688 | XXXXXXXXXXXXXXXXXX |
| 1 | 0.628 | XXXXXXXXXXXXXXXXX |
| 2 | 0.604 | XXXXXXXXXXXXXXXX |
| 3 | 0.559 | XXXXXXXXXXXXXXX |
| 4 | 0.530 | XXXXXXXXXXXXXX |
| 5 | 0.518 | XXXXXXXXXXXXXX |
| 6 | 0.456 | XXXXXXXXXXXX |
| 7 | 0.432 | XXXXXXXXXXXX |
| 8 | 0.382 | XXXXXXXXXXX |
| 9 | 0.382 | XXXXXXXXXXX |
| 10 | 0.427 | XXXXXXXXXXXX |
| 11 | 0.435 | XXXXXXXXXXXX |
| 12 | 0.457 | XXXXXXXXXXXX |
| 13 | 0.429 | XXXXXXXXXXXX |
| 14 | 0.383 | XXXXXXXXXXX |
| 15 | 0.365 | XXXXXXXXXX |
| 16 | 0.340 | XXXXXXXXXX |
| 17 | 0.327 | XXXXXXXXX |
| 18 | 0.284 | XXXXXXXX |
| 19 | 0.241 | XXXXXXX |
| 20 | 0.212 | XXXXXX |
| 21 | 0.178 | XXXXX |
| 22 | 0.231 | XXXXXXX |

## Cont. Appendix (4):

## A4.16: Distance Correlation Function: Zarka Flow MCM/month; Samra Flow MCM/month

CCF - correlates Zarka Flow MCM/month( t ) and Samra Flow MCM/month $(\mathrm{t}+\mathrm{k})$

|  |  |
| :---: | :---: |
| -22 -0.101 | XXXX |
| -21-0.128 | XXXX |
| -20-0.118 | XXXX |
| -19-0.113 | XXXX |
| -18-0.124 | XXXX |
| -17-0.120 | XXXX |
| -16-0.125 | XXXX |
| -15-0.113 | XXXX |
| -14-0.109 | XXXX |
| -13-0.086 | XXX |
| -12-0.086 | XXX |
| -11-0.099 | XXX |
| -10-0.112 | XXXX |
| -9 -0.118 | XXXX |
| -8-0.119 | XXXX |
| -7 -0.130 | XXXX |
| -6-0.116 | XXXX |
| -5 -0.108 | XXXX |
| -4-0.098 | XXX |
| -3-0.075 | XXX |
| -2 -0.042 | XX |
| -1 -0.051 | XX |
| $0-0.013$ | X |
| $1-0.024$ | XX |
| $2-0.041$ | XX |
| $3-0.055$ | XX |
| $4-0.040$ | XX |
| $5-0.031$ | XX |
| $6-0.035$ | XX |
| 7 -0.040 | XX |
| $8-0.035$ | XX |
| $9-0.042$ | XX |
| $10-0.015$ | X |
| $11-0.017$ | X |
| $12-0.042$ | XX |
| $13-0.036$ | XX |
| $14-0.031$ | XX |
| $15-0.034$ | XX |
| $16-0.033$ | XX |
| $17-0.041$ | XX |
| $18-0.038$ | XX |
| $19-0.045$ | XX |
| $20-0.039$ | XX |
| $21-0.032$ | XX |
| $22-0.003$ | X |

## Cont. Appendix (4):

## A4.17: Distance Correlation Function: TSS Zarka River mg/l; TSS Samra mg/l



## Cont. Appendix (4):

A4.18: Distance Correlation Function: BOD5 in Zarka mg/l; BOD5 Samra mg/l


## Cont. Appendix (4):

A4.19: Distance Correlation Function: COD in Zarka River mg/l; COD Samra mg/l


## Cont. Appendix (4):

A4.20: Distance Correlation Function: T-P in Zarka River mg/l; T-P Samra mg/l


Cont. Appendix (4):
A4.21: Distance Correlation Function: T-N in Zarka River mg/l; T-N Samra mg/l


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# التحليل والتتبؤ بكمية ونوعية المياه الاخلة إلى سد الملك طل 

إعداد
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الاكتور أحمد الجمرة

يعتبر سد الملك طلال من أهم المشاريع المائية التي تم إنثائها لزر اعة مناطق شاسعة في وادي الأردن. و بالنالي فإن دراسة نو عية و كمية المياه في سد الملك طلال و المياه التي تصب فيه يجب أن نكون من الأولويات في الأردن.

تتركز هذه اللراسة في دراسة و نتبؤ نو عية و كمية المياه في نهر الزرقاء و الذي يعتبر الر افد الرئيسي لسد الملك طلالعبّر عن كمية المياه المستعملة في هذا البحث بالنتفق, ونوعية المياه بالمو اد العالقة الكلية, COD , ${ }^{\text {COD }}$, الفسفور الكلي, و النبتروجين الكلي في نهر الزرقاء. البيانات التي جمعت لكل متغير سجلت خلال 107 شهراً من العام 911 ا و حتى نهاية العام . ..... الطريقة المستخدمة في تحليل الستة متغيرات في نهر الزرقاء هي من خلال علاقة المتغير بنفسه (autocorrelation) و المتغبر مع متغير أخر عند نقطة معينة ( و علاقة المتغير مع نفسه و لكن بعد مسافة افقية معينة (distance correlation) . و قد تم استخدام التتبؤ الحتمي (deterministic) و الإحتمالي (stochastic) لستة متغير ات لإيجاد أفضل نموذج للتنبؤ .

نتائج اللاراسة مؤشر إلى أن نموذج ARIMA نوذجاً جيداً في التتبؤ بمعظم الستة متغيرات. في تتؤ قيم ال BOD 5 لم يحقق أي من النماذج أقل من • (\% من خطأ المتوسط الحسابي, ومع ذلك فأن نموذج ARIMA أعطى أفضل نموذج وأقل نسبة من خطأ المتوسط الحسابي, أما في فيم ال COD فان نموذج ARIMA لم يعطِ أفضل النتائج. أفل نسبة من خطأ المتوسط الحسابي, و التي أعطيت من خلال نماذج ARIMA, كانت تعادل نسبة ^,؛\% في قيم الفسفور الكلي. أما في علاقة المتغيرات بعض ففد تمت معرفة مدى ترابط المتغيرات ببصن, هيئة الدتغيرات في نهر الزرقاء, مصدر المتنيرات, و معلومات أخرى عنها.

